

**Statistical Inference**  
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**Lecture - 05**  
**Basic Concepts of Point Estimations-III**

In the last lecture I introduced two Basic Concepts of Point Estimation namely unbiased estimation and consistency. One of the properties that is unbiasedness it is related to the estimator being equal to the true value on the average; that means, if we have many samples then the average of that will be equal to the true value. Whereas, consistency is a large sample property; that means, if we take a sample to be large enough then the probability that it is close enough to the true value of the parameter is almost equal to one and. So, the two properties have somewhat different applications and as well as implications and many times we try to combine various properties of the estimators.

So, I had shown in the last lecture some sort of invariance of the consistent estimators for example, if  $T$  is a consistent estimator for  $\theta$  then  $g$  of  $t$  and where  $g$  is a continuous function will be consistent for  $g$   $\theta$ . Similarly, if I have  $T_n$  to be consistent for  $\theta$  and I have sequences of numbers  $a_n$  and  $b_n$  such that  $a_n$  converges to 1 and  $b_n$  converges to 0 then  $a_n T_n$  plus  $b_n$  also is a consistent estimator for  $\theta$ . So, let me give a few examples.


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Lecture-3 ①

Examples. 1. Let  $X_1, \dots, X_n$  be a random sample from an exponential distribution with a location parameter  $\mu$ .

$$f_{X_i}(x) = \begin{cases} e^{\mu-x}, & x > \mu \\ 0, & \text{otherwise} \end{cases} \quad \left| \begin{array}{l} F(x) = \int_{\mu}^x e^{\mu-t} dt \\ = 1 - e^{\mu-x} \end{array} \right.$$

$E(X_i) = \mu + 1$   
 $E(\bar{X}) = \mu + 1$   
 $T_1 = \bar{X} - 1$  is unbiased and consistent for estimating  $\mu$ .



Let us consider say let  $X_1, X_2, \dots, X_n$  be a random sample from an exponential distribution with a location parameter  $\mu$ ; that means, I am considering the density function of say  $X_i$  to be  $e^{-\lambda(x-\mu)}$  where  $x$  is greater than  $\mu$  and 0 otherwise.

So, this is actually the well known shifted exponential distribution here  $\mu$  denotes the minimum guarantee time of the component or the life. So, here if you see expectation of  $X_i$  is equal to  $\mu + 1/\lambda$ . So,  $\mu + 1/\lambda$  is the first moment and therefore, if I consider expectation of  $\bar{X}$  that is also going to be  $\mu + 1/\lambda$ .

So, by weak law of large numbers we get  $\bar{X}$  as a consistent estimator for  $\mu + 1/\lambda$  and if I take  $1/\lambda$  to the left hand side it then we get let me call it  $T_1$ . So,  $T_1$  is equal to  $\bar{X} - 1/\lambda$  is unbiased and consistent for estimating  $\mu$ , that is the minimum guarantee time. Now, in this problem let me introduce another estimator.

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Consider  $Y = X_{(1)} = \min\{X_1, \dots, X_n\}$ .

$$\begin{aligned}
 F_{X_{(1)}}(x) &= P(X_{(1)} \leq x) \\
 &= 1 - P(X_{(1)} > x) \\
 &= 1 - P(X_1 > x, \dots, X_n > x) \\
 &= 1 - P(X_1 > x)P(X_2 > x) \dots P(X_n > x) \\
 &= 1 - \{P(X_1 > x)\}^n \\
 &= 1 - [1 - F_{X_1}(x)]^n \\
 &= 1 - e^{-n(\lambda(x-\mu))} \quad x > \mu
 \end{aligned}$$

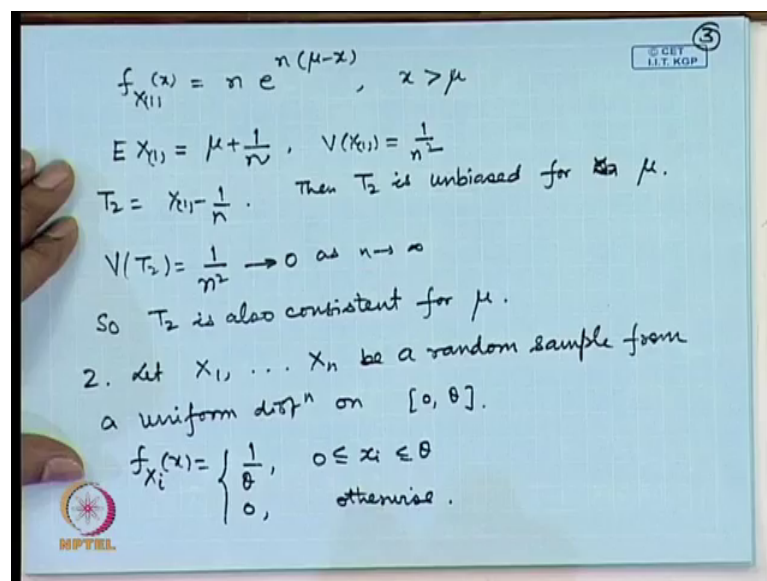
So the prob. density function of  $X_{(1)}$  is

Let us consider say  $Y$  is equal to  $X_{(1)}$  here this  $X_{(1)}$  denotes the minimum of the observations  $X_1, X_2, \dots, X_n$  one can derive the distribution of  $X_{(1)}$ . In fact, in general if I want to find out the distribution of this I can find it in the following way, I consider say c.d.f of this that is probability of  $X_{(1)}$  less than or equal to  $x$ . This can be written as 1 minus probability of  $X_{(1)}$  greater than  $x$  that I can write as 1 minus probability that now, if the minimum is greater than  $x$  this is equivalent to saying each of the  $X_i$  are greater than  $x$ .

Now, here  $X_1, X_2, \dots, X_n$  are a random sample therefore,  $X_1, X_2, \dots, X_n$  are independently and identically distributed random variables. So, this can be actually written as  $1 - \text{probability of } X_1 \text{ greater than } x$  into probability of  $X_2$  greater than  $x$  and so on. Probability of  $X_n$  greater than  $x$  that is  $1 - \text{probability of } X_1 \text{ greater than } x$  to the power  $n$  that is equal to  $1 - e^{-n(\mu - x)}$ , now this is again  $1 - e^{-n(\mu - x)}$  of  $X_1$  itself. Now if I have this as the probability density function I can write down the corresponding c.d.f here that is integral from  $\mu$  to  $x$  of  $e^{-t}$  to the power  $\mu - t$  that is equal to  $1 - e^{-n(x - \mu)}$ .

So, if I substitute  $1 - e^{-n(\mu - x)}$  here I will get  $1 - e^{-n(x - \mu)}$  to the power  $n$  times  $\mu - x$ . So, the probability density function of  $X_1$  is, now this can be obtained by considering derivative of this of course, this value I have written for  $x$  greater than  $\mu$ , if  $x$  is less than  $\mu$  then this is going to be 0.

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So, if we consider derivative of this I will get the density function of  $X_1$  as  $n e^{-n(\mu - x)}$  where  $x$  is greater than  $\mu$ . So, this is the probability density function of the minimum of the observations if I have considered a random sample from an exponential distribution with a location parameter  $\mu$ .

Now, this is the usual two parameter exponential distribution, here the scale parameter is  $\frac{1}{n}$  and location parameter is  $\mu$ . So, if I consider the expectation of  $X_1$  that is equal to  $\mu + \frac{1}{n}$ . So, if I take  $T_2$  as  $X_1 - \frac{1}{n}$  then  $T_2$  is also unbiased for  $X_1$ .

for  $\mu$ . So, I have got another unbiased estimator, now in this one I can consider variance also, what is variance of  $X_1$  for example, why variance of  $X_1$  here is  $1/n^2$ . So, variance of  $T_2$  is also  $1/n^2$  because it is variance of  $X_1$  itself, the variance of a function does not change if I make a change of origin.

Now, consider the result that if expectation is equal to the parameter and the variance goes to 0  $T_2$  is unbiased and its variance goes to 0 as  $n$  tends to infinity. So,  $T_2$  is also consistent for  $\mu$ . So, in this problem we have considered two estimators, one is based on the sample mean this is unbiased and consistent and at the same time we have considered  $T_2$  which is based on the minimum of the observations this is also unbiased and consistent.

So, that brings us to the question that if I have more than one estimator satisfying certain given desirable properties then which one we should use. So, in this direction I will give you one more example, let us consider say  $X_1, X_2, \dots, X_n$  be a random sample from a uniform distribution on the interval 0 to  $\theta$ . That means, I am considering the density of  $x_i$  as  $1/\theta$   $0 \leq x_i \leq \theta$  it is equal to 0 otherwise.

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$E(X_i) = \frac{\theta}{2}$   
 $E(\bar{X}) = \frac{\theta}{2} \Rightarrow E(2\bar{X}) = \theta$   
 $T_1 = 2\bar{X}$ , then  $T_1$  is unbiased and consistent for  $\theta$ .  
 $X_{(n)} = \max\{X_1, \dots, X_n\}$ .  
 $F_{X_{(n)}}(x) = P(X_{(n)} \leq x) = P(X_1 \leq x, \dots, X_n \leq x)$   
 $= [F_{X_i}(x)]^n = \begin{cases} 0, & x < 0 \\ \left(\frac{x}{\theta}\right)^n, & 0 \leq x \leq \theta \\ 1, & x > \theta \end{cases}$   
 $f_{X_{(n)}}(x) = \begin{cases} \frac{n x^{n-1}}{\theta^n}, & 0 \leq x \leq \theta \\ 0, & \text{otherwise.} \end{cases}$

Now in the uniform distribution we know expectation of  $X_i$  is the middle point of the interval that is  $\theta/2$ .

So, immediately we conclude that expectation of  $\bar{X}$  is  $\theta/2$  this implies that expectation of  $2\bar{X}$  is equal to  $\theta$ . So, if I call  $T_1$  is equal to  $2\bar{X}$  then  $T_1$  is unbiased and consistent for  $\theta$ . Now in this problem let me consider another one, let us consider say  $X_{(n)}$  now  $X_{(n)}$  I am calling to be the maximum of the observations. As in the previous case we can derive the distribution of  $X_{(n)}$  let us consider the cdf of this. So, this is equal to probability of  $X_{(n)}$  less than or equal to  $x$ .

Now, this is statement that the maximum is less than or equal to  $x$  is equivalent to that each of the observations is less than or equal to  $x$  and once again using the fact that  $X_i$ 's are independently and identically distributed this is equivalent to saying each of the  $X_i$ 's cdf at  $x$ . So, this is simply this to the power  $n$ , now for the uniform distribution the cdf is it is equal to 0 if  $x$  is less than 0, it is equal to  $x/\theta$  if 0 is less than or equal to  $x$  is less than or equal to  $\theta$  it is equal to 1 if  $x$  is greater than  $\theta$ . So, if we use this cdf here this becomes 0 if  $x$  is less than 0 it is equal to  $x/\theta$  to the power  $n$  if 0 less than or equal to  $x$  is less than or equal to  $\theta$  it is equal to 1 if  $x$  is greater than  $\theta$ . One may find out the probability density function from here, by considering the derivative because this is a continuous distribution.

So, you get the density function as  $n x^{n-1} / \theta^n$  if  $x$  lies between 0 to  $\theta$  and it is 0 otherwise.

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$$E(X_{(n)}) = \int_0^{\theta} x \cdot f_{X_{(n)}}(x) dx = \int_0^{\theta} \frac{n x^n}{\theta^n} dx = \frac{n}{n+1} \theta$$

$$E\left\{\left(\frac{n+1}{n}\right) X_{(n)}\right\} = \theta$$

$T_2 = \frac{n+1}{n} X_{(n)}$ . Then  $T_2$  is unbiased for  $\theta$ .

$$\lim_{n \rightarrow \infty} F_{X_{(n)}}(x) = \begin{cases} 0, & x < \theta \\ 1, & x \geq \theta \end{cases}$$

This is the cdf of a r.v. which takes value  $\theta$  with probability 1. So  $X_{(n)} \xrightarrow{P} \theta$

This fact can also be proved by considering

Let us consider say expectation of  $X_n$  now. So, expectation of  $X_n$  is equal to integral  $x$  into the density function of  $X_n$  from 0 to  $\theta$  that is equal to integral 0 to  $\theta$   $n x$  to the power  $n$  by  $\theta$  to the power  $n$   $dx$ . So, as we can see easily the integral of  $x$  to the power  $n$  will be  $x$  to the power  $n+1$  by  $n+1$  and if we substitute the limits from 0 to  $\theta$  I will get  $\theta$  to the power  $n+1$ . And in the denominator I have  $\theta$  to the power  $n$  so that will cancel and therefore, this value will be equal to  $n$  by  $n+1$   $\theta$ .

If I adjust this  $n$  by  $n+1$  on the left hand side I get this is equal to  $\theta$ . So, if I use the notation  $T_2$  as  $n+1$  by  $n \bar{X}_n$  then  $T_2$  is unbiased we had obtained estimator  $T_1$  as  $2 \bar{X}$  which is unbiased and now I have obtained  $T_2$ . Let us check say probability of  $|X_n - \theta| > \epsilon$  whether it tends to 0 or not. That we can check from here also if we take the limit of this cumulative distribution function now here  $x$  is less than or equal to  $\theta$ . So, this value will tend to 0 if  $x$  is less than  $\theta$  and whenever  $x$  is greater than or equal to  $\theta$  it is becoming 1.

So, if I take the limit of if I consider say limit of  $F_{X_n}$  as  $n$  tends to infinity then this is equal to 0 for  $x$  less than  $\theta$  and it is equal to 1 for  $x$  greater than or equal to  $\theta$ . Now this denotes the distribution of a random variable which takes value only  $\theta$ , this is the c d f of a random variable which takes value  $\theta$  with probability 1. So, basically we have proved that  $X_n$  converges to  $\theta$  distribution, but  $\theta$  is a constant therefore, it is equivalent to saying  $X_n$  converges to  $\theta$  in probability. In fact, this fact can also be proved in a different way, I will consider directly the definition of convergence in probability.

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$$P(|X_{(n)} - \theta| > \epsilon) = P(\theta - X_{(n)} > \epsilon)$$
$$= P(X_{(n)} < \theta - \epsilon) = \left(\frac{\theta - \epsilon}{\theta}\right)^n \rightarrow 0 \text{ as } n \rightarrow \infty$$

So  $T_3 = X_{(n)}$  is consistent for  $\theta$

$\Rightarrow T_2$  is also consistent for  $\theta$ .

Let  $X_1, \dots, X_n$  be a random sample from a continuous population with cdf  $F(x)$  and the range of variables is interval  $[a, b]$ .

Let  $U = X_{(1)}, V = X_{(n)}$ .

$$F_n(u) = \begin{cases} 0, & x < a \\ 1 - [1 - F(u)]^n, & a \leq x < b \\ 1, & x \geq b \end{cases}$$

Let me take say probability of modulus  $X_n$  minus  $\theta$  say greater than  $\epsilon$ .

Now, the distribution of  $X_n$  is in the interval 0 to  $\theta$ ; that means,  $X_n$  is always below  $\theta$ . So, if we consider this modulus of  $X_n$  minus  $\theta$  this is same as  $\theta - X_n$ . So, this is equivalent to probability of  $\theta - X_n$  greater than  $\epsilon$  which I can write as probability of  $X_n$  less than  $\theta - \epsilon$ . Now this is nothing, but the distribution function of  $X_n$  at the point  $\theta - \epsilon$  since  $X_n$  is having a continuous distribution whether we put less than or less than or equal to it does not make a difference.

Therefore, this value is equal to  $\frac{\theta - \epsilon}{\theta}$  to the power  $n$  as we have derived just in the previous sheet here. So, now you can see  $\epsilon$  is a positive number. So,  $\theta - \epsilon$  is less than  $\theta$  therefore, if I take the limit as  $n$  tends to infinity this will go to 0. So,  $X_n$  is consistent for  $\theta$  let me call it  $T_3$ ; if we look at the coefficient  $n + 1$  by  $n$  this goes to 1 as  $n$  tends to infinity.

So, we have the result that if  $T_n$  is consistent for  $\theta$  and  $a_n$  goes to 1 then  $a_n T_n$  is also consistent therefore, if we use this fact  $T_2$  is also consistent for  $\theta$ . So,  $T_2$  is unbiased and consistent  $T_1$  is also unbiased and consistent. So, I have given you two different distributions where for estimating of one parameter I am getting 2 different unbiased and consistent estimators. So, that shows that actually we need additional criteria to distinguish between or to choose between various competing estimators. From

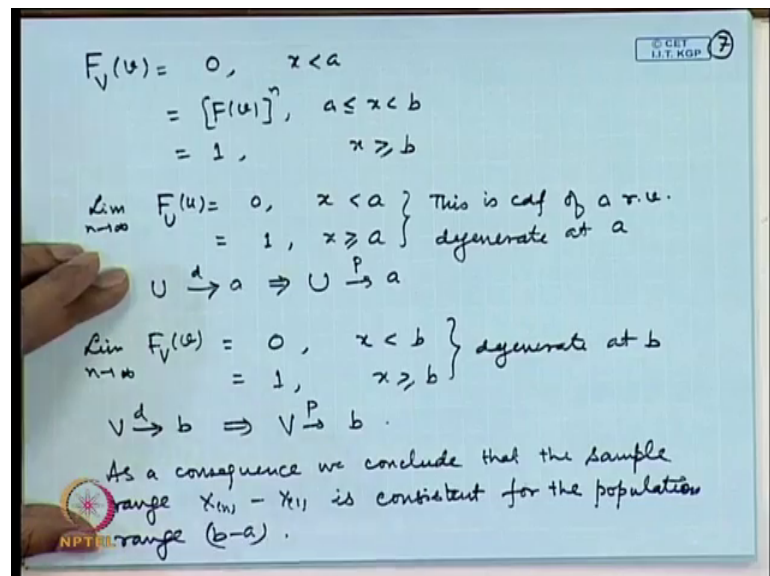


the previous two exercises we can also find something more important. If we look at the form of the distribution function of the maximum and the minimum there is some a specific structure here for example, when I took the limit here I got only 0 1 here.

Similarly, in the distribution of the minimum we had that  $X_1$  is converging to  $\mu$ . So, minimum was converging to the lower limit and here the maximum is converging to the upper limit  $\theta$ . In fact, if we have any continuous distributions then this is a general fact. So, I will state it in the following results. So, let me give it as exercise let  $X_1, X_2, \dots, X_n$  be a random sample from a continuous population with c.d.f say capital  $F(x)$  and the range of variables is interval  $a$  to  $b$ . Of course, this interval may be open or closed that does not make any difference if we are handling a continuous distribution. Let us define say  $U$  is equal to the minimum and  $V$  is equal to  $X_n$  then the claim is that  $U$  is a consistent estimator for  $a$  and  $V$  is a consistent estimator for  $b$ .

So, the proof will use the steps which we have derived just now, that is the c.d.f of  $U$  that is  $1 - F(x)^n$ . So, actually it is equal to 0 for  $x < a$  it is equal to  $1 - F(x)^n$  for  $a \leq x < b$  it is equal to 1 for  $x \geq b$ .

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Similarly if I consider say  $F_V$  then it is equal to 0 for  $x < a$  it is equal to  $F_V$  to the power  $n$  for  $a \leq x < b$  it is equal to 1 for  $x \geq b$ .



equal to  $b$  notice here that this equality or inequality does not make any difference here because it is continuous distribution.

So, if I take the limits here now  $F$  is a number between 0 to 1. So, if I take the limits here this number will go to 0; so, this is going to 1. So, if I take the limit as  $n$  tends to infinity of  $F_U$  I am getting 0 for  $x$  less than  $a$  and it is equal to 1 for  $x$  greater than or equal to  $a$ .

So, this is  $c.d.f$  of a random variable which is simply degenerate at  $a$ . So, we can conclude that  $u$  converges to  $a$  in distribution and therefore,  $u$  converges to  $a$  in probability because convergence in distribution and probability are equivalent if the right hand side is a constant. Similarly if I consider limit of  $F_V$  as  $n$  tends to infinity then this is also 0 for  $x$  less than  $b$  and it is equal to 1 for  $x$  greater than or equal to  $b$ .

So, once again  $V$  is converging to  $b$  this random variable is degenerate at  $b$ . So  $V$  tends to  $b$  in distribution or  $V$  tends to  $b$  in probability. So, if a continuous distribution is having a range  $a$  to  $b$  then the smallest order statistics converges to the lowest value or the lowest value in the range and the largest order statistic converges to.

So, these can be treated as the consistent estimators of these respective parameters. So, this actually gives some easy applications basically for example, we want to find out a consistent estimator for the range. For example, here range may be  $b$  minus  $a$  then easily you can say that  $V$  minus  $u$  that is the maximum minus the minimum, sample range is the consistent estimator for the population range.

As a consequence we conclude that the sample range that is  $X_n$  minus  $X_1$  is consistent for the population range  $b$  minus  $a$ . We have some further special cases here for example, lower limit could be minus infinity or the upper limit could be plus infinity. In that case for example, if the lower limit is minus infinity then  $X_1$  does not converge in probability. Similarly if the highest value is unbounded; that means,  $b$  is infinity then  $x_n$  does not converge in probability.