

Statistical Inference
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Lecture - 39
UMP Tests- III

In the last class I have also considered the cases for two sided composite hypothesis and there is one particular case, when we are having the null hypothesis as a two sided. Like $\theta < \theta_1$ or $\theta > \theta_2$ against θ lying in the interval $\theta_1 < \theta < \theta_2$. In these cases also the uniformly most powerful test exists provided the distributions are in the 1 parameter exponential family. And the $Q(\theta)$ function which is there in the 1 parameter exponential family should be strictly monotone.

So, basically we have given 2 results; one is that if the families of distributions have monotone likelihood ratio then for the one sided testing problems like $\theta < \theta_1$ or $\theta > \theta_2$ against $\theta_1 < \theta < \theta_2$ or the dual of it for these problems UMP test can be derived. So, now I will discuss various applications of these in both the results; let me start with the normal distribution ok.

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Lecture 25

UMP Tests : Applications

Example: Let us return to testing for mean in a normal population. Let $X_1, \dots, X_n \sim N(\theta, 1)$

We want to test $H_0: \theta \leq \theta_1$ or $\theta \geq \theta_2$
 vs $H_1: \theta_1 < \theta < \theta_2$ ($\theta_1 < \theta_2$).

So by the previous theorem, the UMP test is given by

$$f(x, \theta) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x-\theta)^2}$$

$$f(z, \theta) = \frac{1}{(\sqrt{2\pi})^n} e^{-\frac{n\theta^2}{2} - \frac{z^2}{2}}$$

$$\phi(x) = \begin{cases} 1 & \text{if } c_1 < \bar{X} < c_2 \\ \gamma_i & \text{if } \bar{X} = c_i, i=1,2 \\ 0 & \text{if } \bar{X} < c_1 \text{ or } \bar{X} > c_2 \end{cases}$$

where γ_i 's & c_i 's are determined by

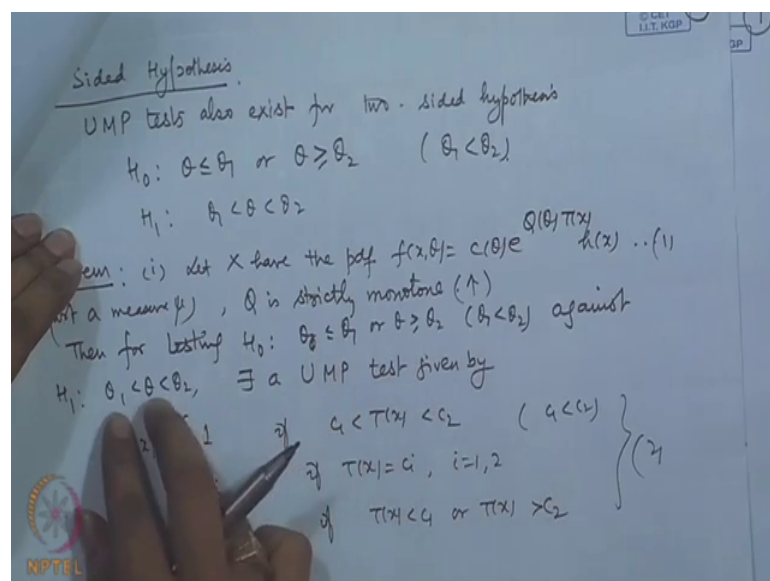
So, let us return to testing for mean in a normal population; that means, we are having a set of that X_1, X_2, \dots, X_n follows a normal; so, its a random sample from a normal $\theta_1 < \theta < \theta_2$

distribution. And we want to test $H_0: \theta \leq \theta_1$ or $\theta \geq \theta_2$ against $H_1: \theta_1 < \theta < \theta_2$. So, here of course, we have assumed $\theta_1 < \theta_2$.

So, in the previous class I have given the theorem. So, here we have 1 parameter exponential family; see if we write down the distribution it is $\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2\theta}}$ that is equal to $\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2\theta}}$. And if we write for $f(x; \theta)$ where $x = x_1, x_2, \dots, x_n$ then this is $\frac{1}{\sqrt{2\pi}^n} e^{-\frac{1}{2\theta} \sum_{i=1}^n x_i^2}$.

So, this is in the form of a 1 parameter exponential family the $Q(\theta)$ function is θ it is strictly increasing function. So, strictly monotone therefore, the theorem which I gave in the last lecture let me just show it again; let us look at the statement of the result.

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If we have $f(x; \theta) = c(\theta) e^{Q(\theta)T(x)} h(x)$, where Q is strictly increasing function. Then for testing two-sided null hypothesis against an interval for the alternative hypothesis UMP test exist. And the tests as this form which is having this it is based on $T(x)$; $T(x)$ function which is available here. Therefore, we can straightforwardly write down here the test based on $\sum x_i$ or \bar{x} . So, by the

previous theorem the UMP test is given by $\phi(x)$ is equal to 1 if c_1 is less than \bar{X} less than c_2 .

It is γ_i if \bar{X} is equal to c_i i is equal to 1, 2, it is equal to 0 if \bar{X} is less than c_1 or \bar{X} is greater than c_2 . So, the presentation of the test is that we reject the null hypothesis H_0 if \bar{X} lies between c_1 and c_2 . And we reject with probability γ_i if \bar{X} is equal to c_i for i is equal to 1, 2. And we accept H_0 if \bar{X} is less than c_1 or \bar{X} is greater than c_2 where, this γ_i 's and c_i 's are determined by the size conditions.

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$(*) E_{\theta_1} \phi(\bar{X}) = E_{\theta_2} \phi(\bar{X}) = \alpha$
 Since $\bar{X} \sim N(\theta, 1/n)$, we may take $r_1 = r_2 = 0$ (wlog)
 Now (*) gives $P_{\theta_i} (c_1 < \bar{X} < c_2) = \alpha, i=1,2$

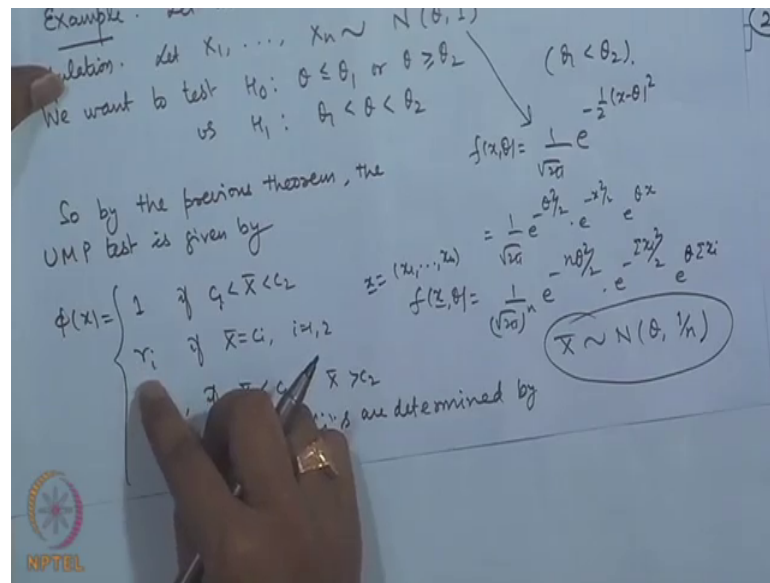
$$P_{\theta_i} (\sqrt{n}(c_1 - \theta_i) < \underbrace{\sqrt{n}(\bar{X} - \theta_i)}_{Z \sim N(0,1)} < \sqrt{n}(c_2 - \theta_i)) = \alpha, i=1,2$$

$$\Rightarrow \Phi(\sqrt{n}(c_2 - \theta_i)) - \Phi(\sqrt{n}(c_1 - \theta_i)) = \alpha, i=1,2$$

 For given values of $\theta_1, \theta_2, n, \alpha$, we can solve the above equations to determine c_1 & c_2
 e.g. $n=9, \theta_1=0, \theta_2=1, \alpha=0.05$. Then the above equations reduce to

Expectation of $\phi(\bar{X})$ under θ_1 and under θ_2 to be equal to α . Now, note here I am considering \bar{X} is equal to c_i that distribution of \bar{X} will be; the distribution of \bar{X} is normal θ_1 by n .

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This is a continuous a distribution therefore, without loss of generality we can take gamma i is to be 0. If it is to be 0 then this point is included here. Since, X bar follows normal theta 1 by n we may take gamma 1 is equal to gamma 2 equal to 0 without loss of generality. So, now we want this condition size condition star probability of c 1 less than X bar less than c 2 is equal to alpha, for i is equal to 1 2. Now, when i is equal to 1 then X bar follows normal theta 1 1 by n.

So, for i is equal to 1 this condition then can be written as we can transfer to the standard normal variable. We will get root n c 1 minus let me write for i minus less than root n X bar minus theta i less than root n c 2 minus theta i is equal to alpha, for i is equal to 1 2. Now, when theta is equal to theta i this is normal 0 1. So, this is reducing to phi of root n c to minus theta i minus phi of root n c 1 minus theta i is equal to alpha i is equal to 1 2.

So, for given values of theta 1, theta 2, n, alpha we can solve the above equation to determine c 1 and c 2 and of course, this will be numerical solutions. As an example let us take say suppose, I take say n is equal to 9 let me take say theta 1 is equal to say 0, theta 2 is equal to say 1 and say alpha is equal to 0.05. Then what will be these equations?

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$\Phi(3c_2) - \Phi(3c_1) = 0.05 \dots (1)$
 $\Phi(3(c_2-1)) - \Phi(3(c_1-1)) = 0.05 \dots (2)$
 These can be solved from the tables of $\Phi \rightarrow$ cdf of standard normal dist.
Remark: The UMP test for the dual problem
 $H_0: \theta_1 \leq \theta \leq \theta_2$ vs $H_1: \theta < \theta_1$ or $\theta > \theta_2$
 or for $H_0^*: \theta = \theta_0$ vs $H_1^*: \theta \neq \theta_0$
 do not exist. We can show this by an example.
Example: X_1, \dots, X_n a random sample from double exponential dist.

$$f(x) = \frac{1}{2\theta} e^{-\frac{|x|}{\theta}}, \quad -\infty < x < \infty$$

Then the above equations reduce to the first equation will become; now root n is 3 c 2 if I am writing theta is equal to 0 then it will be phi of 3 c 2 minus phi of 3 c 1 is equal to 0.05 that will be 1 equation. And the other equation will become 3 of c 2 minus 1 minus phi of 3 into c 1 minus 1 is equal to 0.05. These can be solved from tables of capital phi function, this is the cdf of a standard normal distribution that we have been using. So, once again I have demonstrated here that under the given conditions UMP test for a testing problem can be provided.

And this helps us in taking exact decisions at a given level of significance and of course, the given level of significance may depend upon the problem that is given at hand. Let me give some further applications. Now, another point which I would like to mention here, that I have considered here the region of null hypothesis as two sided. And the region for alternative hypothesis is the complementary of that is it is within an interval.

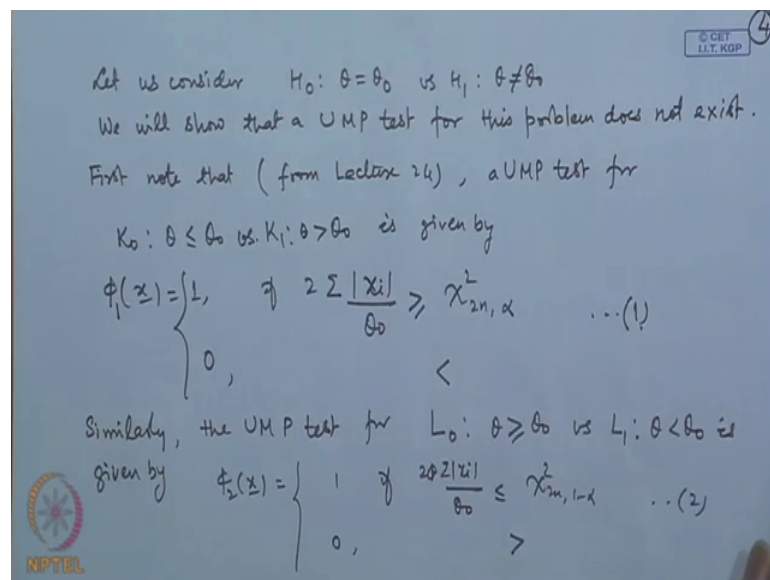
Now, one may think that if the UMP tests exists for this problem the if I interchange it, like if I write this has H naught and this is H 1; that means, the alternative is two sided. Unfortunately, in these cases it can be shown that the UMP test does not exist; I will demonstrate by it an example.

Let me give this comment here; the UMP test for the dual problem H naught theta 1 less than or equal to theta less than or equal to theta 2 versus H 1 theta less than theta 1 or theta greater than theta 2 or for let me say H naught star theta is equal to theta naught

versus $H_1: \theta \neq \theta_0$ does not exist. So, let us take the example we can show so, let us take this example. We have considered earlier a double exponential distribution, let us consider say X_1, X_2, \dots, X_n a random sample from double exponential $1/\theta \exp(-|x|/\theta)$.

And here of course, both θ and x have the range on the whole real line. This problem you have discussed in the previous lecture. I had demonstrated the UMP test for the two sided for one sided testing problem, that is for $\theta \leq \theta_0$ against $\theta > \theta_0$. Now, here I will show that if I consider this type of hypothesis then the UMP test does not exist.

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So, let us consider say $H_0: \theta = \theta_0$ against say $H_1: \theta \neq \theta_0$. We will show that a UMP test for this problem does not exist. So, if you go back to the development that, I gave it in the last lecture; what we have shown here if you considered with example discussed in the last lecture. I have considered the one sided testing problem $\theta \leq \theta_0$ against $\theta > \theta_0$. And we derived the UMP test of the having form that reject H_0 if $\sum_{i=1}^n |X_i|$ is greater than or equal to c .

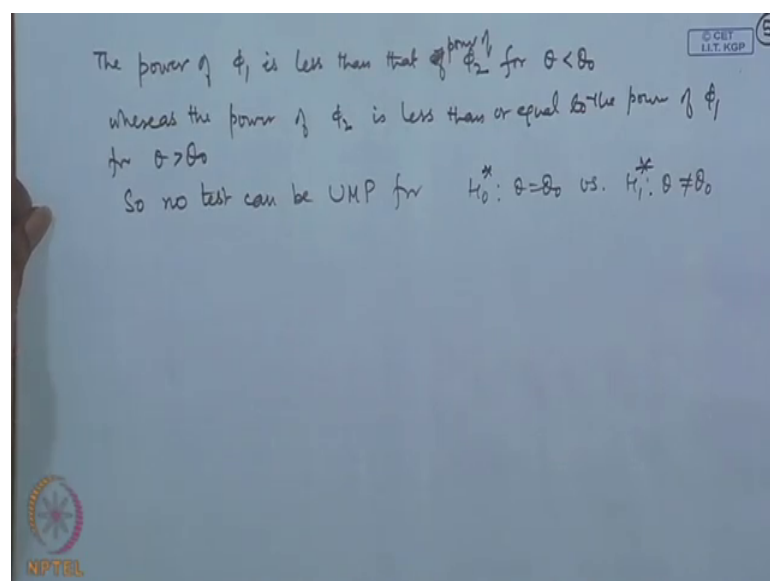
And we were able to determine this constant also; the final form was reject H_0 if $\sum_{i=1}^n |X_i| \geq c$ by $\theta_0 \sum_{i=1}^n |X_i| \geq \chi_{2n, \alpha}^2$. So, let us write this first note that; so, I am giving the reference from lecture 24 a UMP test

for; let me give some different names than H_0 and H_1 . We can consider say K_1 $\theta \leq \theta_0$ against K_2 $\theta > \theta_0$. This is given by $\phi(x) = 1$ for $\frac{2 \sum x_i}{n} \geq \theta_0$ and 0 if it is less.

Now, if you considered the dual problem here, the dual problem is to consider H_0 $\theta > \theta_0$ against H_1 $\theta \leq \theta_0$. Then in that particular case the rejection region will become less here, the reason is that we are having the monotonic equation ratio in θ and $\sum x_i$. So, the rejection region will become less than or equal to here. And when we proceed in this fashion the constant this will become $\chi^2_{2n-1, 1-\alpha}$. Therefore, we can write that form here; let me call this is as say ϕ_1 .

Similarly, the UMP test for technical this L_1 $\theta > \theta_0$ against L_2 $\theta \leq \theta_0$; this is given by $\phi_2(x) = 1$ if $\frac{2 \sum x_i}{n} \leq \theta_0$ and 0 if it is greater. Now, by the property of the UMP test if I look at the power functions then the power function of this will be having the values in the $\theta > \theta_0$ region. Now, $\theta > \theta_0$ region is the region of the null hypothesis for the L_1 . So, naturally this value will be larger than this value for $\theta > \theta_0$. So, let me write this comment here.

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The power of ϕ_1 is less than that of ϕ_2 for θ less than θ_{naught} . Why? Because, for θ less than θ_{naught} ϕ_1 is having the size that is the level of significance or the probability of type 1 error that is less than or equal to α . The maximum value is attained at θ is equal to θ_{naught} . So, the power function of ϕ_1 for θ less than θ_{naught} is actually the probability of type 1 error which is less than or equal to α ; whereas, for the ϕ_2 it is the probability of rejecting when H_1 is true. That is actually it is a 1 minus the probability of type 2 error, that value is greater than or equal to α because, the minimum value that is attained at θ_{naught} .

So, what we are getting here is this is less than or equal to that power of ϕ_2 for this. And if I consider the power of ϕ_2 that is less than or equal to the power of ϕ_1 for θ greater than θ_{naught} . Note here what we are claiming is that for one sided testing problems ϕ_1 and ϕ_2 both are UMP, but in the other region they have the power higher than the other one; that means, like ϕ_1 is UMP for θ less than or equal to θ_{naught} . So, for θ greater than or equal to θ_{naught} the ϕ_2 is UMP.

So, this one is having power less than that and similarly the other way round. So, naturally no test is UMP. So, no test can be UMP for $H_0: \theta = \theta_{naught}$ against $H_1: \theta = \theta_{naught}$; let me call it star here. So, what we have concluded here is that although for one sided testing problems and for some of the two sided testing problems UMP test exist. There are certain two sided testing problems where, the UMP test does not exist and that I will be developing in the next lecture.