

Applied Multivariate Analysis

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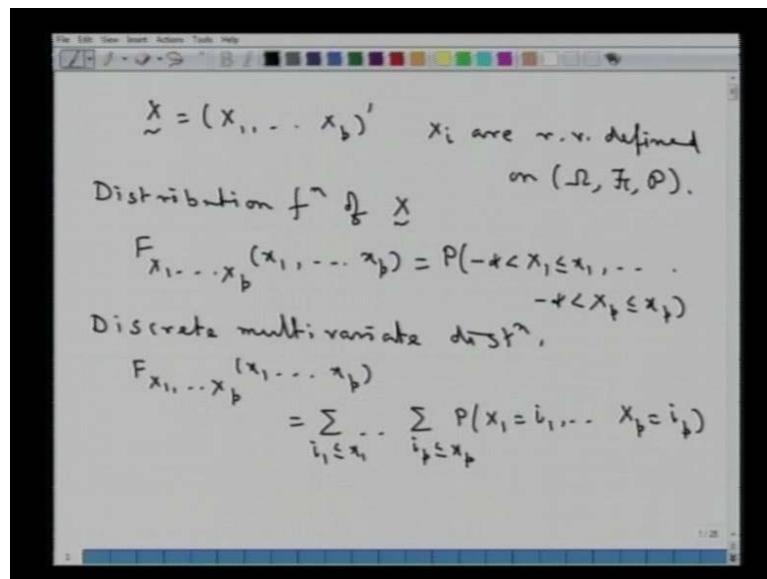
Indian Institute of Technology, Kanpur

Lecture No. # 01

Basic concepts on multivariate distribution – I

We will be starting with some basics concepts on multivariate analysis. So, let us define what we called by a multivariate random vector.

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$\tilde{X} = (X_1, \dots, X_p)'$ X_i are r.v. defined on (Ω, \mathcal{F}, P) .

Distribution f^{\sim} of \tilde{X}

$$F_{X_1, \dots, X_p}(x_1, \dots, x_p) = P(-\infty < X_1 \leq x_1, \dots, -\infty < X_p \leq x_p)$$

Discrete multivariate distⁿ.

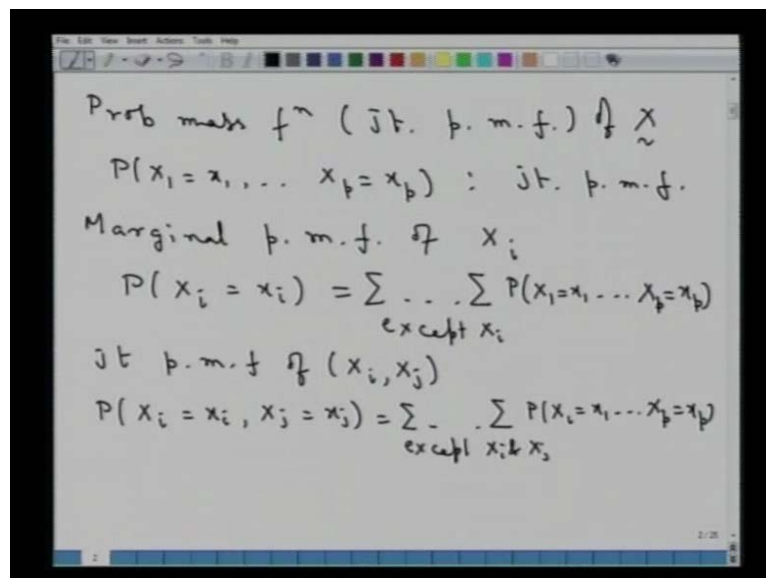
$$F_{X_1, \dots, X_p}(x_1, \dots, x_p) = \sum_{i_1 \leq x_1} \dots \sum_{i_p \leq x_p} P(X_1 = i_1, \dots, X_p = i_p)$$

Let us denote by X ; a set of p random variables x_1, x_2, x_p . So, this is a random vector of random variables, where these X_i 's, actually X_i 's are random variables defined on a probability space ω, \mathcal{F}, P . Now, some basic concepts of multivariate analysis will be actually defining some simple concepts. We will make distinction between discrete multivariate random vector, and continuous multivariate random vectors. So, we define first, the distribution function. Distribution function of this multivariate random vector X is defined to be F_{X_1, X_2, X_p} at the points x_1, x_2, x_p . This is defined as the following that, it is probability that minus infinity less than X_1 less than equal to x_1 and

the last random variable minus infinity less X_p less than equal to small x_p . So, this is how, a distribution function of a multivariate random variable is defined.

Let us now make distinction between discrete multivariate random variable, and continuous multivariate random variable. Suppose we have discrete multivariate distribution; then this distribution function F_{X_1, X_2, \dots, X_p} at small x_1, x_2, \dots, x_p points. This is defined to be summation $\sum_{i_1, i_2, \dots, i_p}$ for rather. So, over all these points, where we have probability that X_1 equal to say i_1 ; X_p equal to i_p ; where these now i_1 is less than equal to small x_1 and i_p is less than equal to this x_p point. So, this particular quantity for a discrete multivariate random variable is defined in this following way. Now, given the joint distribution function of this set of p random variables constituting this multivariate random vector, we can define marginal probability mass functions, joint mass functions.

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So, the probability mass function or rather the joint probability mass function of this multivariate random vector X is defined in the following way: that it is probability that X_1 equal to small x_1 ; X_p equal to small x_p . So, we call this particular quantity, which is the joint probability mass function. Now given this, the joint probability mass function of this random vector X , as we had defined. We can find out, what is called the marginal distributions. The marginal probability mass function of say any random variable in this particular set of p random variables x_1, x_2, \dots, x_p say x_i . This is defined to be probability that, X_i equal to small x_i .

So, this is given by $p - 1$ fold summation, which is except this X_i random variable. So, this is probability over the sum over the joint probability mass function. Like here we have, probability X_1 equal to x_1 extending up to... I am sorry this probability that, this we have the joint probability mass function right. Now, this is the marginal probability mass function of one random variable X_i in the set of the multidimensional random vector, which is comprising of x_1, x_2, x_p .

Now, in a similar way one can actually define the joint probability mass function of any set of variables, taken from this multidimensional random vector x_1, x_2, x_p ; say the joint probability mass function of two random variables X_i and X_j taken from this set of this multidimensional random vector X say X . The joint probability mass function of X_i, X_j is defined to be in a similar way to what we had defined for the marginal probability mass function of one random variable. So, this can be defined as probability X_i equal to x_i, X_j equal to x_j ; this is sum.

Now, the sum is $p - 1$ fold; it is over all X_i 's except X_i and X_j and then we have this joint probability mass function of all the random variables. When we have the joint probability mass function of any set of variables taken from the multidimensional random vector X comprising of those p components; one can define distributions, which are referred to as conditional distributions. say Suppose, we are interested in knowing what is the conditional distribution of any set of variables taken from this p dimensional random vector given another set of random variables.

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Handwritten mathematical derivation on a whiteboard:

$$\begin{aligned} & \text{Conditional dist}^n \text{ of } X_k \text{ given } X_i \text{ \& } X_j \\ & P(X_k = x_k \mid X_i = x_i, X_j = x_j) \\ & = \frac{P(X_k = x_k, X_i = x_i, X_j = x_j)}{P(X_i = x_i, X_j = x_j)} \end{aligned}$$

Say let us be simple and try to look at the following which we say that, say the conditional distribution of X_k given X_i and X_j . So, in order to find what is the conditional distribution of X_k given X_i and X_j ; we look at the following that probability that X_k equal to x_k ; given that X_i equal to small x_i and X_j , the random variable equal to small x_j . So, this would be given by the following that this is the joint probability mass function of these three random variables X_i , X_j and X_k ; this divided by the marginal probability mass function of the two random variables X_i and X_j .

So, here what we have basically, in order to find out the conditional distribution of one random variable given two random variables X_i ; any two random variables X_i and X_j . Of course, X_i and X_j are not included in this X_k here. It can also be. But when we look at this, it is basically **the** this numerator in this conditional distribution of X_k given X_i and X_j is the joint probability mass function of X_i , X_j and X_k and the denominator is X_i and X_j is joint probability mass function. **right**. Let us now move on to the case of conditional rather the marginal distributions, the conditional distributions in case of continuous multivariate random variables.

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Continuous mult. distⁿ

$$F_{X_1, \dots, X_p}(x_1, \dots, x_p) = \int_{-\infty}^{x_1} \dots \int_{-\infty}^{x_p} f_{X_1, \dots, X_p}(x_1, \dots, x_p) dx_1 \dots dx_p$$

$f_{X_1, \dots, X_p}(x_1, \dots, x_p)$: ∂ b. prob density f^n (p.d.f.)

$$= \begin{cases} \frac{\partial^p F_{X_1, \dots, X_p}}{\partial x_1 \dots \partial x_p} & \text{if derivative exists at } (x_1, \dots, x_p) \\ 0 & \text{o/w} \end{cases}$$

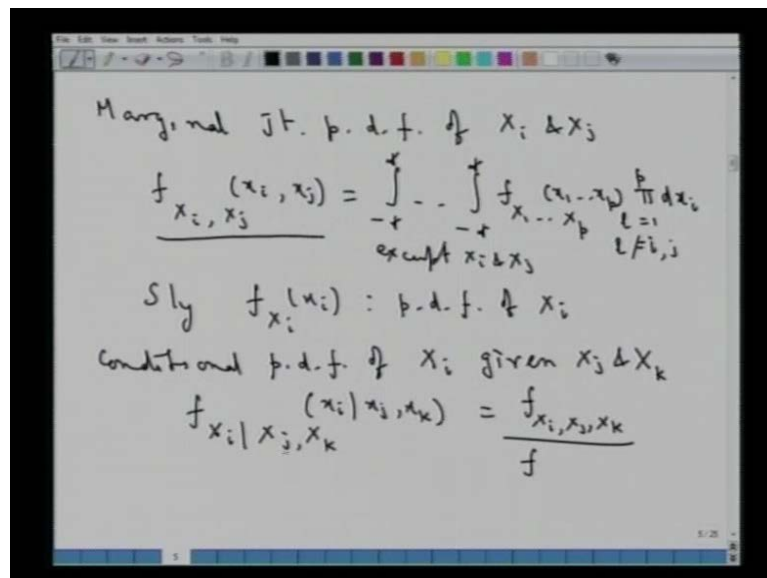
How it looks like? continuous multivariate distribution. So, we had first of all defined, what is the joint distribution function of any set of multivariate random components? So, in case of a continuous multivariate distribution comprising of these p elements, this for a continuous distribution would be defined through the integral. We assume that, the set of random variables form a set of absolutely continuous variables. So, we will have that

been defined through the following functions, small f function which we are going to define shortly.

So, this **this** defined to be the following and then, this is product of x_i 's; i equal to 1 to p . Now, the function that we had introduced here is what is called the joint probability density function of these random variables. So, this quantity is what we referred to as the joint probability density function or in short, pdf of this random vector X . Now, this is given by; so, this quantity is given by the following. So, it is the p th partial differential, this with respect to the variables that we have; if the differentiation exists at the point x_1, x_2, x_p **if the derivative exists at x_1, x_2, x_p** and is equal to zero, if it is otherwise. **Right.**

So, for a continuous random vector X , this quantity is of interest and what we referred to that as, the joint probability density function. Now, given this joint probability density function in a way similar to what we had done for discrete random vector, we can define marginal probability density function of a set of random variables taken from that multivariate random vector. We can define conditional distributions exactly in the same way as, what we had done for a discrete distribution.

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Say for example, we can look at the marginal say, joint pdf of X_i and X_j . So, there are two variables taken from this set of p random variables. So, this would be defined in the way that, this is f_{X_i, X_j} at the point say x_i and x_j and we will have **p minus one** p minus two fold integral here over the entire range of those variables. These integrals are

except the two variables X_i and X_j ; leaving out these two variables, will have p minus two variables integrating over their range and then looking at this as the joint probability density function of x_1, x_2, \dots, x_p . What will be getting here? This coefficient i equal to 1 to p then, i is not equal to j ; it is better to have a different notation here.

Say i equal to 1 to p with i not equal to j ; **this is i equal to one to p then i not equal to j** it is not equal to j , dx_i . So, we will have this particular quantity to give us the joint probability density function of X_i and X_j , two random variables taken from this set. Similarly, what we can look at is **this**. So, this would be a probability density function not the joint **the** probability density function of a particular random variable, X_i taken from this particular set. So, we can also have a conditional distribution, conditional density function of say X_i given X_j and X_k say; X_i and X_k , that would be given by the following that this is X_i given X_j and X_k . So, this would be given by the joint probability density function of X_i, X_j, X_k ; this divided by the joint probability density function of the conditioning variables, that is X_j and X_k . **right**

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$$\underline{X} = (X_1, \dots, X_p)'$$
 is the set of indep. r.v.s

$$P(X_1 = x_1, \dots, X_p = x_p) = \prod_{i=1}^p P(X_i = x_i)$$
 for discrete

$$f_{X_1, \dots, X_p}(x_1, \dots, x_p) = \prod_{i=1}^p f_{X_i}(x_i)$$
 for cont.

This random vector \underline{x} , that we had defined comprising of these p elements X_1, X_2, \dots, X_p is said to be a set of independent random variables; **is the set of independent random variables** independent defined by statistical independent. **is the set of independent random variables**. If we have the joint probability mass function or the joint probability density function given as the product of the respective marginal probability mass function or the **probability** joint probability mass density functions.

Say for example, in the case of discrete distributions, what we will be having for independent random variables is the following that, the joint probability mass function of X_1, X_2, X_p would be given by the product of the respective probability mass functions given as the following. So, this is for the discrete distribution. **for discrete distribution** and If we have continuous distribution, then the joint probability density function of X_1, X_2, X_p would be given by the product of the marginal probability density functions. This would be for the continuous random variable. Now, let us now move on to some concepts, which are going to be important for this particular course.

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Expectation vector of X
 $\tilde{X} \sim p_{X1}$
 $\mu = E(\tilde{X}) = \begin{pmatrix} E(X_1) \\ \vdots \\ E(X_p) \end{pmatrix}$
 $\tilde{X} \xrightarrow{\sim p_{X1}} \tilde{Y} = \underline{A} \tilde{X} + \underline{b}$
 $\underline{A}_{q \times p}$: matrix of constants
 $\underline{b}_{q \times 1}$: vector of constants
 $\mu_Y = E(\tilde{Y}) = \underline{A} E(\tilde{X}) + \underline{b} = \underline{A} \mu + \underline{b}$

So, we define the expectation vector of this random vector X to be the following that let us denote that by the vector μ , which is also a p dimensional vector, which is expectation of this random vector. So, that is defined as the following. So, X component Y is expectation of the corresponding random variables X_1, X_2 and X_p . Now, these expectations **expectation** x_i in appearing in this particular vector; they basically are computed from the probability density function of the respective random variable X_i , that is coming in this particular direction here. Now, given the information that X , the random vector which is a p dimensional random vector with an expectation vector as μ ; if you make a transformation **say suppose we make a transformation** which is say p dimensional random vector to a random vector Y , which is given by say $\underline{A} X$ plus \underline{b} .

Now, here this \underline{A} matrix which may be is q by p which is matrix of constant. So, this is matrix of constants and \underline{b} ; so, this now becomes q by p and this is p by 1 . So, this

particular component out here is q by 1. So, b is say q by 1 vector. So, this is a vector of constants. So, we take here this A matrix and this b vector out here to be non-stochastic elements actually. And then, if we are now interested to find out, what is the expectation vector of this newly defined random vector, which is Y; say let us denote that by mu Y. So, this would be given by A matrix, which is a matrix of constants and that would then be given by expectation of X plus this b vector **right**. So, this would be given by A mu **a mu** plus this b. **right** Having defined this expectation vector, we move on to defining what we mean by covariance matrix of this random vector X?

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Covariance matrix of \underline{X} .

$$\Sigma = \text{Cov}(\underline{X}) = E(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})'$$

$\underline{\mu} = E(\underline{X})$

$(i, j)^{\text{th}}$ element of \underline{X}

$$\sigma_{ij} = E(X_i - \mu_i)(X_j - \mu_j)$$

$\mu_i = E(X_i)$
 $\mu_j = E(X_j)$

$$\sigma_{ii} = E(X_i - \mu_i)^2$$

Correlation matrix of \underline{X}

$$\rho = \text{Corr}(\underline{X}) = \Sigma^{-1/2}$$

Let us define the covariance matrix of this random vector X. Now, let us denote that by sigma, which is covariance matrix of X; which is given by expectation of X minus expectation of X, which **which** we have already denoted by this mu vector that into X minus mu transpose. So, this mu is basically expectation vector of the random vector, which is X **right**. Now, the i j th element in this covariance matrix of X i j th element of this X is given by say sigma i j, which is simply the covariance between these two random variables X i and X j. So, let us denote by mu i to be expectation of the i th component of the mu vector. Then, this is X j minus mu j, where we have mu i is expectation of X i and similarly, mu j is expectation of the X j component **right**.

Now, this is the i j th element of this sigma matrix. So, the diagonal elements basically are giving us the variances of the respective components, which are p in number. So, this is just the variance of the i th component, which is expectation of X i minus mu i whole

square **right**. Now, having defined this covariance matrix of this random vector X, we can define what is referred as the correlation matrix? Correlation matrix of X similarly can be defined. Say, this is a row matrix, which is holding the correlation components between the elements of this random vector. So, it is a correlation matrix of this.

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The image shows a whiteboard with handwritten mathematical derivations. The first line is $V = \text{diag}(\sigma_{11}, \dots, \sigma_{pp})$. The second line is $\underline{X}_{p \times 1} \rightarrow E(\underline{X}) = \underline{\mu}, \text{Cov}(\underline{X}) = \Sigma$. The third line is $\underline{X} \rightarrow \underline{Y} = A \underline{X} + \underline{b}; \text{Cov}(\underline{Y}) = ?$. The fourth line is $\Sigma_Y = \text{Cov}(\underline{Y}) = E(\underline{Y} - E(\underline{Y}))(\underline{Y} - E(\underline{Y}))'$. The fifth line is $= E(A \underline{X} + \underline{b} - (A \underline{\mu} + \underline{b}))(\dots)'$. The sixth line is $= E(A(\underline{X} - \underline{\mu}))(\dots)'$. The seventh line is $= E(A(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})' A')$. The eighth line is $\text{Cov}(\underline{Y}) = A E(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})' A'$. The ninth line is $\text{Cov}(\underline{Y}) = A \Sigma A'$.

Let us define that to be a matrix V half inverse times the sigma matrix into V half minus half. Let me also define, what we mean by this V matrix? So, this V matrix is a diagonal matrix holding the elements sigma 1 1, sigma 2 2 and the sigma p p. So, the V matrix is basically the matrix, which is having the diagonal entries only and comprising of the variances of the respective components. Now, we had seen what happens to when we have a random vector X? And then, we make a transformation from the random vector X to the random vector Y; which was from p dimension to a q dimension lower or higher dimension.

What we are, now what we can also look at is given the information that, X is a p dimensional random vector with expectation of X as a mu vector and the covariance matrix of X to be given by the sigma matrix. If we now make a similar transformation as to what we had done earlier; to a random vector Y, which is A X plus b. A and b are defined similarly as to what we had defined previously in this slide. So, this basically is defined exactly in the same way that, **I am sorry**. So, this is that A matrix. This A matrix is a matrix of constants; b is vector of constants.

So, given that information, we are now trying to look at, what is the covariance matrix of the new random vector that we have introduced, which is Y . Now, by definition that, the covariance matrix of Y would be given by Σ_Y . Let us denote that by Σ_Y , for example. So, this Σ_Y is expectation of Y minus expectation of Y into the transpose of this particular quantity. So, we have already computed, what is expectation of Y from this random vector A . So, Σ_Y we just replace this by the component in terms of X . So, this would be expectation of $A X$ plus b ; this minus, now expectation Y as we had seen earlier is A times the μ vector; this plus this b vector.

So, we have this quantity here given by this b and then, the transpose of this particular quantity comes out here. So, what we will be having is, this b component cancelling out. So, we will have expectation of $A X$ minus μ that multiplied by the transpose of this quantity. So, it is $A X$ minus b . So, this would now be given by expectation of A into X minus μ X minus μ transpose and then we will have this A transpose. So, what we can see here is that this basically is the stochastic component here. This A matrix and the A transpose matrix basically are the two matrixes of constants.

And hence, we can take the expectation operator inside and what we will be having here is, $A X$ minus μ to X minus μ transpose A transpose. And then this expectation of X minus μ into X minus μ transpose is nothing but the covariance matrix of X which is given earlier by Σ_X matrix. So, this is what we will be having as the covariance matrix of the newly defined random vector, which is Y . Now, suppose we have two random vectors X and Y ; two different dimensions. We can also define; what is the covariance matrix between the two components here.

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$\underline{X} \sim p \times 1$: p -dim random vector
 $\underline{Y} \sim q \times 1$: q -dim random vector
 $\text{Cov}(\underline{X}, \underline{Y}) = E(\underline{X} - E(\underline{X}))(\underline{Y} - E(\underline{Y}))'$
 Partition of Cov matrix \rightarrow
 $\underline{X} \rightarrow \text{Cov}(\underline{X}) = \Sigma$
 $\underline{X} = \begin{pmatrix} \underline{X}^{(1)} & r \times 1 \\ \vdots & \vdots \\ \underline{X}^{(2)} & p-r \times 1 \end{pmatrix}$ $\underline{\mu} = \begin{pmatrix} \underline{\mu}^{(1)} & r \times 1 \\ \vdots & \vdots \\ \underline{\mu}^{(2)} & p-r \times 1 \end{pmatrix}$ $E(\underline{X}^{(1)})$

So, suppose we have this p dimensional random vector. So, suppose we have this as p dimensional random vector and we have Y , another q dimensional random vector. We can define, what is the covariance matrix between the elements of X and the elements of Y ; similar to what we had defined by covariance of the components in that particular vector. This would be given now by expectation of X minus expectation of X vector; this multiplied by Y into expectation of this Y vector whole transpose. So, this is how, one defines the covariance matrix between two different sets of random vector.

It is important actually to look at the following, which we look at the partitions of the covariance matrix. What the partitions and what the elements actually can; what sort of interpretation actually can; the elements of those partitioned elements can be having actually? **partition of covariance matrix**. So, suppose we have this random vector X with the covariance matrix of X as sigma matrix. Let us make the following partition. Suppose, this X vector which was p dimensional vector; now is partitioned into two following sub vectors say X_1 and X_2 .

Say suppose this is r by 1 and this is p minus r by 1 ; similar to this X vectors partition, we can look at the corresponding partition in the μ vector, which is the expectation vector. So, we can write the similar partition as μ_1 vector, which is r by 1 vector out here and then, we will have the second sub vector as p minus r dimensional. So, this quantity here is expectation of the corresponding sub vector, what we had as X_1 written

out there. So, this would be expectation of X_1 and similarly, this would just be expectation of this particular component X_2 from here.

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The image shows a handwritten derivation of the covariance matrix of a random vector \underline{X} . The derivation starts with the definition of covariance:
$$\text{Cov}(\underline{X}) = E[(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})']$$

$$= E\left(\begin{pmatrix} X^{(1)} - \mu^{(1)} \\ X^{(2)} - \mu^{(2)} \end{pmatrix} \begin{pmatrix} X^{(1)} - \mu^{(1)} & X^{(2)} - \mu^{(2)} \end{pmatrix}'\right)$$

$$= E\left(\begin{pmatrix} X^{(1)} - \mu^{(1)} & X^{(2)} - \mu^{(2)} \end{pmatrix} \begin{pmatrix} X^{(1)} - \mu^{(1)} \\ X^{(2)} - \mu^{(2)} \end{pmatrix}'\right)$$

$$= \begin{pmatrix} \sum_{11} & \sum_{12} \\ \sum_{21} & \sum_{22} \end{pmatrix}$$

Below the matrix, the elements are defined:
$$\sum_{11} : \text{Cov}(X^{(1)})$$

$$\sum_{22} : \text{Cov}(X^{(2)})$$

$$\sum_{12} : \text{Cov}(X^{(1)}, X^{(2)})$$

Now, if we look at the covariance, the partition that we have corresponding to the partitioning of the X vector that is what we had written. So, this by definition of the covariance matrix is given by the following that it is this particular quantity. Then we look at what happens to this particular element, when we look at this as the sub vectors. So, this would be X_1 minus μ_1 . So, this vector would be the second component X_2 minus μ_2 and then, the transpose of this comes out here. Now, if one multiply the two and then take the expectation inside or $E(\cdot)$ to that one can just look at, what would be this particular matrix?

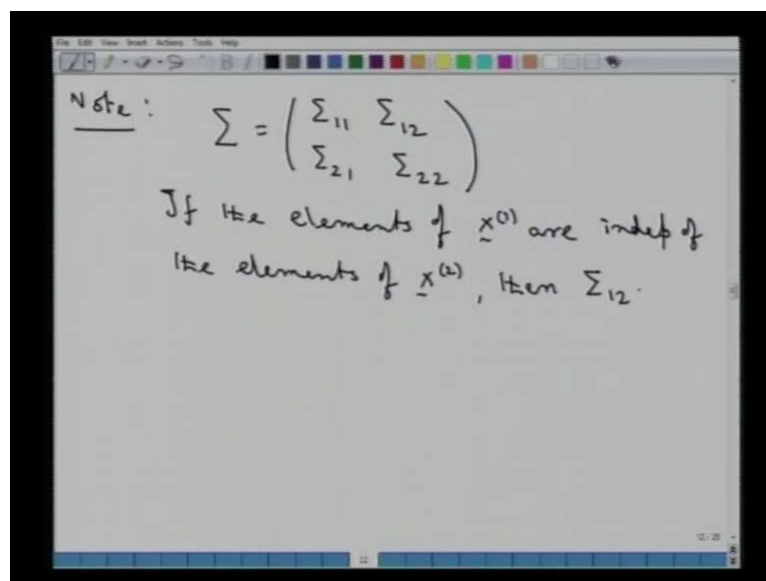
This would be given by this X_1 minus μ_1 X_1 minus μ_1 transpose and then this block here. So, this would comprise of four blocks. This would be the second component X_2 minus μ_2 ; then this is X_2 minus μ_2 transpose. And then this element would be X_1 minus μ_1 into X_2 minus μ_2 transpose and this would just be the transpose of this particular element there. So, this element once we take the expectation operator inside would be **this would be** a p by p matrix, which we denote by σ_{11} say; this would be a p by q matrix, which we denote by σ_{12} . And then the expectation of the quantity which is here would be a q by p , which is denoted by say σ_{21} .

This element, the last block here will have this partitioning actually carried forward. So, when we take the expectation of this block here, that is denoted by σ_{11} . When we

take expectation of this particular block here, that is going to be denoted by Σ_{11} block, which is a q by q dimensional matrix there. When we take the expectation of this block here, it is basically denoted by Σ_{11} . Now, this Σ_{12} element; if you look at carefully, this Σ_{12} block here is expectation of this particular matrix here and what is that? **That** that is, basically is the covariance matrix of the sub vector, what we had denoted by X_1 .

Similarly, if we are looking at Σ_{22} ; Σ_{22} is nothing but the covariance matrix of the second block actually that, this particular X_2 block of elements and this Σ_{12} matrix, which is the off diagonal blocks in this sigma matrix. So, this is the covariance matrix of the two sub vectors X_1 and X_2 that is what we had defined. So, this is the covariance matrix of the two blocks X_1 and X_2 **right**. This is just the transpose of this particular matrix and that one can say that, it is the covariance matrix of X_2 and X_1 . So, it just differing by that particular transpose.

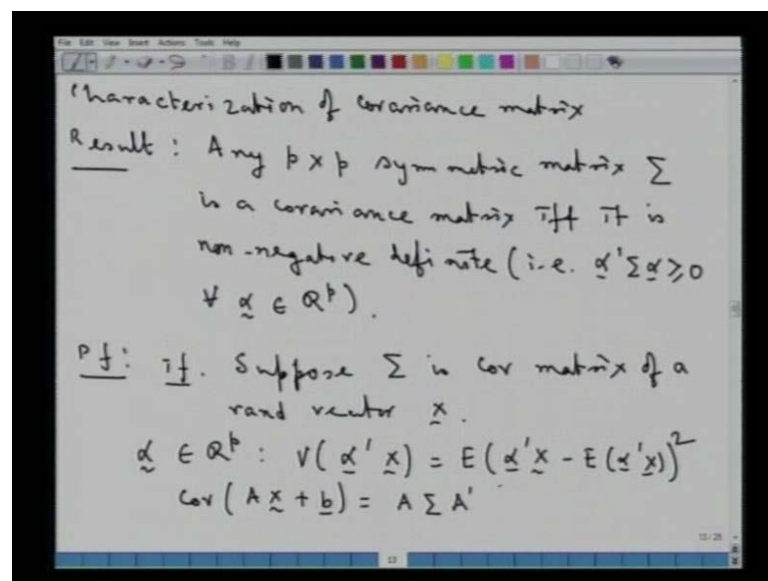
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Now, we make a small note here; that, we had this sigma matrix in the partition form written as **sigma 1 1** sigma 1 2 sigma 2 1 sigma 2 2. Now, if we have X_1 and X_2 , independent set of random variables then, what will be having is this sigma 1 2 will be given by a null matrix. Now, if the elements of X_1 are **independent** independently distributed of the elements of the X_2 sub vector; then we will be having this off diagonal block here sigma 1 2 to be given by a null matrix. However, the converse of this particular result is not true.

What I mean by saying that is that, if we have Σ to be given by null matrix then, it does not imply that, the elements of X_1 and X_2 are independent. It does not imply. It would only be implied, if we have some other condition like, what we had for bivariate distribution that; if we have the joint distribution of X_1 and X_2 to be given by a multivariate normal. Then only, we will have Σ to be given by a null matrix; that would imply in such a situation that, the element of X_1 and X_2 are independent **right**. Now, let us move on to one characterization of covariance matrix, what we are defining. So, let us now look at one characterization.

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How we can characterize a covariance matrix? So, characterization of a covariance matrix, we have the following result. Let me first state the result that if we have any matrix which is positive semi definite, it can be associated with covariance matrix of some random variable. So, if we have a p by p symmetric matrix Σ , will be a covariance matrix if and only if, it is positive semi definite or non-negative definite. So, let me write the result first that, any p by p symmetric matrix Σ is a covariance matrix if and only if, it is non-negative definite. That is, $\alpha' \Sigma \alpha \geq 0$ for every α belonging to \mathbb{R}^p **right**.

So, that gives us a characterization of a covariance matrix. So, let us look at the proof of this particular simple fundamental result. So, what it says is that, if we have Σ to be the covariance matrix of **a of** some random vector X , then it is going to be positive semi definite, non-negative definite. Other way round, if we have any matrix which is

symmetric and it is non-negative definite, then we can associate some random vector to that particular matrix. For which, the random vector is going to have that matrix as its covariance matrix. So, let us look at the proof of this result. Let us first **proof a** look at the proof of the if part.

Now, suppose sigma is **is covariance matrix covariance matrix of some random variable** **x** covariance matrix of a random vector X; then we take a alpha belonging to R to the power p. Then, what we can say is that the following that variance of alpha prime X. Now, note that this alpha prime X is a scalar random variable. So, it is a linear combination of the elements of that p dimensional random vector. By definition, this actually would be given by expectation of this random variable alpha prime X minus expectation of this alpha prime X whole square **right**.

Now, we actually had proved a result that variance of or rather the covariance matrix we had proved the result; so that, we can even use that particular result. We had proved the result that, covariance matrix of A X plus b; this matrix is given by A covariance matrix of X, which is denoted by sigma; A sigma A prime. So, we can use this particular result, in order to find what is the variance of X? Otherwise from definition, one can also find out **what is the** what is the expectation of this particular quantity? Let us look at that.

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$$\begin{aligned}
 V(\alpha' \underline{x}) &= E(\alpha' \underline{x} - \alpha' \underline{\mu})^2 \\
 &= E(\alpha' (\underline{x} - \underline{\mu}))^2 \\
 &= E(\alpha' (\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})' \alpha) \\
 &= \alpha' \Sigma \alpha \geq 0 \quad \forall \alpha \in \mathbb{R}^p. \\
 &\Rightarrow \Sigma \text{ is n.n.d.}
 \end{aligned}$$

only if: Suppose Σ is n.n.d. of rank $r \leq p$
 $\Sigma = C C'$; C is $p \times r$ of $r(C) = r$
 Let \underline{y} be a random vector of in

So, this is variance of alpha prime X; this is given by expectation of this alpha prime X minus alpha prime mu. mu is the expectation vector of that covariance that particular element, which is alpha prime X this whole square. This would be given by expectation;

we take alpha prime out. So, this is X minus μ whole square. So, this would be given by expectation of alpha prime X minus μ . Then this is the transpose of that; that multiplied this alpha vector. So, this would just be given by alpha prime; then the covariance matrix of X , which is the σ alpha. This is the variance of X .

Now, variance of alpha prime X is going to be given by this alpha is belonging to \mathbb{R} to the power p . So, since this is the variance of this random variable alpha prime X , which is going to be greater than or equal to 0. So, we have proved that, if σ is actually the covariance matrix of any random vector X , then for every alpha belonging to \mathbb{R} to the power p alpha prime σ alpha is going to be greater than or equal to 0. So, that it is non-negative. It is going to be non-negative definite for every alpha belonging to \mathbb{R} to the power p . So, this basically implies that, σ matrix is non-negative definite. So, let me just write that as, $n \times n$ d. So, that is non-negative definite.

So, the other way round, the only if part we look at, suppose σ is non-negative definite we will prove that, σ is a covariance matrix associated with some random vector x . Now, since we have assumed that σ is non-negative definite, say suppose it is a non-negative definite of rank r which is less than or equal to the dimension of this **random of this** matrix σ . Since σ is non-negative definite of rank r , we can write this σ matrix in the following form that, it is $c c'$; where c is p by r of full column rank **right**. Now, since we have factorized σ into this $c c'$ form, let us now define a random vector Y .

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Variables with $E(\underline{y}) = \underline{0}$, $\text{Cov}(\underline{y}) = I_r$

$$\underline{y} \rightarrow \underline{x} = c \underline{y}$$

$$\Rightarrow E(\underline{x}) = \underline{0} (= c E(\underline{y}))$$

$$\text{Cov}(\underline{x}) = \text{Cov}(c \underline{y}) = c \text{Cov}(\underline{y}) c'$$

$$= c I_r c'$$

$$= c c' = \Sigma$$

$\Rightarrow \Sigma$ is a cov matrix.

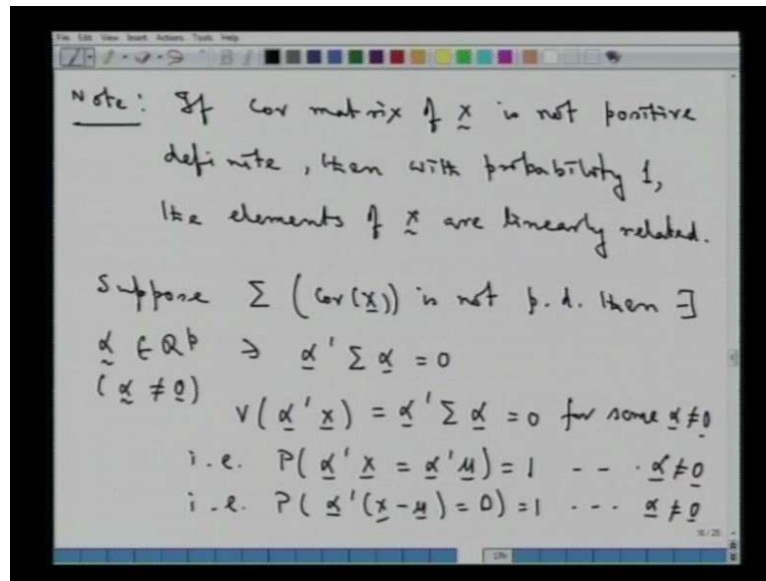
Let Y be a random vector of independent random variables of independent random variables with say expectation of Y given by a null vector. Now, since we assume that, the components Y_1, Y_2, \dots, Y_r ; what we have for the random vector Y ? They are independent. Then, the covariance matrix actually would be given by a diagonal matrix. Using the result that we had stated earlier that, if we have the components of X_1 , that sub vector to be independent of the components of the X_2 sub vector. Then, the block which is holding the covariance's of the components of X_1 and X_2 are going to be 0. So, this would imply that, the covariance matrixes diagonal without loss of generality, we take that to be an identity matrix.

So, the covariance matrix of Y is given by say this identity matrix of dimension r right. Now, Y is that set of independent random variables with the expectation of Y vector to be a null vector and the covariance matrix of this Y vector to be an identity matrix of dimension r . Now, let us make a transformation from Y to X , where X is given by c times this Y . Now, the matrix c what we had defined earlier, which is a p by r matrix; which is a matrix of constants of full column rank. So, that is of rank r . So, with that c matrix, if we now look at what is this X matrix characteristic?

As far as its expectation vector and its covariance matrix are concerned, we will have; this implies that, expectation of this X vector would be given by a null vector; because this is equal to C times expectation of this y vector and expectation of a Y vector being a null vector. This would imply that, it is just a null vector. And then, the covariance matrix of this X vector is covariance matrix of this C times Y vector. That would be given by C times covariance matrix of Y into this C transpose matrix. Covariance matrix of Y , what we have assumed is an identity matrix. It is a the covariance matrix associated with iid random variables actually with variance is equal to 1.

So, this would be given by C times this I_r matrix C transpose. So, this is equal to $C C'$ prime and which is nothing but that sigma matrix. This would imply that, this sigma matrix what we had just taken as a non-negative definite matrix. It is now the covariance matrix of this random vector X right. So, this would actually prove that, sigma is a covariance matrix. So, we have proved the result that any p by p symmetric matrix which is non-negative definite is if a p by p symmetric matrix is non-negative definite is a covariance matrix if and only if, it is non-negative definite.

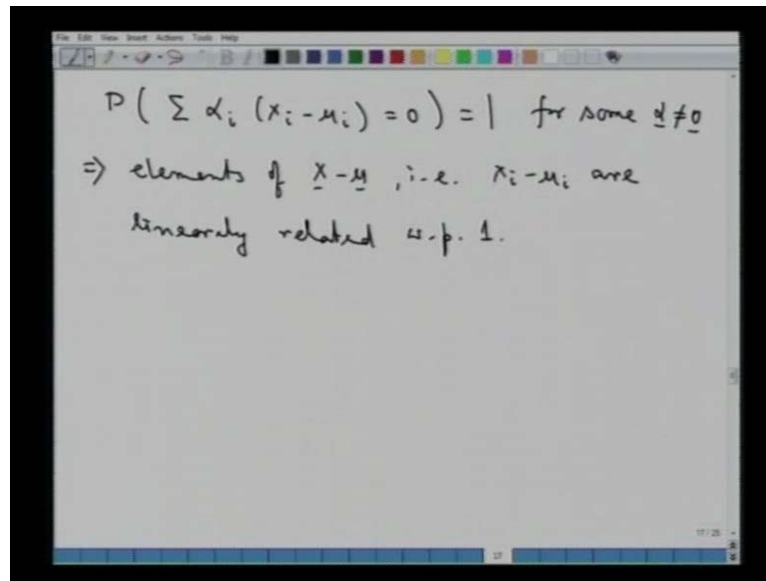
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We just put it as a note the following result that, if covariance matrix of X is not positive definite, then with probability 1, the elements of X are linearly related with probability 1 the elements of x are linearly related. What do we mean by this is the following that, suppose σ , the covariance matrix of this X random vector is not positive definite, then there exists an α belonging to \mathbb{R} to the power p such that, we will have α prime. Now, these α vectors are belonging to \mathbb{R} to the power p of course, α is not equal to 0. So, this goes without saying that, this α is not equal to 0. Then, this α prime $\sigma \alpha$ that is equal to 0. So, there exists some α such that, this α prime $\sigma \alpha$ is equal to 0. Now, what does this mean?

This means basically that, what we had seen earlier is that, this α prime $\sigma \alpha$ is nothing but variance of this α prime X . So, we have this variance of α prime X given by α prime $\sigma \alpha$; this is equal to 0 for some α which is not equal to a null vector. What it means is the following that, probability that α prime X is equal to its expectation α prime μ ; this is equal to 1 for some α which is not equal to 0 right. That is what we have is α prime X minus μ ; this is equal to 0; this probability is 1 for some α , which is not equal to 0 right. So, this basically would imply from the statement that we have made here that, probability that this particular linear combination is equal to 0; this is equal to 1 for α not equal to 0.

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So, this would imply that, probability that summation α_i ; the i th component X_i minus μ_i ; this is equal to 0 is equal to 1 for some α , which is not equal to a null vector. So, this would imply that, the elements of $X - \mu$; that is, these $X_i - \mu_i$'s are linearly related with probability 1 **right**. So, we will have the elements of the X_i components; because these μ_i components, they are basically constant components. So, we will have the elements X_i is actually being linearly related, if we have the matrix σ , the covariance matrix to be a matrix such that, it is not positive definite **right**. Before we move on to random sampling **of** from a multivariate distribution, we look at to concept, which is moment generating function. From where, we can get actually joint moments of the components, which are there in that multidimensional random vector.

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Moment Generating f^n

$$\underline{X} \sim \text{bxi} : E(\underline{X}) = \underline{\mu} , \text{Cov}(\underline{X}) = \Sigma$$

m.g.f. of \underline{X}

$$M_{\underline{X}}(\underline{t}) = E(\exp(\underline{t}'\underline{X}))$$

Marginal m.g.f. \wedge

$$M_{X_1, \dots, X_n}(t_1, \dots, t_n) = M_{\underline{X}}(\underline{t})$$

Marginal m.g.f. of X_i

$$M_{X_i}(t_i) = M_{X_1, \dots, X_n}(0, \dots, t_i, \dots, 0)$$

So, let us look at that moment generating function. So, the concept of this moment generating function of the random vector X is, basically trying to generalize what we actually have for univariate random variable. So, we have **say** this X , a p dimensional random vector with expectation of X given by this μ vector. Covariance matrix of X is being given by the sigma matrix. Then the moment generating function mgf of this random vector X is defined to be the following that, it is expectation of E to the power t prime X **right**.

So, this is how, one defines the moment generating function of the random vector, the p dimensional random vector X . Now, given the information; now, this provided the expectation exists, now the moment generating function of this random vector X actually would also lead us to the marginal moment generating function of the components of X in the following way. Now, if you look at, how to get the marginal moment generating functions **marginal moment generating functions**. Say, we are looking at this **the** moment generating function X_1, X_2, X_n at these points **x_1, x_2, x_n** ; that is what is our $M_X(t)$, the moment generating function.

So from here, if we want to find the marginal mgf of X_i , that would be derived from this; which is the joint moment generating function of the components of X in the following way. That, the marginal moment generating function of X_i at the point t_i ; that would be given from the joint moment generating function of X_1, X_2, X_n with 0 at all other points, except the i th position which is say t_i . So, this is the i th point. So, this is

what is going to give us. The marginal moment generating function of the component, which is X_i .

To see how that is true, it is trivial to look at the expression of this particular joint moment generating function of the random vector X . So, here if we take all the t_i 's except one t_i component here as 0, then this particular term here; which is this term is the summation $t_i X_i$. So, here if all the t_i 's except one particular t_i is 0. Then this sum is nothing but just $t_i X_i$. So, expectation of e to the power $t_i X_i$ is just the moment generating function of this particular component which is X_i . Now, given **given** this particular moment generating function, one can actually look at the joints moments.

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The image shows a whiteboard with the following handwritten text and formula:

joint moments

$$E(x_1^{k_1} x_2^{k_2} \dots x_p^{k_p}) = \left. \frac{\partial^{k_1 + \dots + k_p} M_x(t)}{\partial t_1^{k_1} \dots \partial t_p^{k_p}} \right|_{t_1 = \dots = t_p = 0}$$

Joint moments, actually of any order provided the **moment the** corresponding moments exists can be obtained from the following. Say, suppose we are looking at this particular joint moment, we can obtain **the** this joint moment for this set of random variables X_1, X_2, \dots, X_p or any set of random variables derived from it. Using the moment generating function in the following way that, it is the partial derivative of the joint moment generating function; this, the derivative with respect to all these t_i 's. This evaluated at t_1 equal to t_2 equal to t_p ; this equal to 0. So, this is how, from the joint moment generating function $M_x(t)$, one can get to the joint moments of any order provided that order moment exists; one can get to the joint moments derived from that **right**.

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$$\underline{X} \sim N(\underline{\mu}, \Sigma) \rightarrow E(\underline{X}) = \underline{\mu}, \text{Cov}(\underline{X}) = \Sigma$$

Q. f. $\underline{X}' A \underline{X}$: A $p \times p$ matrix of constants

$$\begin{aligned}
 E(\underline{X}' A \underline{X}) &= E(\text{tr } \underline{X}' A \underline{X}) \\
 &= E(\text{tr } A \underline{X} \underline{X}') \\
 &= \text{tr } A E(\underline{X} \underline{X}') \quad \text{--- (1)}
 \end{aligned}$$

$$\begin{aligned}
 \Sigma &= E(\underline{X} - \underline{\mu})(\underline{X} - \underline{\mu})' = E(\underline{X} \underline{X}') - \underline{\mu} \underline{\mu}' \\
 \Rightarrow E(\underline{X} \underline{X}') &= \Sigma + \underline{\mu} \underline{\mu}' \quad \text{--- (2)}
 \end{aligned}$$

using (2) in (1)

Now, a result which is particularly useful in some applications; as we move, when this particular multivariate, the normal course; we will see that, we are very frequently concerned with the a quadratic form of the following nature that, suppose we have this \underline{X} ; this multivariate vector with expectation of \underline{X} to be this $\underline{\mu}$ vector and the covariance matrix of \underline{X} to be this Σ matrix. Then, if we define the quadratic form as, $\underline{X}' A \underline{X}$; this is a quadratic form in \underline{X} . So, where this A has ofcourse, the interpretation; that it is a matrix of constants. So, it is a non-stochastic matrix.

So, this is say A p by p matrix of constants, then expectation of $\underline{X}' A \underline{X}$ is the expectation of this particular quadratic form. So, this actually can be obtained in the following way that, now note that this is scalar quantity. This, we can write as expectation of the trace of this $\underline{X}' A \underline{X}$. We can take, we can use the result that trace of $A B$ equal to trace of $B A$. So, what we can write is the following that, this is trace of $A \underline{X} \underline{X}'$. We can now take the expectation operator inside the trace. So, this would be the trace of A expectation of $\underline{X} \underline{X}'$. Now, let me give this equation number 1.

Now, in order to find out what is the expectation of $\underline{X} \underline{X}'$? We see that, this Σ is expectation of $\underline{X} - \underline{\mu}$ into $\underline{X} - \underline{\mu}$ transpose. If you look at, what the expectation of this quantity is. So, it would turn out that, this is expectation of $\underline{X} \underline{X}'$ minus $\underline{\mu} \underline{\mu}'$; just open it up, then take expectation inside. So, this would tell us that, expectation of this $\underline{X} \underline{X}'$ is nothing but Σ plus $\underline{\mu} \underline{\mu}'$

transpose. So, if you use this equation number 2 in equation 1. So, using this equation number 2 in 1, what we get is the following that what we had there was trace of this particular quantity.

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The image shows a whiteboard with the following handwritten equations:

$$\begin{aligned}
 E(\underline{x}' A \underline{x}) &= \text{tr} A E(\underline{x} \underline{x}') \\
 &= \text{tr} A (\Sigma + \underline{\mu} \underline{\mu}') \\
 &= \text{tr} A \Sigma + \text{tr} A \underline{\mu} \underline{\mu}' \\
 E(\underline{x}' A \underline{x}) &= \text{tr} A \Sigma + \underline{\mu}' A \underline{\mu}
 \end{aligned}$$

So, we have expectation of $\underline{x}' A \underline{x}$, that was given by trace of A expectation of $\underline{x} \underline{x}'$. Now, we have obtained that in the form of this two here; that expectation of $\underline{x} \underline{x}'$ is $\Sigma + \underline{\mu} \underline{\mu}'$. So, this would be now be given by $\text{tr} A \Sigma + \text{tr} A \underline{\mu} \underline{\mu}'$. So, this expression now would just be given by trace of $A \Sigma$; this plus trace of $A \underline{\mu} \underline{\mu}'$. And then, that can be written actually as the following which is $\underline{\mu}' A \underline{\mu}$; using once again the result that trace of $A b$ equal to trace of $b A$. So, what we have is the following result that if we have that multi-dimensional random vector \underline{x} with a mean vector equal to $\underline{\mu}$ and a covariance matrix equal to Σ , then expectation of this quadratic form is just given by this trace of $A \Sigma$ plus $\underline{\mu}' A \underline{\mu}$; similar result for the variance of this quadratic form can also be derived.