

Applied Multivariate Analysis

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Lecture No. # 02

Basic concepts on Multivariate Distribution – II

Let us start the day with some definitions. Suppose we have x , a p dimensional random vector such that, we have expectation of x as the mean vector μ and the covariance matrix of x to be denoted by Σ . We define the characteristic function.

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$X \sim N(\mu, \Sigma) \Rightarrow E(X) = \mu, \text{Cov}(X) = \Sigma$

Characteristic function of X

$$\phi_X(t) = E(\exp(i t' X)) \quad ; \quad i = \sqrt{-1}$$

$t \in \mathbb{R}^p$

Defⁿ: $\text{Cov}(X) = \Sigma$

(a) Total variation in X : $E \text{tr} \Sigma$

(b) Generalized variance of X : $|\Sigma|$

Note: $E \text{tr} \Sigma, |\Sigma|$ can't replace Σ

e.g. $\Sigma^1 = \begin{pmatrix} 2 & 1 \\ 1 & 3 \end{pmatrix}_{2 \times 2}; \Sigma^2 = \begin{pmatrix} 2 & -1 \\ -1 & 3 \end{pmatrix}_{2 \times 2}$

$$E \text{tr} \Sigma^1 = E \text{tr} \Sigma^2 = 5$$
$$|\Sigma^1| = |\Sigma^2| = 2 \times 3 - 1 = 5$$

Characteristic function of x is defined as $\phi_X(t)$ to be expectation of $e^{it'x}$ where we have i to be square root of minus 1 and t , a vector belonging to \mathbb{R}^p to the power p right. Now, this is the joint characteristic function of the components of x . So, it is the joint characteristic function for the random vector x . Now given the information about this particular joint characteristic function of the elements of x , one can actually get to the marginal characteristic functions of the respective components, that make up this particular random vector x ; similar to what we had seen in the last lecture for the moment generating function.

Now, this characteristic function of course completely determines a multivariate distribution. And the knowledge of this characteristic function about the multivariate random vector completely determines also the characteristic function and hence, the distribution of the respective components of x . Sometimes a quantity which is used to condense the information about, which is present in this covariance matrix σ ; other two following quantities just put it as definitions. So, suppose the covariance matrix of x is σ ; it is a p by p matrix. So, there are unknown elements which are present in this p by p matrix; it is a symmetric matrix.

So, the total number of unknown quantities in this particular matrix is p in to p plus 1 by 2. Now, in order to summarize this **the** information, that is present in the covariance matrix. The two following quantities are defined. The first is the total variance or total variation in x . So, the total variation in x is given by trace of the σ matrix. So, it condenses the information that we get through the covariance matrix in terms of single quantity, which is given through the trace of this particular matrix. The second quantity which is of interest, what is referred to as a generalized variance?

Generalized variance of this random vector x and that is given by determinant of σ . So, once again it summarizes or rather condenses the information, that is present in this covariance matrix σ in terms of two single quantities either trace of σ or determinant of σ . The first one is called the total variation in x . As we will see later on, then this total variation in x is the quantity of interest and which actually is looked upon as preserving type of approach, when we look at principle components analysis. Now, it should be noted that, the two quantities that we have defined just now.

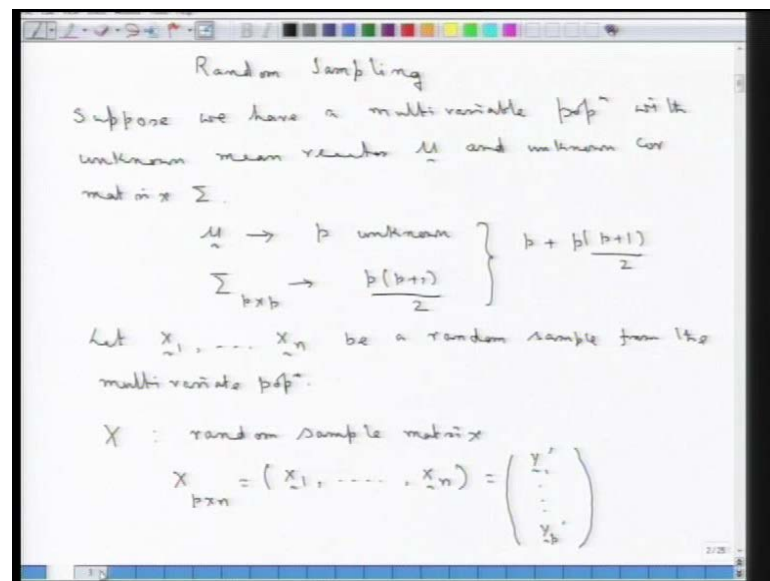
The trace of σ , which is the total variation in x and the generalized variance of x are not sufficient to replace this p by p matrix σ . So, just put it as a note that, this trace of σ and determinant of σ cannot actually replace this σ . It is obvious actually, if you look at this determinant and then, the corresponding condensation in terms of trace of σ and determinant of σ . Say for example, if you have a σ matrix, suppose σ is a 2 by 2 matrix which has elements say 2 3 on the diagonal; 1 1 on the off diagonal.

Suppose this is the first σ matrix covariance matrix; we can have another σ matrix, which is of 2 by 2 matrix; which has entries as 2 3, the same entries on the diagonal; minus 1 and **plus** 1 on the two off diagonal. So, these are two different

covariance matrices. However, if we look at the total variation that is, it is given through trace of sigma. Trace of sigma 1 will be equal to trace of sigma 2 and that is equal to the sum of the two diagonal elements that is 5. We can look at the determinant of the two covariance matrices sigma 1 and sigma 2. So, these two quantities would be 2 into 3 minus 1. So, that is equal to 5 as well **right**.

So, the total variation that is there in this sigma 1 and sigma 2 are both equal to 5. The generalized variance of the underlined random vector x, which is given through this sigma 1 and sigma 2; both of them are also equal to 5. So, they have both of these two covariance matrices have the same measures as far as total variation and the generalized variance of the associated random vector is concerned. However, the two covariance matrices themselves are quite or clearly different **right**. So, these two quantities are not enough actually to replace the covariance matrix as such. However, the condensation that one gets is of interest at times.

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Now, let us move in to an important concept, which is random sampling from multivariate distributions or rather random sampling from multivariate distributions. Now, random sampling from multivariate distribution; **what is** what is basically the need of this random sampling? **Suppose we have a multivariate multivariable population** Suppose we have a multivariable population with unknown mean vector as mu, which is p by 1 and unknown covariance matrix as sigma. Since for all practical purposes, when

we considering a multivariate population these two quantities; μ , the mean vector and the covariance matrix σ are unknown.

So, the numbers of unknown quantities, the parameters which are actually present in that particular population are the following. So, this mean vector will lead us to p unknown quantities, because it is a p dimensional vector. So, p unknown quantities and when we look at the σ matrix, which is p by p matrix. This is a symmetric matrix; it is a covariance matrix. Hence, it is a symmetric matrix and hence, the total number of unknown quantities which are present in this σ matrix are p into p plus 1 by 2 . So, the total number of unknown quantities which are present in this μ vector and the covariance matrix are p plus p into p plus 1 by 2 right.

So, in order to have inference about these unknown quantities and to have some estimators corresponding to the mean vector μ and the covariance matrix σ , what is done is to go for random sampling from that particular population. Now, let us denote by the following quantities. Let x_1 vector, x_2 vector, x_n vector be a random sample from the multivariate population. So, if we have that, then on the basis of this n random samples, now each of these x_i 's; x_i vectors are p by 1 ; because that is basically random sample drawn from the p variate, a multivariate population. And hence in order to have the inference about the mean vector μ and the covariance matrix σ , we will be using this particular set of random sample.

Now, let us denote by capital x , the random sample matrix. The random sample matrix, which is a p by n matrix basically comprises of the following vectors. So, where this x_1 vector, which is the first random sample to up to x_n vector, which is the n th random sample. So, we have this p by n random matrix, which is actually having this entire random sample of dimension n . Now, this can alternatively be written in terms of the following. So, let us write this as y_1 prime, y_2 prime and y_p prime, where these quantities are having the following interpretation.

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x_j : j^{th} observation vector $j = 1, \dots, n$
 $x_j = \begin{pmatrix} x_{j1} \\ \vdots \\ x_{jp} \end{pmatrix}$
 $y_i = (x_{i1}, \dots, x_{in}) \quad i = 1, \dots, p$
 \bar{x} : Sample mean vector
 $\bar{x} = \begin{pmatrix} \bar{x}_1 \\ \vdots \\ \bar{x}_p \end{pmatrix} = \begin{pmatrix} \frac{1}{n} \sum_{j=1}^n x_{j1} \\ \vdots \\ \frac{1}{n} \sum_{j=1}^n x_{jp} \end{pmatrix} = \frac{1}{n} \begin{pmatrix} y_1' \mathbf{1}_n \\ \vdots \\ y_p' \mathbf{1}_n \end{pmatrix} = \frac{1}{n} X \mathbf{1}_n$
 where $\mathbf{1}_n = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix}$

So, we have this x_j vector is basically the j th observation vector. We have n such observations. So, j equal to 1 to up to n , where this x_j vector is given by the following, that this is say x_{1j} , x_{pj} . So, each of these are p dimensional, because that is j th observation vector corresponding to this p dimensional **vector** random vector. Similarly, this y_i vector that we had defined here, that this is y_{1i} , y_{2i} , y_{ni} . This y_{i1} is what is holding the observations corresponding to the i th variable. So, these are basically y_{i1} . So, corresponding to the i th row of that matrix.

So, this would be x_{i1} , x_{i2} , x_{in} ; where this i now is from 1 to up to p . So, this basically is the structure of data matrix; where this y_{i1} is basically the i th row of the x matrix and x_j is the j th column of the x matrix **right**. In terms of these vectors, that we have defined through the random sampling; we will introduce the two important quantities, which is the random mean vector or sample mean vector and the sample variance covariance matrix, which would be used actually in order to have inference or rather the inference about the unknown mean vector μ and the unknown covariance matrix, that is σ **right**.

So, we will have this being defined, which is the \bar{x} matrix or \bar{x} vector rather. This is the sample mean vector. **well** In order to have the same notation as this one, what we will denote this by capital x_i 's. So, let us be consistent with this definition. So, you will have a capital x_i 's to denote the corresponding random variables and we will use small x

i 's to denote the observations. So, \bar{x} capital x bar vector is the sample mean vector and it is given by the following. So, this is a p dimensional vector, which holds \bar{x}_1, \bar{x}_2 and the p th variables mean \bar{x}_p .

So, this would be given by $\frac{1}{n}$ then the summation of all the observations corresponding to the first variable. So, that is $\bar{x}_1 = \frac{1}{n} \sum_{j=1}^n x_{1j}$ say, this j is equal to 1 to up to n and the last element is $\frac{1}{n} \sum_{j=1}^n x_{pj}$ right. So, these are the corresponding sample means of the respective variables in the x vector component. Now, in terms of this observation vector, this can be written in the following way. So, you can take $\frac{1}{n}$ outside and then, what we have here is? This is basically $\frac{1}{n} Y'$ multiplied by I which is of dimension $n \times 1$.

So, it is a column vector of dimension n with 1 as all the entries and the last one is, what is corresponding to \bar{x}_p all the variables for the p th variable; all the observation corresponding to the p th variable, where this will just complete this particular stuff here. So that, this in terms of the random matrix, what we have is $\frac{1}{n} Y' I$; where this I is an $n \times 1$ column vector, which has 1 as all its entries. So, this \bar{x} which is the sample mean vector; it is a random vector once again; because it is comprising of the random variables that we have taken. So, the entries of that particular x vector are given by these, which in terms of the observation vectors that we had defined earlier, the p observation vectors.

So, these basically are these observation vector, y_i primes and that random sample mean vector in terms of the random matrix, that the data matrix that we have defined. It is $\frac{1}{n} Y' I$. Now, let us move on to looking at now this \bar{x} vector, which is the sample mean vector is what is going to be used for inference that is based; that is inference about the population mean vector, which is unknown; which is the μ quantity. Now, in order to look at the other unknown quantity which is present; which is σ^2 the variance covariance matrix.

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Sample variance covariance matrix

$$S_n = \frac{1}{n} \begin{pmatrix} \sum_{j=1}^n (x_{1j} - \bar{x}_1)^2 & \sum_{j=1}^n (x_{1j} - \bar{x}_1)(x_{2j} - \bar{x}_2) & \dots & \sum_{j=1}^n (x_{1j} - \bar{x}_1)(x_{pj} - \bar{x}_p) \\ \sum_{j=1}^n (x_{2j} - \bar{x}_2)^2 & \dots & \dots & \dots \\ \dots & \dots & \dots & \dots \\ \sum_{j=1}^n (x_{pj} - \bar{x}_p)^2 & \dots & \dots & \dots \end{pmatrix}$$

$$= \begin{pmatrix} S_{11} & S_{12} & \dots & S_{1p} \\ S_{21} & S_{22} & \dots & S_{2p} \\ \dots & \dots & \dots & \dots \\ \dots & \dots & \dots & S_{pp} \end{pmatrix} ; n S_n = (n-1) S_{n-1}$$

We introduce the sample variance covariance matrix. So, the sample variance covariance matrix can have divisor either n similar to the univariate set up **can have a divisor as n** can have a divisor n minus 1. Suppose we define that in terms of a divisor n, then this is how the sample variance covariance matrix is going to look like; so, this sample variance covariance matrix which is going to be based once again on the random samples that we have drawn. It is a p by p random matrix, which has the following entries. So, the 1 1th element of that is going to be given by this $\sum_{j=1}^n (x_{1j} - \bar{x}_1)^2$.

Now, \bar{x}_1 is what we have already defined through this random vector here, the sample mean vector. So, that is what is going to be used here. Then, the 1 2th element is going to be given by this $\sum_{j=1}^n (x_{1j} - \bar{x}_1)(x_{2j} - \bar{x}_2)$; this into $\sum_{j=1}^n (x_{1j} - \bar{x}_1)(x_{pj} - \bar{x}_p)$ and the last entry here do not have enough space actually to write it. So, this would be given by this $\sum_{j=1}^n (x_{1j} - \bar{x}_1)(x_{pj} - \bar{x}_p)$ that multiplied by $\sum_{j=1}^n (x_{pj} - \bar{x}_p)^2$ **right**. So, this is the matrix here. Then, this is going to be a symmetric matrix. So, this we only need to write the **upper triangle** of this particular matrix.

The 2 2th element would be the one, which is corresponding to the second variable; this is $\sum_{j=1}^n (x_{2j} - \bar{x}_2)^2$. The last entry in this row would be $\sum_{j=1}^n (x_{2j} - \bar{x}_2)(x_{pj} - \bar{x}_p)$ and the last diagonal entry the p pth element of this matrix would be $\sum_{j=1}^n (x_{pj} - \bar{x}_p)^2$. Now, if you look carefully at this sample variance covariance matrix that we have listed

out here, it basically is holding say for example, this particular element; this you can associate with the variance covariance variance for the first component.

So, we can denote that by s_{11} say, then s_{12} upon n of this quantity is what is sample covariance between the first and the second variable. The last entry in the first row here is the sample covariance between the first and the p th variable. So, this would be s_{1p} ; s_{21} is just this element s_{12} itself; because this is the variance covariance matrix. This is s_{22} , s_{2p} and this would be s_{pp} element, where the s_{ij} is the ij th entry of the particular sample variance covariance matrix **right**. Now, if we define the sample variance covariance matrix through divisor $n - 1$, we will see later on that; one of them would be unbiased estimator of the population covariance matrix σ .

And the other is going to be associated with the maximum likelihood estimator, when we talk about random sampling from a multivariate normal distribution. So, one could have also define this, in terms of this **your yes** $n - 1$ quantity. So, we can say that n times s_n would be in such a situation will be given by $n - 1$ s_{n-1} **right**; where the s_{n-1} matrix would be this matrix what we have with a divisor here as $n - 1$ **right**. Now, let us try to write this particular sample variance covariance matrix, what we have defined in terms of the data matrix x that we had introduced.

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The image shows a handwritten derivation of the sample variance-covariance matrix. It starts with the definition of $(n-1)s_{n-1} = n s_n$ and expresses it as a matrix of sums of products of data points. The matrix is then simplified to $X X' - n \bar{x} \bar{x}'$, where X is the data matrix, X' is its transpose, n is the sample size, and \bar{x} is the vector of sample means. Further simplification shows $X X' - \frac{1}{n} X 1_n 1_n' X'$, where 1_n is a vector of ones. The final result is $(n-1)s_{n-1} = n s_n = X (I_n - \frac{1}{n} 1_n 1_n') X'$, with a note that I_n is the identity matrix of size $n \times n$.

So, we will have the $n - 1$ s_{n-1} is equal to n times s_n . Now, that if we look at this particular matrix here, now this element here can be written in terms of summation

following. I will just write it out here that this entry here; we can write it as $\sum_{j=1}^n x_{1j}^2$ minus n times $\sum_{j=1}^n x_{1j} x_{2j}$. So, all these entries similarly can be written like this. Say for example, the 1 2 th element here can be written as $\sum_{j=1}^n x_{1j} x_{2j}$ minus n times $\sum_{j=1}^n x_{1j}^2$. So, keeping that in mind, we can write this particular matrix in two parts. The first part will hold the sum of squares and the cross products.

So, we can right that as $\sum_{j=1}^n x_{1j}^2$; this j equal to 1 to n . The second entry, we will just have the first quantity which I said $\sum_{j=1}^n x_{1j} x_{2j}$; the last entry here is summation from j equal to 1 to n $\sum_{j=1}^n x_{1j} x_{pj}$. This is going to be $\sum_{j=1}^n x_{2j}^2$ $\sum_{j=1}^n x_{1j} x_{pj}$ and the last entry p p th element of this matrix is just the sum of squares corresponding to the p th component. So, this is the first entry, the first block actually and then, this minus n times the sum of squares entries which will be getting like here; n times $\sum_{j=1}^n x_{1j}^2$ will come here; n times $\sum_{j=1}^n x_{1j} x_{2j}$ will come here and similarly, the other entries will follow.

So, this can be written in terms of $\sum_{j=1}^n x_{1j}^2$. The second entry will be $\sum_{j=1}^n x_{1j} x_{2j}$. The last entry in this first row would be $\sum_{j=1}^n x_{1j} x_{pj}$. Then, the last diagonal entry would be $\sum_{j=1}^n x_{pj}^2$ **right**. So, once we have written it in this particular form, it is easy to realize that the first matrix that we have written here in terms of the data matrix; that we had introduced that x . It is just $x x^T$ **right**. The x matrix we had defined here which was **well** we can actually we had written this as this matrix here. Let me see **yeah** this particular matrix here, this basically is the x matrix.

So, if we write the corresponding entries here x 1 vector, the first component; so, it will have the p components out there. Then, this x n will have once again the p components of that n th observation vector. And then, this matrix p by p would actually lead us to the product $x x^T$, which would now be having the diagonal entries as the sum of the squares of the respective components and the off diagonal entries will hold the products. Now, this particular matrix block here can similarly be written in terms of the x bar vector. So, this is \bar{x} vector that multiplied by this \bar{x} vectors transpose.

What is \bar{x} vector? \bar{x} vector is what we have defined here as the respective x bar components for the p entries **(())** p elements. Now, **we** from this particular form, it is sometimes useful to reduce it to this particular form. Now, this is not yet in terms of

entirely the data matrix. Now, as we had seen that, this \bar{x} vector can be written in terms of the data matrix as \bar{x} vector equal to $\frac{1}{n} \mathbf{1}_n' X$. So, we can use that same thing out here. So, it is $\frac{1}{n} \mathbf{1}_n' X$ and then, the transpose of that particular vector. So, this can be written as $X' \frac{1}{n} \mathbf{1}_n$; this one will get canceled out.

So, will have **one** $\frac{1}{n} \mathbf{1}_n'$ remaining and then, this is $\frac{1}{n} X' \mathbf{1}_n$, then the transpose of this quantity. So, what will be having is $\frac{1}{n} \mathbf{1}_n' X X' \frac{1}{n} \mathbf{1}_n$; this multiplied by $X' X$. So, this particular form what we have is what is now expressing this **the** sample variance covariance matrix with either the divisor n or the divisor $n - 1$ in terms of the observed data matrix. Now, here \mathbf{I}_n of course is an n by n identity matrix. So, this is what we get. Now, an alternate form of this sample variance covariance matrix is also sometimes useful.

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$$(n-1) S_{n-1} = n S_n = X'X - n \bar{x} \bar{x}'$$

$$= \sum_{j=1}^n x_j x_j' - n \bar{x} \bar{x}'$$

$$= \sum_{j=1}^n (x_j - \bar{x})(x_j - \bar{x})'$$

Sample Correlation matrix

$$R = (D^{1/2})^{-1} S (D^{1/2})^{-1}$$

$$D = \text{diag}(s_{11}, \dots, s_{pp})$$

Geometric Interpretation
 Let x_1, \dots, x_n be n obsn vectors
 & y_1, \dots, y_p obsn corresp to p variables
 (i) Projection of y_i on $\mathbf{1}_n$ $i = 1, \dots, p$

$$\left(\frac{y_i \mathbf{1}_n'}{\mathbf{1}_n' \mathbf{1}_n} \right) \mathbf{1}_n = \bar{x}_i \mathbf{1}_n = \begin{pmatrix} \bar{x}_i \\ \vdots \\ \bar{x}_i \end{pmatrix}$$

I will just write that, say $n - 1$ s_{n-1} n times s_n that what we have derived as $X'X$ transpose into \bar{x} transpose. So, since this X vector, X matrix rather; X matrix is of this particular form $X'X$ transpose can be given by the following that, this is summation j equal to 1 to n $x_j x_j'$ transpose minus n times $\bar{x} \bar{x}'$ transpose. So, this can be written as following $X'X$ minus \bar{x} vector into $X'X$ minus \bar{x} transpose j equal to 1 to n **right**. So, this form also we will today itself use this form, in order to derive an unbiased

estimator for the population covariance matrix, that is σ . Now, similar to the population correlation matrix, one can also define the sample correlation matrix.

Say, the sample correlation matrix **say** denoted by R would be $\frac{1}{n} D^{-1/2} S D^{-1/2}$, where S is either with the divisor n or $n - 1$; this into $D^{-1/2}$ to the power minus 1; where this D matrix is the diagonal matrix which is holding the sample variances of the respective components p in number **right**. So, through this diagonal matrix, if we look at pre and post multiplication using $D^{-1/2}$; So, what will be getting is the sample correlation matrix; where the i, j th element will actually be the sample correlation random variable for the x_i and x_j component of this **right**. Now, let us look at a bit of geometric interpretation for this random sampling.

So, when we talk about this geometric interpretation, we will look at too simple interpretations. Suppose we look at these as observation vectors, let x_1, x_2, \dots, x_n be n observation vectors. Now, from these observation vectors we had seen that, one can also write this in terms of y_i 's. So, this y_1 vector, y_2 vector, y_p vector; this would be observations corresponding to p variables. Now, if we look at the **projection of any y_i on one** projection of say any y_i vector on I_n is going to be given by $y_i \cdot \frac{1}{\|I_n\|} I_n$ divided by $\|I_n\|$; this multiplied by this I_n vector and what is this equal to?

This is the projection **this is the projection** vector, which is the projection of y_i on I_n . So, there are p such vectors, i equal to 1 to up to p . Corresponding to each of these p vectors, which now are holding all the observations corresponding to these variables. So, this is what is going to give us the sum of all the entries, which we have corresponding to the i th variable and $\frac{1}{\|I_n\|}$; this is $\frac{1}{\|I_n\|} I_n$ just drop this I_n from all these places. So, later on actually without loss of any generality will just drop this particular I_n index will say that, it is a vector which is holding 1's and its **(())** particular dimensions conforming to the other vector.

So, this is just the sum of all these observations and this is going to give us n . So, this is what we will be having is $\bar{x}_i \cdot I_n$. So, this basically is the vector, which is of dimension n and has entries \bar{x}_i at all the places. So, that is basically the projection. So, this is the mean **mean** of the i th variable and this is the n dimensional vector and that is what has got the interpretation that, it is just a projection of the y_i vector. The i th

vector **correspond the vector** corresponding to the i th variable are holding all the n observations on this $I \times n$ vector and that is what is this.

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(ii) Deviation vectors

$$\underline{d}_i = y_i - \bar{x}_i \mathbf{1} \quad i = 1(1)n$$

$$= (x_{i1} - \bar{x}_i, \dots, x_{in} - \bar{x}_i)'$$

$$\underline{d}_i' \underline{d}_i = \sum_{j=1}^n (x_{ij} - \bar{x}_i)^2 = (n-1) s_{ii} \quad i = 1(1)n$$

also $\underline{d}_i' \underline{d}_k = \sum_{j=1}^n (x_{ij} - \bar{x}_i)(x_{kj} - \bar{x}_k)$

$$= (n-1) s_{ik}$$

Let θ_{ik} be the angle betⁿ \underline{d}_i & \underline{d}_k , then

$$\cos \theta_{ik} = \frac{\underline{d}_i' \underline{d}_k}{[\underline{d}_i' \underline{d}_i]^{1/2} [\underline{d}_k' \underline{d}_k]^{1/2}}$$

$$= \frac{(n-1) s_{ik}}{[(n-1) s_{ii}]^{1/2} [(n-1) s_{kk}]^{1/2}}$$

Now, let us also look at the following deviation vectors. Let us denote by say d_i , the deviation vector. Now, the deviation vector is the deviation, which we are going to define as y_i ; this minus x_i bar times $\mathbf{1}$ **right**. So, this is what is going to hold the following quantities that, we have x_{i1} minus x_i bar; this is x_{in} , the n th observation minus x_i bar. So, this deviation vector is what is going to give us in each of the components, the deviation of the respective observations corresponding to that i th variable.

This is the first observation in the i th variable that minus the mean corresponding to all the n observation of that i th variable. So, **these are** this is the deviation of the first observation **from its mean** from the mean corresponding to that variable and likewise, we have all these n entries like this. If we look at the square of the norm of this deviation vector; if we just look at this d_i prime d_i , what is that going to give us? That is going to give us the sum of squares of these deviation quantities. So, this is going to be given by x_{ij} minus x_i bar square this j equal to 1 to n . So, what is this quantity? This quantity is n minus 1 times s_{ii} .

So, this is actually if we have the deviation vectors, d_i is being denoted by this deviation as we have discussed that, these are the deviations basically from the respective mean

components. Then, d_i' is nothing but the sum of squares of these entries in the deviation vector, which is associated with $n - 1$ times s_{ii} ; where s_{ii} is the sample observed variance components corresponding to the i th components. So, we have this deviation vectors p in numbers and hence, this also will be p in numbers. So, it is basically the norms square of this deviation vector are associated with the sample variances of the respective components with the multiplier $n - 1$ as the divisor.

Now, similarly also what we can actually see that this $d_i' d_k$, which is the dot product of the deviation for the i th variable and the dot product for the k th variable. This would be the cross product $\sum_{j=1}^n (x_{ij} - \bar{x}_i)(x_{kj} - \bar{x}_k)$. So, this is $\sum_{j=1}^n (x_{ij} - \bar{x}_i)(x_{kj} - \bar{x}_k)$. And this is what this quantity, which is the cross product between the deviation vector for the i th variable and k th variable. It is nothing but the covariance $n - 1$ times s_{ik} . So, where s_{ik} is 1 upon $n - 1$ of this product, which is the covariance between the i th and the k th variable.

Let us also now look at the angle between these two deviation vectors. Let us denote by θ_{ik} . Let θ_{ik} be the angle between d_i , the deviation vector for the i th variable and d_k , the deviation vector corresponding to the k th variable. Then, what we have is this cosine of this θ_{ik} is going to be given by this, which is $d_i' d_k$; this divided by $d_i' d_i$, this for the i th vector d_i ; this into $d_k' d_k$ whole raise to the power half. So, if we have these two vectors d_i and d_k , then the cosine of the angle between the two vectors two deviation vectors is given by this and what is this equal to?

As we have seen that, this $d_i' d_k$ is nothing but our $n - 1$ times s_{ik} , the covariance that divided by now this $d_i' d_i$, as we had seen out here is $n - 1$ times s_{ii} . So, this comes down to $n - 1$; this as s_{ii} . So, that is coming from this and this $d_k' d_k$ is $n - 1$ times s_{kk} . So, this is this s_{kk} whole raise to the power half. So, what we sees that, this $n - 1$ factor cancels out and what we have as the cosine of the angle between these two deviation vectors is s_{ik} divided by under root of s_{ii} into s_{kk} .

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i.e. $\cos \theta_{ik} = \frac{\Delta_{ik}}{[\Delta_{ii} \Delta_{kk}]^{1/2}} = r_{ik}$

Thus if $\theta_{ik} = 0$, then $r_{ik} = 1$
if $\theta_{ik} = \pi/2$, then $r_{ik} = 0$
if $\theta_{ik} = \pi$, then $r_{ik} = -1$

Note : (a) Sample total variation is S
(b) Sample generalized variance is $|S|$

That is, the cosine of this angle theta i k is s i k that divided by s i i in to s k k whole raise to the power half and what is that equal to? That is, just **the correlation** the sample correlation between the i th and the k th variable. So, this gives us a nice geometric interpretation about these deviation vectors, that are associated with this random sampling; that the cosine of the angle between the two is nothing but the correlation between two random variables i and k. Now, given this particular expression here we can say that, if this theta i k the angle between the two deviation vector is 0. So, they are basically in the same direction.

So, if this theta i k the angle between the two deviation vector is 0, then what we have this r i k is cosine of that angle 0 and hence, this is equal to 1. That justifies actually our intuition that, if two deviation vectors are in the same direction, then the correlation between the two random variables would be perfect linear correlation. So, that the correlation is equal to 1. Now if on the other hand, the two are orthogonal if theta i k equal to pi by two. So, we have the two deviation vectors orthogonal to one another. Then, what is the value of the correlation coefficient between the two as we would expect the two are orthogonal?

So, it is moving in orthogonal directions and hence, the correlation would just be equal to 0. In the other extreme, if they move in opposite direction not orthogonal, move perfectly in opposite direction; that is, if the angle between the two deviation vectors is pi. So, the

two vectors are moving in completely different, but perfectly opposite direction. Then, what we have this r_{ik} is what we expect that, they are exactly in the opposite perfect negative correlation. So, this gives us a nice feeling about verifying the intuition that, what happens to the deviation vectors and the angle that they are making and the corresponding values of the measure of association between the **two** any two variables.

So, this cosine θ_{ik} of course, it is for every pair i and k taken from the set of possible p variables **right**. Now you remember that, we had defined at the start that, two quantities which are associated with the covariance matrix σ , the total variation in x given through the trace of σ matrix and then, the generalized variance of x given by the determinant of this σ matrix. Similar to that, one can also define the sample quantities. So, the sample quantities would be the sample total variance. One can define the sample total variation as trace of s matrix s either with a divisor n or with a $n - 1$.

And the sample generalized variance as the determinant of s matrix and as we had argued that these of course, gives us compression of the sample variance covariance matrix. However, these are not sufficient enough to replace s . Because we can construct in a similar way to what we had seen for the σ matrix that, it is possible to have two different sample covariance matrix giving us the same total sample variation and the sample generalized variance **right**. Now, we move on to one important result regarding this estimation procedure.

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$\bar{\underline{x}} \rightarrow \underline{\mu}$
 $S_{n-1} \rightarrow \Sigma$

Result: Let $\underline{x}_1, \dots, \underline{x}_n$ be a r.s. from a multivariate popⁿ with mean vector $\underline{\mu}$ and covariance matrix Σ , then

(i) $E(\bar{\underline{x}}) = \underline{\mu}$ and $\text{Cov}(\bar{\underline{x}}) = \frac{\Sigma}{n}$
↳ (ii) $E(S_{n-1}) = \Sigma$

Pf: (i) $E(\bar{\underline{x}}) = E\left(\frac{1}{n} \sum_{i=1}^n \underline{x}_i\right) = \frac{1}{n} \sum_{i=1}^n E(\underline{x}_i)$
 $= \frac{n\underline{\mu}}{n} = \underline{\mu}$
 $\text{Cov}(\bar{\underline{x}}) = E(\bar{\underline{x}} - \underline{\mu})(\bar{\underline{x}} - \underline{\mu})'$
 $= E\left(\frac{1}{n} \sum_{i=1}^n (\underline{x}_i - \underline{\mu})\right) \left(\frac{1}{n} \sum_{i=1}^n (\underline{x}_i - \underline{\mu})\right)'$

Now, as we had seen that, **we are** we have the two following quantities, which is the $\bar{\underline{x}}$ vector; which we are going to associate with the population mean vector, which is $\underline{\mu}$ and we have the sample variance covariance matrix say with a divisor $n - 1$. Now, that is what is going to be used in order to have inference about the population variance covariance matrix that is Σ . The way of the following result, let me first state the result. So, let our x_1, x_2, \dots, x_n be a random sample **be a random sample from a** from a multivariate population with mean vector unknown as $\underline{\mu}$ and covariance matrix as Σ .

Then, the two quantities that we have defined as $\bar{\underline{x}}$ and S_{n-1} has the following properties. Number one: expectation of this $\bar{\underline{x}}$ vector is going to be equal to the mean vector $\underline{\mu}$ and covariance matrix of this $\bar{\underline{x}}$ vector is going to be Σ/n and number two: for the **covariance matrix** sample variance covariance matrix S_{n-1} , expectation of S_{n-1} only is equal to Σ . In other words, we are trying to say that this $\bar{\underline{x}}$ vector, the sample mean vector is an unbiased estimator of the population mean vector, which is $\underline{\mu}$.

The sample variance covariance matrix with a divisor $n - 1$ is an unbiased estimator of the population variance covariance matrix, that is Σ **right**. And the covariance matrix of this $\bar{\underline{x}}$ vector is going to be given by Σ/n . This gives us the feeling of generalization of the univariate result what we have to the multivariate set up. We had

ofcourse; similar result when we had univariate set up. Let us quickly look into the proof of this particular result. Now, in order to prove this one, one is quite straight forward that, expectation of this \bar{x} is expectation of the respective components.

So, its expectation of $\frac{1}{n}$ upon n \bar{x} ; now \bar{x} is going to be given by this x_i bar. So, its **all the** all these observations i equal to 1 to n . Now, when we look at expectation of this; its expectation operators comes inside and what we have is i equal to 1 to n . Expectation of these x_i components and **expect** x_1, x_2, x_n are random sample from the same multivariate population. Each having identical mean as μ and hence, expectation of each of these exercise will be equal to μ . So, what will have this as n times μ ; this divided by n and hence, that is equal to μ . So, we have this first component of this result.

We can also now look at the second result, which gives us the covariance of this \bar{x} bar vector. Which by definition of the covariance matrix of any random vector would be given by \bar{x} bar minus its expectation vector, which is μ as what we have derived; that multiplied by the transpose of that same quantity. So, what we have this is the following that, one can write this as $\frac{1}{n}$ upon n summation i equal to 1 to n . Then, we will have this x_i minus μ and then, the transpose of the same quantity, which is this i equal to 1 to n x_i minus μ transpose. Now, what is this quantity equal to? This quantity would be equal to the following.

(Refer Slide Time: 47:10)

The image shows a handwritten derivation for the covariance matrix of the sample mean vector \bar{x} . The derivation is as follows:

$$\begin{aligned} \text{Cov}(\bar{x}) &= E\left(\frac{1}{n}[(x_1 - \mu) + \dots + (x_n - \mu)]\right) \\ &\quad \left(\frac{1}{n}[(x_1 - \mu)' + \dots + (x_n - \mu)']\right) \\ &= \frac{1}{n^2} \left[\underbrace{\Sigma + \Sigma + \dots + \Sigma}_n \right] \\ &= \frac{1}{n^2} \cdot n \Sigma \\ &= \frac{\Sigma}{n} \end{aligned}$$

(ii) $(n-1)S_{n-1} = \sum_{j=1}^n (x_j - \bar{x})(x_j - \bar{x})'$

$$= \sum_{j=1}^n x_j x_j' - n \bar{x} \bar{x}'$$

$$E((n-1)S_{n-1}) = \sum_{j=1}^n E(x_j x_j') - n E(\bar{x} \bar{x}') \quad \text{--- (1)}$$

Let us just split this or rather write it term by term. So, the covariance matrix of this \bar{x} vector is going to be given by expectation. Now, what is the first element? The first element is $1/n$, as we had seen it is summation of those components. So, that it will hold these quantities. So, this is $\bar{x} - \mu$ plus the last term is n th random sample vector. Now, this multiplied by the transpose of this. Transpose of this would be given by $1/n$ and then, the transpose of the first entry, which is $\bar{x} - \mu$ transpose plus $x_n - \mu$ transpose **right**.

Now, when we look at taking expectation term by term of each of these entries with these; now, if we look at expectation of this term multiplied by this term, the first term here; what will be getting is the covariance matrix of this x , which is σ apart from this multiplier. Now, when we look at the cross product say this with the next element here; now remember that this x_1 vector, x_2 , x_n ; they are random sample and hence, they are independent. So that, we will have the expectation of the cross product of this with any of these here, except the first entry would be 0; because we have x_i **x** any x_i be independent of x_j , if i is not equal to j .

And hence, the covariance matrix between the two would be equal to 0. So that, what will be having is the following; after we take expectation $1/n^2$ will be there and then, we will have this. When this is multiplied with each of these entries and then, expectation being taken only the first element would give us σ and all the rest of the elements will be 0. When we look at the second entry here, once again the same entry here will give us the σ matrix and the rest of the entries will be zeros. So, we will have these n σ components. So, these are n in numbers.

Corresponding to this type of product **1 being** 1 with 1, 2 with 2 and n with n , all the cross product entries will be zeros; because of independence of these components \bar{x}_1 , x_1 , x_2 and x_n . So, this thus gets us down to $1/n^2$ n times σ matrix and that is nothing but, this σ by n matrix **right**. So, we have proved the first part of this particular result. Let us now move on to proving the second part of the result, which establishes actually the unbiasedness elements of this covariance matrix. So, what we had seen? this is the second part of the result. So, what we now have is, this $n-1$ s $n-1$; that is what, we had seen earlier is $x_j - \bar{x}$ into $x_j - \bar{x}$ transpose, this j equal to 1 to up to n **right**.

Now, we had also seen that, this particular term is written equivalently in the form that, this is j equal to 1 to $n \times j \times j$ prime minus n times $\bar{x} \bar{x}$ prime **right**. Now, we look at proving the results. So, if we now look at expectation of this n minus 1 s n minus 1. This is going to be given by expectation of this particular entire quantity. So, we take expectation term by term. We will have this as expectation of $x_j \times x_j$ prime, this minus n times expectation of $\bar{x} \bar{x}$ transpose. Let us give an equation number 1 to this because will be requiring this latter on. So, we need to find out those two quantities. What those expectations are? Expectation of $x_j \times x_j$ prime and expectation of $\bar{x} \bar{x}$ prime.

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The image shows a whiteboard with the following handwritten derivations:

$$\begin{aligned} \text{Cov}(x_j) &= \Sigma = E(x_j - \mu)(x_j - \mu)' \\ &= E(x_j x_j') - \mu \mu' \\ \Rightarrow E(x_j x_j') &= \Sigma + \mu \mu' \quad \text{--- (2)} \end{aligned}$$

$$\begin{aligned} \text{Cov}(\bar{x}) &= \frac{\Sigma}{n} = E(\bar{x} - \mu)(\bar{x} - \mu)' \\ &= E(\bar{x} \bar{x}') - \mu \mu' \\ \Rightarrow E(\bar{x} \bar{x}') &= \frac{\Sigma}{n} + \mu \mu' \quad \text{--- (3)} \end{aligned}$$

using (2) & (3) in (1)

$$E((n-1)S_{n-1}) = \sum_{j=1}^n (\Sigma + \mu \mu') - n \left(\frac{\Sigma}{n} + \mu \mu' \right)$$

$$=$$

So, what we realize is that **expectation** this covariance matrix of $x_j \times x_j$ is the random sample drawn from that multivariate population. So, the covariance matrix of x_j is nothing but sigma. So, that is equal to expectation of this x_j minus μ into x_j minus μ transpose. And hence, this is equal to as we had seen in the last lecture, this is nothing but $x_j \times x_j$ prime; this minus $\mu \mu$ prime. And hence, this would imply that expectation of $x_j \times x_j$ prime, which is a quantity that we would be requiring in order to evaluate that expression 1 is given by this **right**.

Furthermore, if we recall the result that we had proved in a first part, covariance matrix of \bar{x} is σ by n and that is equal to expectation of \bar{x} minus μ , its mean into \bar{x} minus μ transpose **right**. So, by the same approach what one can show is that, this

is expectation of $\bar{x} - \mu$; this minus μ is **right**. So, what we will be having is this also, that expectation of $\bar{x} - \mu$ is going to be equal to $\frac{\sigma^2}{n}$; this minus μ . So, this is one that we have going to use and this is one we are going to use.

So, we will use this 1 and 2 or rather using 2 and 3 in 1. 1 is what; this particular expression. So, what will be having is this that expectation of $\sum_{j=1}^{n-1} s_j^2$; that is summation j equal to 1 to $n-1$, then expectation of this particular quantity. So, that what will be having here is $\sum \mu^2$; this minus we have this as n times expectation of this quantity. So, that \sum this is \sum by n ; this minus μ^2 **right**. So, if one simplifies, this will cancel out.

So, will have one σ^2 from here; you will have $n-1$ σ^2 from here and then, this μ^2 term. **Just a minute** this is plus sign out here. Because if you take this expectation of $\bar{x} - \mu$ to this side, then we will have $\frac{\sigma^2}{n} + \mu^2$. So, this is the plus sign out here not a minus sign; then what will be having here is just $n-1$ times σ^2 . This would imply that, expectation of $\sum_{j=1}^{n-1} s_j^2$ is just going to be equal to σ^2 . So, this will imply that, $\sum_{j=1}^{n-1} s_j^2$ is an unbiased estimator of the population variance covariance matrix that is σ^2 .

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The image shows a whiteboard with the following handwritten text:

$$S_n \text{ is not u.e. of } \Sigma$$

$$n S_n = (n-1) S_{n-1}$$

$$E(S_{n-1} (n-1)) = E(n S_n) = (n-1) \Sigma$$

$$\Rightarrow E(S_n) = \frac{n-1}{n} \Sigma \neq \Sigma$$

Now, if on the other hand, we take $\sum_{j=1}^n s_j^2$ is not going to be an unbiased estimator of σ^2 . Why? Simply, because this n times $\sum_{j=1}^n s_j^2$ is $n-1$ times $\sum_{j=1}^{n-1} s_j^2$ and what

we have proved is that, expectation of s_{n-1} that multiplied by $n-1$. So, that would also be equal to expectation of n times s_n ; that is equal to σ . This would imply that, this was actually equal to $n-1$ times this σ . This is $n-1$ times σ . So, this would imply that, expectation of s_n is going to be equal to $n-1$ by n times σ and which is not equal to σ . So, this proves that, though s_{n-1} with the sample covariance matrix with a divisor $n-1$ is an unbiased estimator of σ . However, this s_n is not an unbiased estimator of σ . However, as n goes to infinity, this will be an unbiased system estimator and hence, this s_n will be an unbiased estimator in the limit.