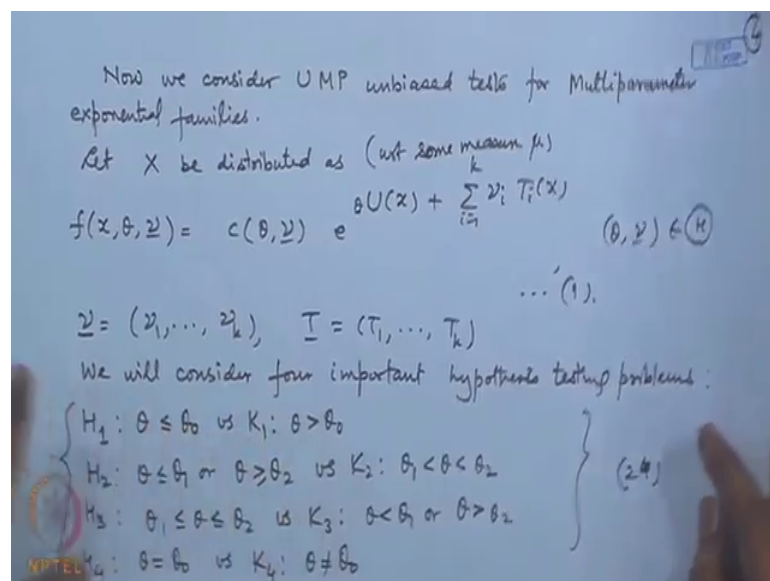


Statistical Inference
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Lecture – 47
Unbiased Test for Normal Populations – I

Yesterday, we have discussed in the previous lecture UMP unbiased tests for the multi parameter exponential family of distributions let me just recollect the discussion.

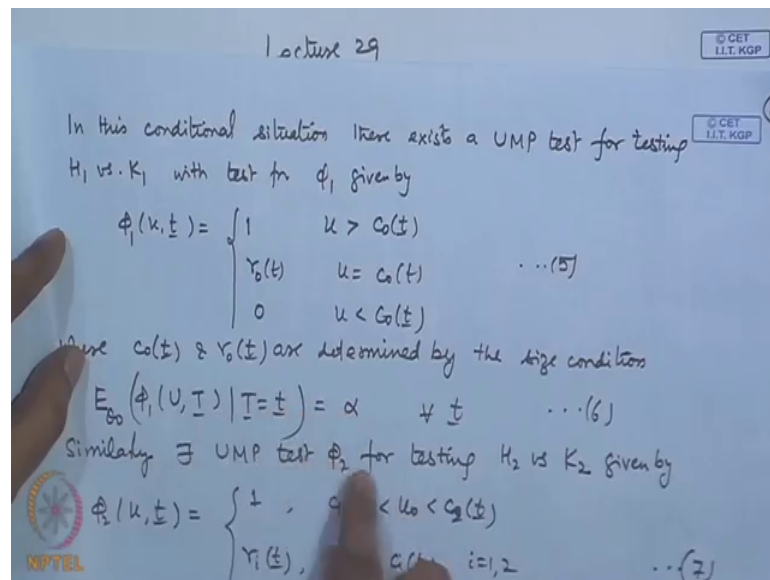
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We had considered a distribution of the type $e^{\theta U(x) + \sum_{i=1}^k \nu_i T_i(x)}$. Here we are able to derive the UMP unbiased test for 4 types of hypothesis called H_1, H_2, H_3, H_4 versus K_1, K_2, K_3, K_4 respectively for θ and here $\nu_1, \nu_2, \dots, \nu_k$ are considered as the nuisance parameter of course, we had considered here that the parameter space is $k+1$ dimension and convex also.

A peculiar nature of these tests was that these tests were conditional.

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Let me show the form here for at least one of them. So, if you recollect the form the tests were of the form phi 1. So, I am calling it condition on u greater than C naught t and here this term is gamma naught t and this C naught and gamma naught are determined from the conditional thing. A extension of this was done later on where we called this as the unconditional tests also.

However the derivation of this we demonstrated by application to comparison of two binomial proportions two Poisson arrival rates and so, on and in also in testing for independence and contingency table in each of these cases we saw that we have to actually determine the conditional distribution of u given t.

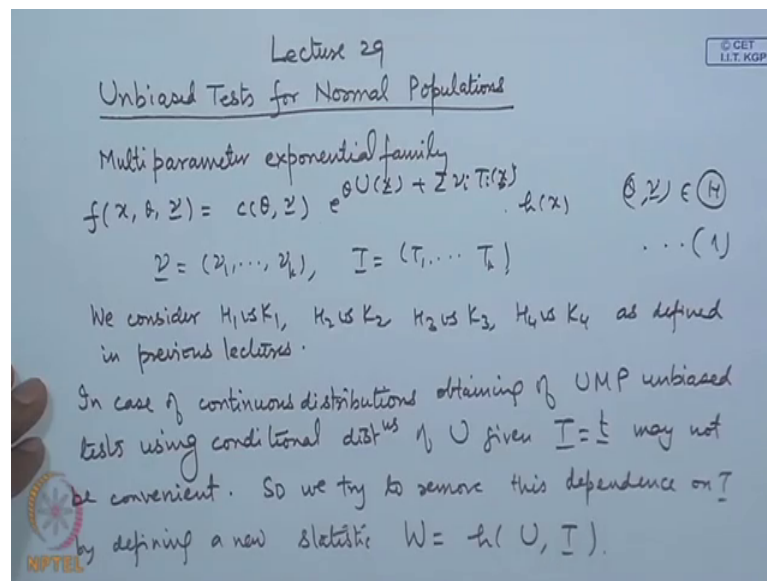
Now in the case of continuous distributions for example if we consider normal distributions gamma distributions etcetera then these conditional distributions are not so, easy. Because in the discrete case we are able to write down the conditional distribution in terms of probability and we are able to apply the formula for the condition probability that is probability of a given b is equal to probability of a intersection b divided by probability of b, and we are able to actually derive the exact form of the test.

In the case of a continuous, this conditional distribution may not be so, easy. Therefore, we apply a method by which we can modify these conditions u greater than or u less than some constant which is dependent upon t to something some another statistic let us call it say w so, that this condition becomes free from t. That means, this new w variable which

I am saying it could be a function of u and t should be defined in such a way that first of all it should be an increasing function or monotonic function so, that the conditions of inequality remain same or they get reversed.

And second thing is that the independence. So, fortunately there is a method and I will explain that method now in this lecture here.

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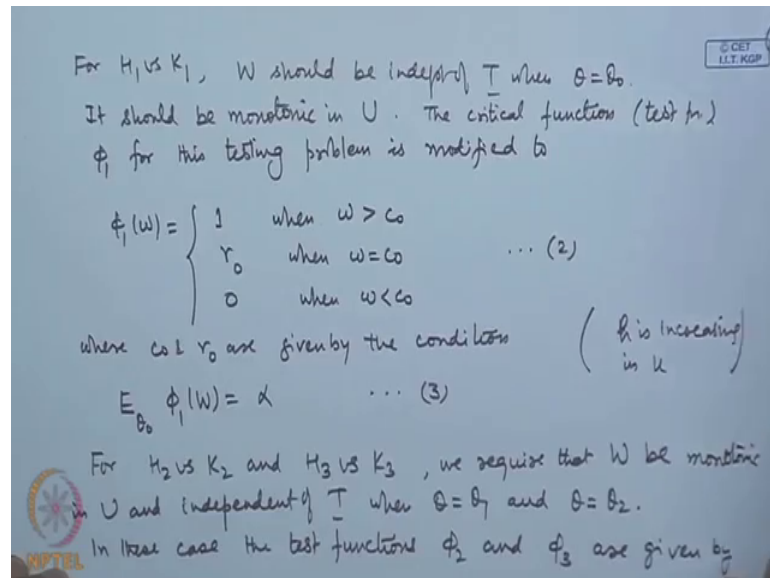


So, let us consider multi parameter exponential family; multi parameter exponential family. So, the density is of the form let me write from the yesterdays discussion f of x theta nu that is equal to some constant times e to the power theta U x plus sigma mu i T i x and of course, there will be some term here which is consisting of x here. Here nu is equal to nu 1 nu 2 nu K and t is equal to $T_1 T_2 T_k$ and this theta nu belongs to certain parameter space say script theta.

So, I call this density 1 as before I consider hypothesis testing problems H_1 versus K_1 , H_2 versus K_2 , H_3 versus K_3 H_4 versus K_4 as defined in previous lectures.

In case of continuous distributions obtaining of UMP unbiased tests using conditional distributions of U given T is equal to t may not be convenient. So, we try to remove this dependence on T by defining a new statistic let us call it say w is equal to a function of U and T .

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Now, the conditions that we need to impose are as follows for H_1 versus K_1 , W should be independent of T when θ is equal to θ_0 . Remember here the hypothesis H_1 here the condition was based on θ_0 if you remember the exact condition here. Let me recollect from yesterday's lecture expectation θ_0 here similarly if you see ϕ_2 and ϕ_3 here the condition is on θ_1 and θ_2 . And similarly for H_4 the condition is on again θ_0 the condition is on θ_0 . So, we have to look at the independence of W from T at these points

So, W should be independent of θ of T when θ is equal to θ_0 it should be monotonic in U the critical function or the test function ϕ_1 for this testing problem is modified to ϕ_1 of w is equal to 1 when w is greater than C_0 it is equal to γ_0 when w is equal to C_0 it is equal to 0 when w is less than C_0 where C_0 and γ_0 are given by the condition expectation of ϕ_1 w is equal to α .

One thing we should notice here I have maintained the same sign like here U was greater than C_0 and here it is w is greater than C_0 . If we do that then we have assumed that h is increasing in u if h is decreasing in u then this will get reversed here. For H_2 versus K_2 and H_3 versus K_3 we require that W be monotonic in U and independent of t when θ is equal to θ_1 and θ is equal to θ_2 . So, in these cases the test functions ϕ_2 and ϕ_3 are given by ϕ_2 w is equal to 1, then c_1 is less

than w less than c_2 it is equal to γ_i when w is equal to c_i , i is equal to 1, 2 and it is equal to 0.

If w is less than c_1 or w is greater than c_2 and c_i is and γ_i is are determined by expectation of $\phi_2(W)$ at θ_1 and expectation of $\theta_2 \phi_2(W)$ is equal to α .

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$$\phi_2(w) = \begin{cases} 1 & \text{when } c_1 < w < c_2 \\ \gamma_i & \text{when } w = c_i, i=1,2 \\ 0 & \text{when } w < c_1 \text{ or } w > c_2 \end{cases} \quad (4)$$

where c_i 's & γ_i 's are determined by

$$E_{\theta_1} \phi_2(W) = E_{\theta_2} \phi_2(W) = \alpha \quad \dots (5)$$

$$\phi_3(w) = \begin{cases} 1 & \text{when } w < c_1 \text{ or } w > c_2 \\ \gamma_i & \text{when } w = c_i, i=1,2 \\ 0 & \text{when } c_1 < w < c_2 \end{cases}$$

where c_i 's & γ_i 's are determined by

$$E_{\theta_1} \phi_3(W) = \alpha \quad i=1,2 \quad (6)$$

Similarly ϕ_3 is given by the reciprocal of this one that is it will become 1 when w is less than c_1 or w is greater than c_2 it is equal to γ_i when w is equal to c_i for i is equal to 1, 2 and it is equal to 0 if c_1 is less than w is less than c_2 . c_i is and γ_i is are determined by the size conditions expectation $\theta_i \phi_3(w)$ is equal to α for i is equal to 1, 2.

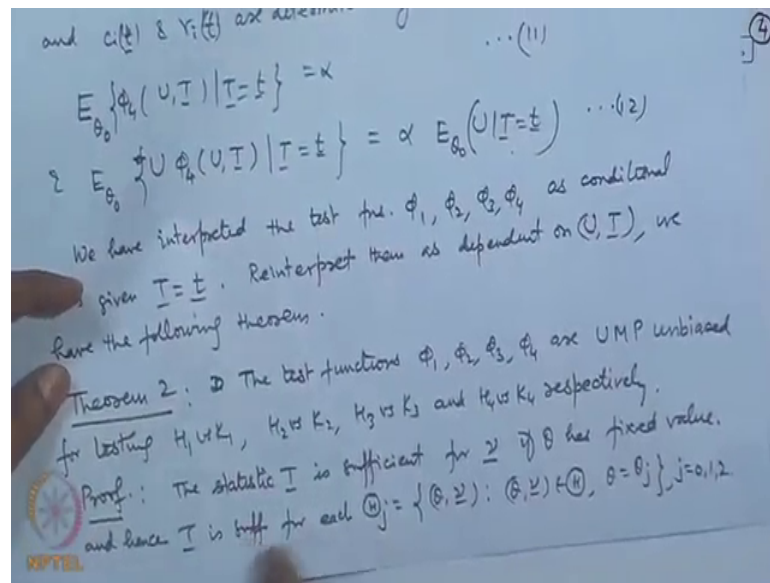
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For H_4 vs K_4 , we need $h(u, t) = a(t)u + b(t)$, ($a(t) > 0$)
 and W to be independent of T when $\theta = \theta_0$.
 Then ϕ_4 can be described as
 $\phi_4(w) = \phi_3(w) \dots (7)$
 where c_i 's & γ_i 's are given by
 $E_{\theta_0} \phi_4(w) = \alpha \dots (8)$
 $E_{\theta_0} W \phi_4(w) = \alpha E_{\theta_0} \phi_4(w) \dots (9)$
Theorem 3: Suppose X has dist^n in multiparameter expo family.
 & $W = h(U, T)$ is increasing in U . Then for H_1 vs K_1 , ϕ_1 is UMP
 unbiased of W is indep of T when $\theta = \theta_0$. For H_2 vs K_2 , ϕ_2
 is UMP unbiased of W is independent of T when $\theta = \theta_1$ & $\theta = \theta_2$.

Now, for the testing problem H_4 versus K_4 , now here if you remember for ϕ_4 we had two conditions. The second condition was involving product of U with ϕ_4 . So, now, if you translate we need another condition that U should be H should be linear function

So, let me write it here we need $h(u, t)$ to be a linear function of u of course, again I am taking a 2 to be 0 positive if it is negative and the region will get reversed and w to be independent of t , when θ is equal to θ_0 . Then this ϕ_4 can be described as $\phi_4(w)$ is equal to say $\phi_3(w)$ where c_i is and γ_i is are given by expectation of $\phi_4(w)$ is equal to α at θ_0 and expectation of $w \phi_4(w)$ is equal to α times expectation of $\phi_4(w)$ at θ_0 .

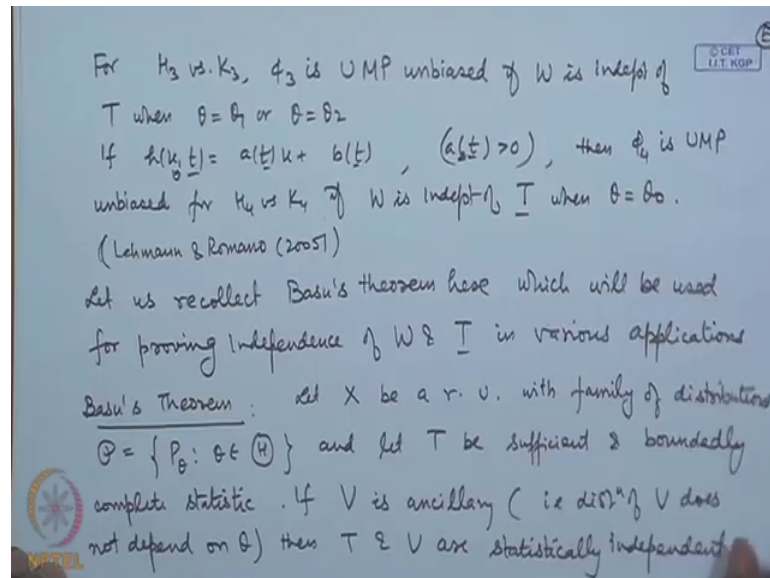
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So, note here unlike the theorem 2 which I gave yesterday, we had there all the constants depending upon t that is they were calculating from the conditional distribution of u given t . But here you see that none of them is dependent upon t , the expressions that I have written they have become free from t . So, this is the advantage of this technique and in order to imply this technique, we need to suitably define this function H for various hypothesis testing problems.

So, I summarize all these results in the following theorem, let me call it theorem 3. Suppose X has distribution in multi parameter exponential family 1, and W is equal to h of U T is increasing in U and increasing in U . Then for H_1 versus K_1 ϕ_1 is UMP unbiased if W is independent of T when θ is equal to θ_1 for H_2 versus K_2 ϕ_2 is UMP unbiased. If W is independent of t then θ is equal to θ_1 and θ is equal to θ_2 .

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For H_3 versus K_3 ϕ_3 is UMP unbiased if W is independent of T then θ is equal to θ_1 or θ is equal to θ_2 . If $h(u, t)$ is of the form $a(t)u + b(t)$ where $a(t)$ is positive, then ϕ_4 is UMP unbiased for H_4 versus K_4 , if W is independent of t when θ is equal to θ_0 .

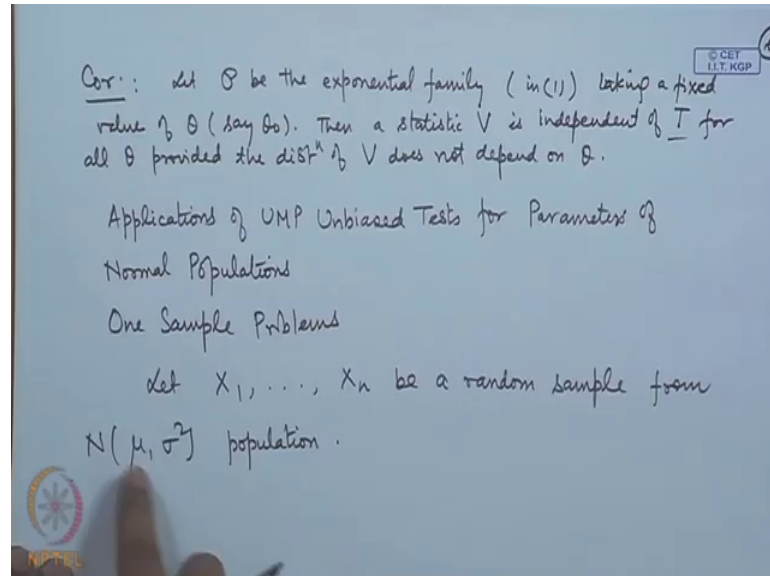
For the proofs one may look at the Lehmann and Romano's book as we have been following this theory from this text here. Now in order to have the independence of w and t we will usually require certain result which we can use for independence. Now one of the important results for proving independence now remember here either u or t they are sufficient or. In fact, if I were to look at this in the full version then u and t is complete and sufficient.

If we fix θ then t is sufficient and conversely if we fix θ then u is sufficient. So, what happens that certain independence is there if I can use Basu's theorem. So, sufficiency and then we should have completeness or bounded completeness and then we should have an ancillary. Let me recollect the Basu's theorem here which will be used for proving independence of w and t in various applications.

As I mentioned I will be briefly discussing mainly discussing the applications for testing problems in normal distributions. So, the Basu's theorem statement let me repeat here, let X be a random vector with family of distributions \mathcal{P}_θ , $\theta \in \Theta$ and let T be sufficient and boundedly complete a statistic.

If V is ancillary that is distribution of V does not depend on θ then T and V are having independent distributions.

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So, what is important is that we should have one sufficient and complete or sufficient and boundedly completed statistic, and another one should have a distribution free from the parameters. If that happens then T and V are statistically independent.

So, in general we will try that, that W should be complete and sufficient and then T should become ancillary or vice versa. As a corollary we have the following important result let P be the exponential family in 1 and taking a fixed value of say θ , then a statistic V is independent of T for all θ provided the distribution of V does not depend on θ .

If you look at the family in one if I fix the value of θ say θ_0 , then sufficient statistic will become $T_1 T_2 \dots T_k$. Now if I have another statistic V whose distribution will not depend upon θ then certainly V and T will be independent. So, we will try to use this now let us consider applications to applications of UMP unbiased tests for parameters of normal populations.

Ah To start with let us consider one sample problems, since we are having $X_1 X_2 \dots X_n$ a random sample from normal μ σ^2 distribution. Remember here I have discussed this normal distribution earlier also, I have considered testing for the mean of a

normal distribution, testing for the variance in a normal distribution, but the crucial difference was that when I was testing for the mean I had considered variance to be known.

And accordingly the tests which were either UMP for H_1 and H_2 and for H_3 and H_4 it was uniformly unbiased we were had obtained. Similarly for sigma square when I was doing the testing the ν was taken to be known and I have taken without loss of generality to be 0 and once again we had the UMP test for H_1 and H_2 and UMP unbiased test for H_3 and H_4 .