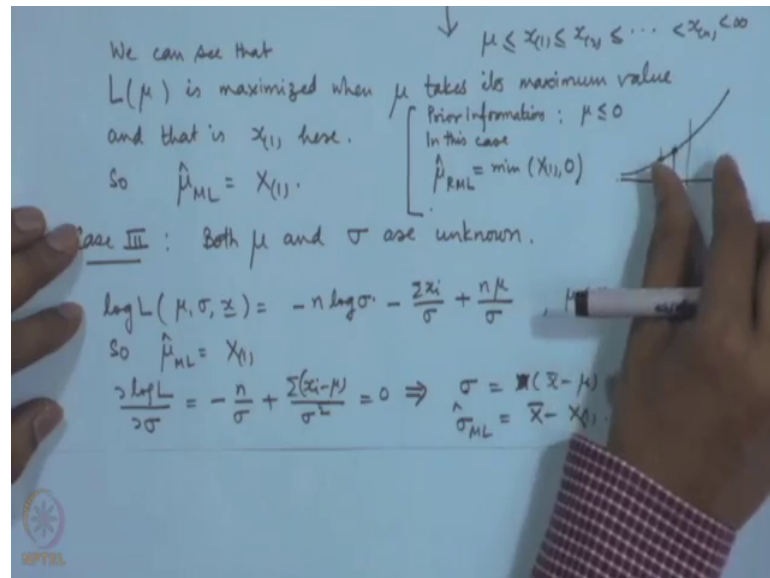


Statistical Inference
Prof. Somesh Kumar
Department of Mathematics
Indian Institute of Technology, Kharagpur

Lecture – 12
Finding Estimators - VI

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Now, let us take the more important case when both the parameters μ and σ are unknown. Now, let us go back to the original likelihood function, it was $\frac{1}{\sigma^n} \prod_{i=1}^n \frac{1}{\sigma} e^{-\frac{x_i - \mu}{\sigma}}$. So, we consider the log of this, so log of likelihood function that is equal to $-n \log \sigma - \sum_{i=1}^n \frac{x_i - \mu}{\sigma}$.

Now, you can see here the role of μ is quite different. And when we consider the maximization with respect to μ , it will be attained at the maximum value of μ . So, we can easily then see that as before the maximum value that it can take is so μ head ML will remain to be $X_{(1)}$. However, for maximization with respect to σ , we can apply the usual calculus here. So, you can consider derivative with respect to σ that will be equal to $-\frac{n}{\sigma} + \frac{\sum(x_i - \mu)}{\sigma^2}$ so, we may actually put it together because this was this term. Now, this is equal to 0, if you put this, you get σ is equal to $\sqrt{\frac{\sum(x_i - \mu)}{n}}$.

times \bar{x} divided by n . So, this n gets cancelled out. So, the maximum likelihood estimator for σ will be obtained by simply replacing μ by $\hat{\mu}$ ML.

So, $\hat{\sigma}$ ML is equal to $\bar{x} - s$. Now, you can see here the effect of partial information and the effect of no information. When the partial information about the parameters was there, then in the case of the estimator of σ we got \bar{x} , but now you see it is changed to $\bar{x} - s$. Whereas, the effect on the estimation of μ is not there for when σ was known or σ is unknown, the estimation of μ is still the same.

Now, in this case I will also consider some special cases. Here let us consider when σ was known, suppose we have additional prior information, suppose additional prior information about μ is there in the form say $\mu \leq 0$. Basically it means that the minimum guarantee time is upper bounded by some number say μ_0 which we have brought down to 0. Now, in this case what will happen? If we look at the form of the likelihood function, this function is an increasing function; this function is an increasing function of μ it starts from minus infinity that means, it is 0 and then at 0, it will be to the power something and then thereafter.

Now, if you see if $\bar{x} > 0$ is here, then the maximum is occurring at this point. Whereas, if $\bar{x} < 0$ is here with respect to 0, then the maximum is occurring here. Then in this case $\hat{\mu}$ ML which I will call restricted ML, this will become minimum of \bar{x} and 0. So, the role of prior information is important here. You consider a second situation suppose in place of $\mu \leq 0$, we had $\mu \geq 0$ or greater than or equal to 0, in that case there will be no change because \bar{x} is greater than or equal to μ it will remain greater than 0. So, the maximum occurrence at \bar{x} which is within the zone. So, there will not be any change in the maximum likelihood estimator when I am considering the prior information $\mu \geq 0$.

So, you can actually see that the role of the prior information is different in different situations and this is the beauty of the maximum likelihood procedure that it takes care of each situation individually. So, this is totally based on the likelihood function.

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4. Laplace or Double Exponential distⁿ.

Let X_1, \dots, X_n be a random sample from double exponential distribution with pdf.

$$f(x, \mu, \sigma) = \frac{1}{2\sigma} e^{-\frac{|x-\mu|}{\sigma}}, \quad x \in \mathbb{R}, \mu \in \mathbb{R}, \sigma > 0.$$

Case: μ is known, say $\mu = 0$.

$$L(\sigma, \mathbf{x}) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i|}{\sigma}}$$

$$\ell(\sigma) = \log L(\sigma, \mathbf{x}) = -n \log 2 - n \log \sigma - \frac{\sum |x_i|}{\sigma}$$

$$\frac{d\ell}{d\sigma} = -\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2} = \frac{-\sum |x_i| + n\sigma}{\sigma^2} > 0 \quad \text{if } \sigma < \frac{1}{n} \sum |x_i|$$

$$< 0 \quad \text{if } \sigma > \frac{1}{n} \sum |x_i|$$

So $\ell(\sigma)$ attains its maximum at $\frac{1}{n} \sum |x_i|$

Now, let us consider another important distribution which is known as Laplace or double exponential distribution. So, let X_1, X_2, \dots, X_n be a random sample from double exponential distribution with the probability density function. Here x is any real number; the parameter μ is any real number and σ is a positive parameter. As before we may have different situations like μ may be known, so we may put it to be 0; when σ may be known and we may put it to be 1 etcetera.

So, let us consider the case when say μ is known say μ is equal to 0. So, in this case the likelihood function is $\frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i|}{\sigma}}$. So, the log likelihood function that is equal to $-n \log 2 - n \log \sigma - \frac{\sum |x_i|}{\sigma}$. So, if we consider $\frac{d\ell}{d\sigma}$ that is equal to $-\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2}$.

If you put this equal to 0, of course, you can adjust the term this is equal to $\frac{\sum |x_i|}{\sigma} - n$ as you can easily see that it is greater than 0. If σ is greater than $\frac{1}{n} \sum |x_i|$ and it is less than 0 if σ is less than $\frac{1}{n} \sum |x_i|$. So, the maximum occurs at $\frac{1}{n} \sum |x_i|$. So, $\ell(\sigma)$ attains its maximum at $\frac{1}{n} \sum |x_i|$.

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Case: μ is known, say $\mu=0$.

$$L(\sigma, \mathbf{z}) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i|}{\sigma}}$$

$$\ell(\sigma) = \log L(\sigma, \mathbf{z}) = -n \log 2 - n \log \sigma - \frac{\sum |x_i|}{\sigma}$$

$$\frac{d\ell}{d\sigma} = -\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2} = \frac{-\sum |x_i| + n\sigma}{\sigma^2}$$

> 0 if $\sigma < \frac{1}{n} \sum |x_i|$
 < 0 if $\sigma > \frac{1}{n} \sum |x_i|$

So $\ell(\sigma)$ attains its maximum at $\frac{1}{n} \sum |x_i|$
 So $\hat{\sigma}_{ML} = \frac{1}{n} \sum |x_i|$.

So, the maximum likelihood estimator of equal to 1 by n sigma modulus x i.

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Case II: σ is known, say $\sigma=1$ (WLOG).

$$L(\mu, \mathbf{z}) = \frac{1}{2^n} e^{-\sum |x_i - \mu|}$$

L is maximized with respect to μ when $\sum_{i=1}^n |x_i - \mu|$ is minimized.
 We can show that $\sum |x_i - \mu| = S$ is minimized when μ is a median of x_1, \dots, x_n .

Write $S = \sum_{i=1}^n |x_{(i)} - \mu|$, $x_{(i)}$'s are ordered values of x_1, \dots, x_n .

Case (i): let n be odd i.e. $n = 2k+1$.

$$S = |x_{(1)} - \mu| + |x_{(2)} - \mu| + \dots + |x_{(2k)} - \mu| + |x_{(2k+1)} - \mu|$$

$$= (|x_{(1)} - \mu| + |x_{(2k+1)} - \mu|) + (|x_{(2)} - \mu| + |x_{(2k)} - \mu|) + \dots + (|x_{(k)} - \mu| + |x_{(k+1)} - \mu|)$$

Let us take the second case when sigma is known. And once again since sigma is a scale parameter, we may take it to be 1 without loss of generality. In this case the likelihood function is equal to 1 by 2 to the power n e to the power minus sigma modulus x i minus mu. Now, you see here, this will be maximized with respect to mu if sigma of modulus x i minus mu is minimized. L is maximized with respect to mu, when sigma of modulus x i minus mu is minimized.

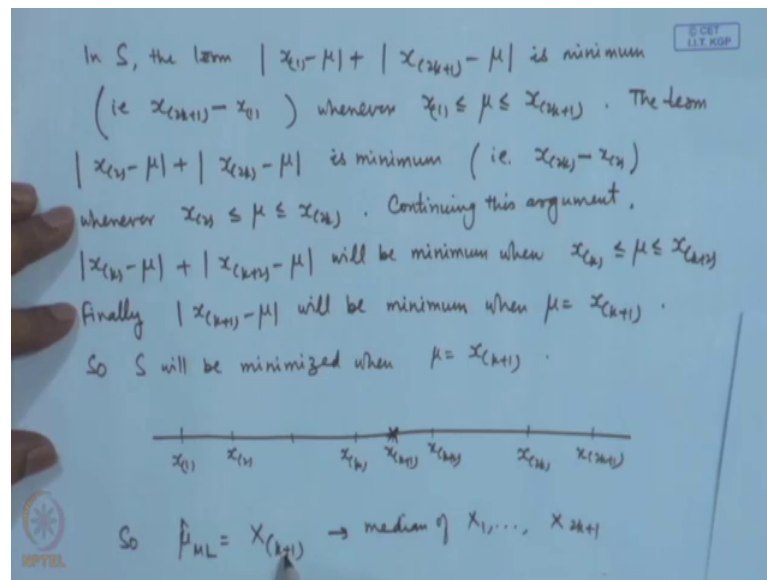
Now, one can show that this is minimized when μ is the median of the observations, because, this modulus term is coming, therefore, you cannot use the usual differentiation procedure here. However, we can give a direct argument, we can show here that sigma modulus of x_i minus μ let me call it S is minimized when μ is a median of x_1, x_2, \dots, x_n .

Let me consider two cases. So, we write S as a sigma and in place of the x_i 's we consider the ordered x_i 's ordered value of x_1, x_2, \dots, x_n , that means, x_1 is the minimum x_2 is a second minimum and so on as before. Now, we give argument in two cases let us take say n that is say n is equal to something like $2k + 1$. Now, this some S , we place like this $x_1 - \mu + x_2 - \mu + \dots + x_{2k} - \mu + x_{2k+1} - \mu$.

This we express as say $x_1 - \mu + x_{2k+1} - \mu$ that means, I have taken the first term and the last term. Then I take the second term and the second last term $x_2 - \mu$ and $x_{2k} - \mu$ and so on that is finally, we will have $x_k - \mu$ minus $x_{k+2} - \mu$. And the last term then will be remaining that is $x_{k+1} - \mu$.

What we do, we look at the minimization of each of these terms which are clubbed together. So, if you look at these two, here it is the x_1 and this is x_{2k+1} . If I consider μ to be any value between these two, then this will turn out to be $x_{2k+1} - x_1$ that will be the minimum value.

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So, let us write the complete argument here. In S , the term $x_{(1)} - \mu$ plus $x_{(2k+1)} - \mu$ is minimum that is the value will be $x_{(2k+1)} - x_{(1)}$, whenever I choose μ to be a number between $x_{(1)}$ and $x_{(2k+1)}$. Similarly, the term $x_{(2)} - \mu$ plus $x_{(2k)} - \mu$; this is minimum and of course, the minimum value will be $x_{(2k)} - x_{(2)}$, whenever $x_{(2)}$ is less than or equal to μ less than or equal to $x_{(2k)}$. So, in that way, if you look at all the sums, they will be minimum, whenever μ lies between the two values which are involved in those two terms.

So, if we continue this argument, the term $x_{(k)} - \mu$ plus $x_{(k+1)} - \mu$ will be minimum, when $x_{(k)}$ is less than or equal to μ less than or equal to $x_{(k+1)}$. Finally, $x_{(k+1)} - \mu$ will be minimum, when μ is equal to $x_{(k+1)}$. We have considered the term by term minimization of this S . So, we have taken these this and this together, then these two together and so on. We have derived the condition for the minimization of each of this.

Now, therefore, the overall minimum will be attained if all the conditions are simultaneously satisfied. Now, if you see all the conditions to be simultaneously satisfied, what will be the condition. This is the widest interval because this is from minimum to the maximum. This interval is the second and so on. So, if I look at this scale here $x_{(1)}, x_{(2)}, x_{(k)}, x_{(k+1)}, x_{(k+2)}, x_{(2k)}, x_{(2k+1)}$ somewhere you have $x_{(k)}$ and $x_{(k+1)}$ and $x_{(k+2)}$. So, from the 1st one, μ should be any value between these two; from the 2nd one, μ

should be any value between these two; from the 3rd one and so on and finally you are getting the value that is $x_k + 1$.

So, if μ is $x_k + 1$, each of these terms that I have clubbed together, they will be minimum. Therefore, overall S will be minimized. So, S will be minimized when μ is equal to $x_k + 1$, because this will satisfy all the conditions. So, we conclude that μ_{ML} is equal to $x_k + 1$ that is actually the median of $x_1, x_2, \dots, x_{2k+1}$ because when the number of observations is odd, the middle value will be the median here.

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Case (ii) n is even say $n = 2k$.

$$S = (|x_{(1)} - \mu| + |x_{(2k)} - \mu|) + (|x_{(k)} - \mu| + |x_{(k+1)} - \mu|) + \dots + (|x_{(k)} - \mu| + |x_{(k+1)} - \mu|)$$

Arguing as before, S will be minimum when $x_{(k)} \leq \mu \leq x_{(k+1)}$ i.e. μ is a median of x_1, \dots, x_{2k} . We may take it to be $(x_{(k)} + x_{(k+1)})/2$.

So we have $\hat{\mu}_{ML} = \text{Med}(x_1, \dots, x_n) = M$.

Case III: Both μ and σ are unknown.

$$L(\mu, \sigma, \mathcal{X}) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i - \mu|}{\sigma}}$$

Now, let us consider the case when n may be even; n is even say n is equal to $2k$. Now, in this case once again we may consider the clubbing in the similar fashion. However, this last term will not be there. Therefore, we will write the clubbing in this fashion $x_1 - \mu + x_2 - \mu + \dots + x_{2k} - \mu$ and so on. In the final, it will be $x_n - \mu + \dots + x_k - \mu + x_{k+1} - \mu$.

So, if we give the argument as before; arguing as before, S will be minimum when x_m is less than or equal to μ less than or equal to x_{m+1} . Because, now on a scale $x_1, x_2, \dots, x_n, x_{n+1}, \dots, x_{2m-1}, x_{2m}$, so here this first sorry this will be k here. So, here if you see the first term will be minimum when μ lies between the largest interval;

the second one will be minimum when the μ lies between x_2 to x_{2k-1} and so on. The last sum will be minimized when μ lies between x_k to x_{k+1} .

Now, if μ lies between x_k to x_{k+1} when we have even number of observations that is x_1, x_2, \dots, x_{2k} any number between x_k to x_{k+1} is called a median. For convenience many times we take the average of these two value that is $x_k + x_{k+1}$ by 2. So, this we conclude that μ is a median of X_1, X_2, \dots, X_{2k} , so where we may take it to be $X_k + X_{k+1}$ by 2. So, we have μ_{ML} equal to the median of X_1, X_2, \dots, X_n . In both the cases, we are getting median; we denote it by say M .

So, now let us consider the important case. When both the parameters may be unknown, so both μ and σ are unknown, in this case the likelihood function is equal to $\frac{1}{\sigma^n} \prod_{i=1}^n e^{-\frac{|x_i - \mu|}{\sigma}}$.

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$$l(\mu, \sigma) = \log(\mu, \sigma, \mathbf{z}) = -n \log 2 - n \log \sigma - \frac{\sum |x_i - \mu|}{\sigma}$$

$$l \text{ is maximized w.r.t } \mu \text{ when } \sum |x_i - \mu| \text{ is minimum}$$

$$\text{ie at } \mu = \text{Med}(x_1, \dots, x_n)$$

$$\text{So } \hat{\mu}_{ML} = M$$

$$\frac{\partial l}{\partial \sigma} = -\frac{n}{\sigma} + \frac{\sum |x_i - \mu|}{\sigma^2} \text{ is attaining the maximum}$$

$$\text{value at } \sigma = \frac{1}{n} \sum |x_i - \mu|$$

$$\text{So } \hat{\sigma}_{ML} = \frac{1}{n} \sum |x_i - M| \quad (\text{mean deviation about median})$$

So, we take the log here that is equal to minus $n \log 2$ minus $n \log \sigma$ minus $\sum |x_i - \mu|$ by σ . So, as before the maximization with respect to μ will occur when $\sum |x_i - \mu|$ is minimum. And we have already shown that this is occurring when μ is a median. So, l is maximized with respect to μ , when $\sum |x_i - \mu|$ is minimized that is at μ equal to median of X_1, X_2, \dots, X_n . So, μ_{ML} is equal to the median which we are calling M .

Now, you look at the solution for sigma. If we consider the derivative of l with respect to sigma, we get $-\frac{n}{\sigma} + \frac{\sum x_i - n\mu}{\sigma^2}$. And as before if we argue this is attaining the maximum value at sigma is equal to $\frac{1}{n} \sum x_i - \mu$. Now, we have already obtained the solution for mu. If we substitute it here you get the maximum value of the maximized value of likelihood function for sigma equal to $\frac{1}{n} \sum x_i - \mu$. So, sigma head ML is equal to $\frac{1}{n} \sum X_i - \bar{M}$ which is nothing but the mean deviation about median mean deviation about median.

So, today friends we have discussed various probability models and we have discussed the maximum likelihood estimators for those models, I have tried to cover various cases here. And another thing is that we will take up some different cases where either the maximum likelihood estimator is not unique, it may not exist. And then we will consider the large sample properties of the maximum likelihood estimators in the next class.