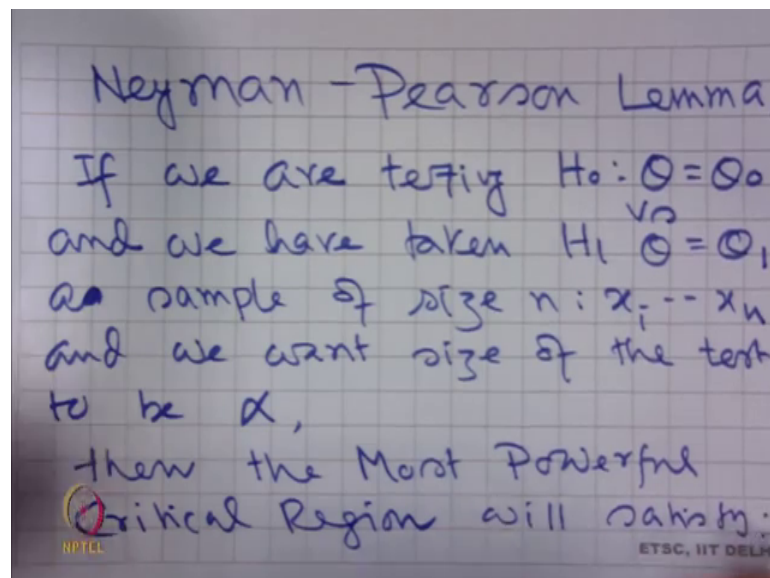


Statistical Inference
Prof. Niladri Chatterjee
Department of Mathematics
Indian Institute of Technology, Delhi

Lecture – 21
Statistical Inference

Welcome students to the MOOCs series of lectures on Statistical Inference. This is lecture number 21 and you remember that we have been discussing testing of hypothesis. In particular our focus has been on Neyman-Pearson Lemma.

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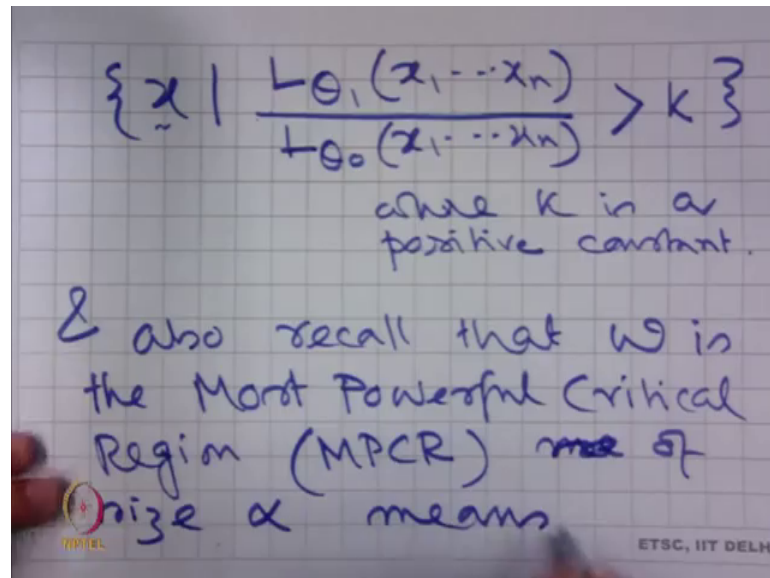
So, what it says? That if we are testing $H_0: \theta = \theta_0$ versus $H_1: \theta = \theta_1$, and we have taken a sample of size n : namely x_1, x_2, \dots, x_n and we want size of the test to be α then the most powerful critical region will satisfy the following.

(Refer Slide Time: 02:23)

$$\left\{ \underline{x} \mid \frac{L_{\theta_1}(x_1, \dots, x_n)}{L_{\theta_0}(x_1, \dots, x_n)} > k \right\}$$

where k is a positive constant.

& also recall that W is the Most Powerful Critical Region (MPCR) of size α means



It is the set of observations \underline{x} , such that $L_{\theta_1}(x_1, x_2, \dots, x_n)$ where this is the likelihood function and under the alternative hypothesis is H_1 upon $L_{\theta_0}(x_1, x_2, \dots, x_n)$ which is the likelihood function under H_0 has to be greater than K , where K is a positive constant. And also recall that W is the most powerful critical region that is MPCR means of size α , means:

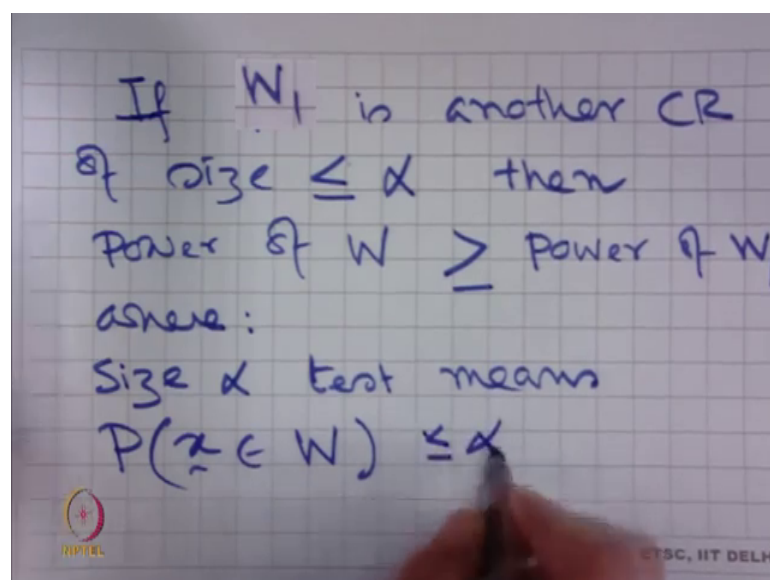
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If W_1 is another CR of size $\leq \alpha$ then

$$\text{Power of } W \geq \text{Power of } W_1$$

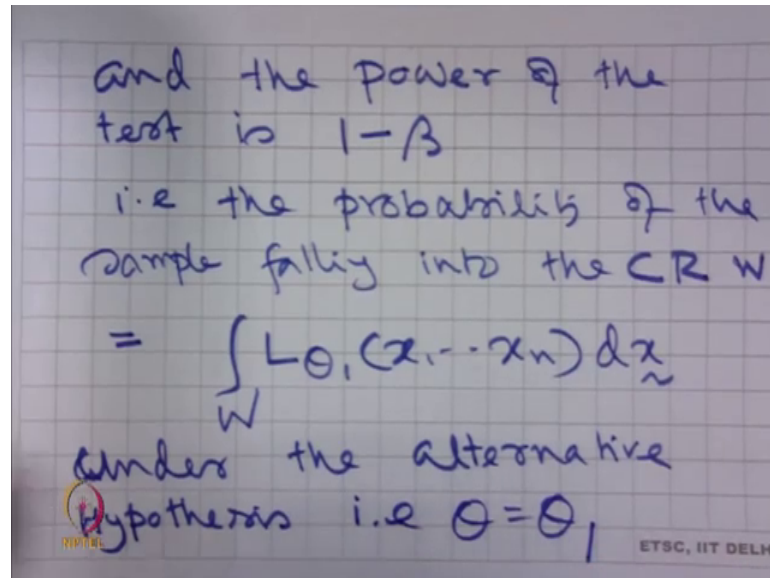
where:

Size α test means

$$P(\underline{x} \in W) \leq \alpha$$


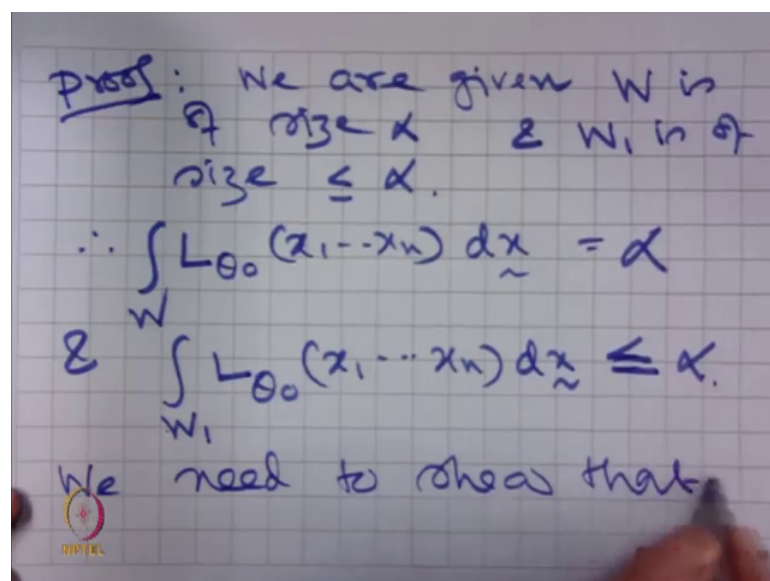
If W_1 is another critical region of size less than equal to α , then power of W is greater than equal to the power of W_1 , where size α test means probability x belonging to the W is less than equal to α and the power of the test is $1 - \beta$.

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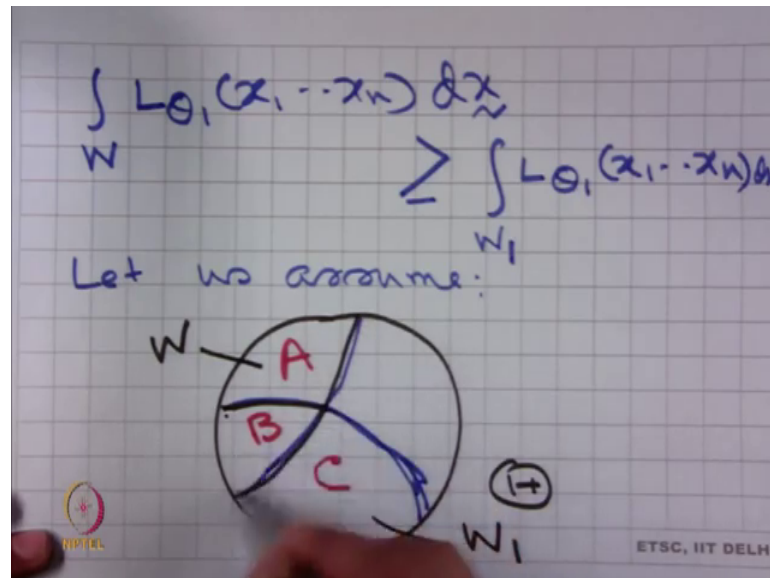
That is the probability of the samples falling into the critical region W and that is equal to integration over W $L_{\theta_1}(x_1, x_2, \dots, x_n) dx$ under the alternative hypothesis that is θ is equal to θ_1 .

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So, we want to prove this we are given W is of size α and W_1 is of size less than equal to α therefore, integration over W of $L_{\theta_0}(x_1, x_2, \dots, x_n) dx$ is equal to α and integration over W_1 $L_{\theta_0}(x_1, x_2, \dots, x_n) dx$ is equal to or less than equal to α . We need to show that integration over W $L_{\theta_1}(x_1, x_2, \dots, x_n) dx$ is greater than equal to integration over W_1 of $L_{\theta_1}(x_1, x_2, \dots, x_n) dx$.

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So, let us assume, this is the sample space θ . Let W be this is W which is A union B and suppose this is W_1 which is B union C . So, that is going to be that general diagram W is a subset of the parameter space θ and W_1 is another subset of the parameter space θ . In general they may have an intersection and according to the diagram the intersection is B .

(Refer Slide Time: 09:00)

$$\frac{L_{\theta_1}(x_1, \dots, x_n)}{L_{\theta_0}(x_1, \dots, x_n)} \geq k$$
 Since $W = A \cup B$,
 on A also $\frac{L_{\theta_1}(x_1, \dots, x_n)}{L_{\theta_0}(x_1, \dots, x_n)} \geq k$

$$\therefore \int_A L_{\theta_1}(x_1, \dots, x_n) dx \geq k \int_A L_{\theta_0}(x_1, \dots, x_n) dx$$

So, W is constructed in such a way that on W L_{θ_1} of x_1, x_2, \dots, x_n upon L_{θ_0} of x_1, x_2, \dots, x_n is greater than equal to K . Since W is equal to A union B , on A also L_{θ_1} of x_1, x_2, \dots, x_n upon L_{θ_0} of x_1, x_2, \dots, x_n is greater than equal to K . Therefore, integration over A of L_{θ_1} of x_1, x_2, \dots, x_n , dx is greater than equal to integration over A of L_{θ_0} of x_1, x_2, \dots, x_n dx multiplied by K . Let us call it inequality 1.

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Now let us consider $W_1 = B \cup C$.
 Since size of $W_1 \leq \alpha$

$$\therefore \int_{B \cup C} L_{\theta_0}(x_1, \dots, x_n) dx \leq \alpha$$
 2 Since W is of size α

$$\int_{A \cup B} L_{\theta_0}(x_1, \dots, x_n) dx = \alpha$$

Now, let us consider W_1 is equal to B union C . Since size of W_1 is less than equal to α , therefore integration over B union C L_{θ_0} of x_1, x_2, \dots, x_n dx is less than equal

to alpha. And since W is of size alpha integration over A union B, $L_{\theta_0}(x_1, x_2, \dots, x_n)$ dx is equal to alpha. Therefore, together we get integration over A union B of $L_{\theta_0}(x_1, x_2, \dots, x_n)$ dx is greater than equal to the integration over B union C $L_{\theta_0}(x_1, x_2, \dots, x_n)$ dx or integration of over A, $L_{\theta_0}(x_1, x_2, \dots, x_n)$ dx is greater than equal to integration over C, $L_{\theta_0}(x_1, x_2, \dots, x_n)$ dx this happens because B is common to both and therefore we can take away the B part and we get this inequality.

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\therefore Together we get

$$\int_{A \cup B} L_{\theta_0}(x_1, \dots, x_n) dx \geq \int_{B \cup C} L_{\theta_0}(x_1, \dots, x_n) dx$$

or

$$\int_A L_{\theta_0}(x_1, \dots, x_n) dx \geq \int_C L_{\theta_0}(x_1, \dots, x_n) dx$$

②

Let us call it 2.

(Refer Slide Time: 12:59)

Now from (1) we have

$$\int_A L_{\theta_1}(x) dx \geq k \int_A L_{\theta_0}(x) dx$$

$$\geq k \int_C L_{\theta_0}(x_1, \dots, x_n) dx$$

③

Now, from 1, we have you can see from here that L_{θ_1} is greater than equal to K times L_{θ_0} . So, if I use little bit of shortcut let me write it as integration over A , L_{θ_1} of dx is greater than equal to K times integration over A L_{θ_0} of dx . And from 2 since we have integration over A , L_{θ_0} is greater than equal to integration over C of L_{θ_0} we can write this is greater than equal to K integration over C L_{θ_1} of $x_1, x_2, \dots, x_n dx$. So, let us call it 3.

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Now since on C

$$\frac{L_{\theta_1}(x_1, \dots, x_n)}{L_{\theta_0}(x_1, \dots, x_n)} < K$$

on C we have

$$\int_C L_{\theta_1}(x_1, \dots, x_n) dx < K \int_C L_{\theta_0}(x_1, \dots, x_n) dx$$

\therefore from (3) & (4)

$$\int_C L_{\theta_1}(x_1, \dots, x_n) dx < \int_A L_{\theta_1}(x_1, \dots, x_n) dx$$

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Now, since on C L_{θ_1} of x_1, x_2, \dots, x_n upon L_{θ_0} of x_1, x_2, \dots, x_n is less than K on C we have integration over C L_{θ_1} of $x_1, x_2, \dots, x_n dx$ is less than K times L_{θ_0} of $x_1, x_2, \dots, x_n dx$. So, let us call it 4.

So, from 3 and 4 from here we have integration over C , L_{θ_1} of x_1, x_2, \dots, x_n is less than this which in turn is less than this which in turn is less than this. Therefore, together we write the integration over C L_{θ_1} of $x_1, x_2, \dots, x_n dx$ is less than the integration over A , L_{θ_1} of $x_1, x_2, \dots, x_n dx$.

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$\therefore \int_{B \cup C} L_{\theta_1}(x_1, \dots, x_n) dx < \int_{A \cup B} L_{\theta_1}(x_1, \dots, x_n) dx$

Power of the CR W_1 < Power of CR W

\therefore We get that Power of $W >$ Power of W_1

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Therefore, the integration over B union C of $L_{\theta_1}(x_1, x_2, \dots, x_n) dx$ is less than integration over A union B of $L_{\theta_1}(x_1, x_2, \dots, x_n) dx$. And by definition we know this is the power of the critical region W_1 and this is the power of critical region W therefore, we get that power of W is greater than power of W_1 .

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\therefore The crux is to construct the MPCR of size α . we need to design the CR W such that on W $\frac{L_{\theta_1}(x_1, \dots, x_n)}{L_{\theta_0}(x_1, \dots, x_n)} \geq K$

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Therefore the crux is to construct the most powerful critical region of size alpha we need to design the critical region W such that on W $L_{\theta_1}(x_1, x_2, \dots, x_n)$ upon $L_{\theta_0}(x_1, x_2, \dots, x_n)$ is greater than equal to K .

Now, in the last class I have given you two exercises to find the most powerful critical region, to test λ is equal to λ_0 , against λ is equal to λ_1 for exponential distribution. So, let me relook at the problem.

(Refer Slide Time: 19:00)

Ex $\text{Exp}(\lambda)$: To test
 $H_0: \lambda = \lambda_0$
 vs.
 $H_1: \lambda = \lambda_1$
 \therefore To obtain the
 MPCR

$$\frac{(\lambda_1)^n e^{-\lambda_1 \sum x_i}}{(\lambda_0)^n e^{-\lambda_0 \sum x_i}} \geq k$$
 or
$$\left(\frac{\lambda_1}{\lambda_0}\right)^n e^{-(\lambda_1 - \lambda_0) \sum x_i} \geq k$$

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So, example, the exponential λ_2 test $H_0: \lambda$ is equal to λ_0 versus $H_1: \lambda$ is equal to λ_1 . Therefore, to obtain the most powerful critical region we have tried this that λ_1 to the power n , e to the power minus λ_1 sigma x_i upon λ_0 to the power n , e to the power minus λ_0 sigma x_i is greater than equal to K or λ_1 upon λ_0 whole to the power n , e to the power minus λ_1 minus λ_0 sigma x_i is greater than equal to K .

(Refer Slide Time: 20:09)

Or Taking natural log
 $-(\lambda_1 - \lambda_0) \sum x_i \geq \text{const } C_1$
Case 1 $\lambda_1 - \lambda_0 > 0$ or $\lambda_1 > \lambda_0$
 \therefore CR will be $\sum x_i \leq K_1$
Case 2 $\lambda_1 - \lambda_0 < 0$ i.e. $\lambda_1 < \lambda_0$
 \therefore CR is $\sum x_i \geq K_2$

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Or taking natural log minus lambda 1 minus lambda 0 sigma x i is greater than equal to some constant say C 1 which we get by manipulating the algebraic terms. But I am not interested in that I am interested in the relative values of lambda 1 and the lambda 0.

So, case 1, lambda 1 minus lambda 0 is greater than 0 or lambda 1 greater than lambda 0. In that case this term becomes negative with the minus sign therefore, the critical region will be sigma x i is less than equal to some constant say K 1. And case 2 lambda 1 minus lambda 0 is less than 0, that is lambda 1 is less than lambda 0 in that case this is going to be positive therefore, critical region is going to be sigma x i is greater than equal to some constant K 2. Up to this we have done in the last class.

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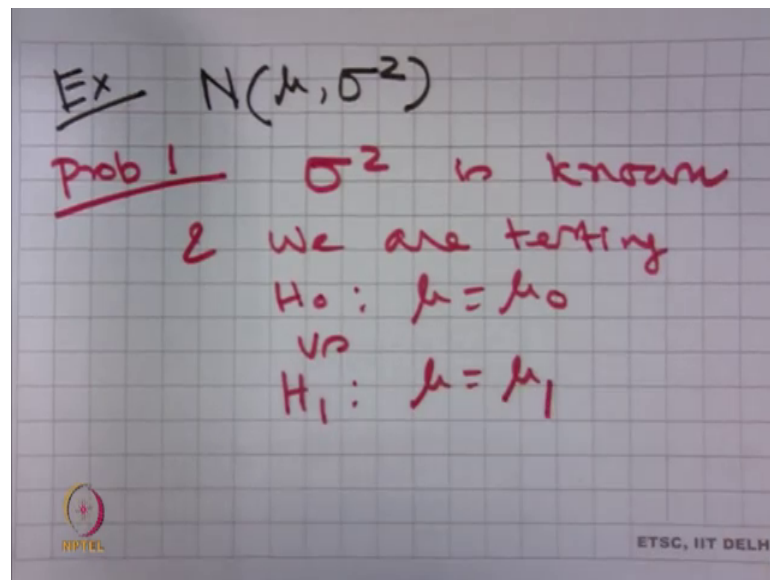
Typically $\sum x_i \sim \Gamma(\lambda_0, n)$
under H_0
Pmt if $\lambda_0 = \frac{1}{2}$
Then $\sum x_i \sim \Gamma(\frac{1}{2}, n) = \Gamma(\frac{1}{2}, \frac{2n}{2})$
 $\therefore \sum x_i \sim \chi^2_{(2n)}$
Hence if $\lambda_0 = \frac{1}{2}$ we can
get the cut off values or
the thresholds from χ^2 table.

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Typically, $\sum x_i$ is distributed as gamma with λ_0 and n under H_0 , but if λ_0 is equal to half, then $\sum x_i$ is distributed as gamma half comma n which is equal to gamma half comma $\frac{2n}{2}$ and therefore, $\sum x_i$ is distributed as chi square with $2n$ degrees of freedom. Hence if λ_0 is equal to half we can get the cutoff values or the thresholds from chi square table.

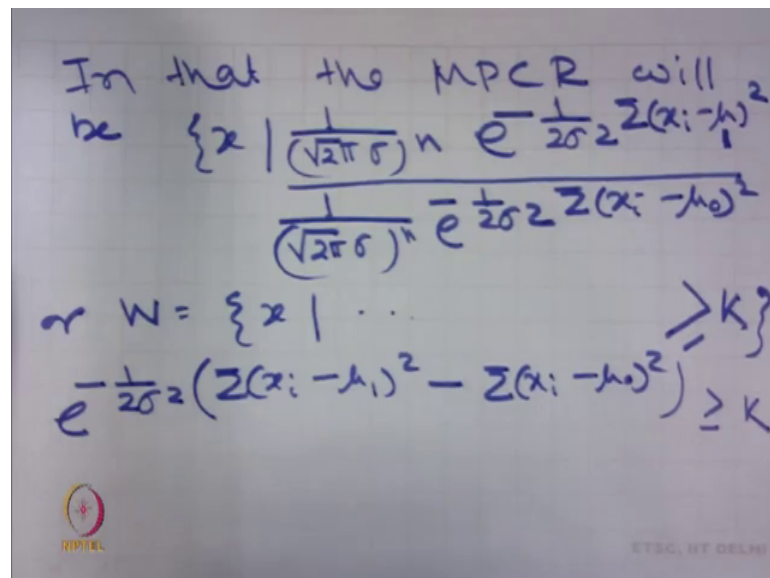
This is a very special case of exponential distribution. In general it will not work, but if we are testing with λ_0 is equal to half then we get from chi square table. Otherwise as I mentioned in the last class that we have to integrate or we have to do it numerically to find that constant, so that we can construct the most powerful critical region another example, that I did in the last class is normal μ comma σ^2 .

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Let us look at problem 1, sigma square is known and we are testing H_0 μ is equal to μ_0 versus H_1 μ is equal to μ_1 .

(Refer Slide Time: 25:05)



And we found that in that case, the most powerful critical region will be all those x such that $\frac{1}{\sqrt{2\pi}\sigma}$ whole to the power n , e to the power minus 1 upon $2\sigma^2$ sigma square sigma x_i minus μ_1 whole square upon μ_1 whole square upon $\frac{1}{\sqrt{2\pi}\sigma}$ whole to the power n , e to the power minus 1 upon $2\sigma^2$ sigma square. Sigma square is known for both of them and it is the same sigma x_i minus μ_0

whole square this ratio has to be greater than equal to K or W is equal to all those x, such that e to the power minus 1 upon 2 sigma square into sigma x i minus mu 1 whole square minus sigma x i minus mu naught whole square is greater than equal to K.

(Refer Slide Time: 26:43)

$$\text{or } -\frac{1}{2\sigma^2} \left(\sum x_i^2 - 2\mu_1 \sum x_i + n\mu_1^2 - \sum x_i^2 + 2\mu_0 \sum x_i - n\mu_0^2 \right) \geq K_1$$

$$\text{or } -\frac{1}{2\sigma^2} \left(n(\mu_1^2 - \mu_0^2) - 2\sum x_i (\mu_1 - \mu_0) \right) \geq K_1$$

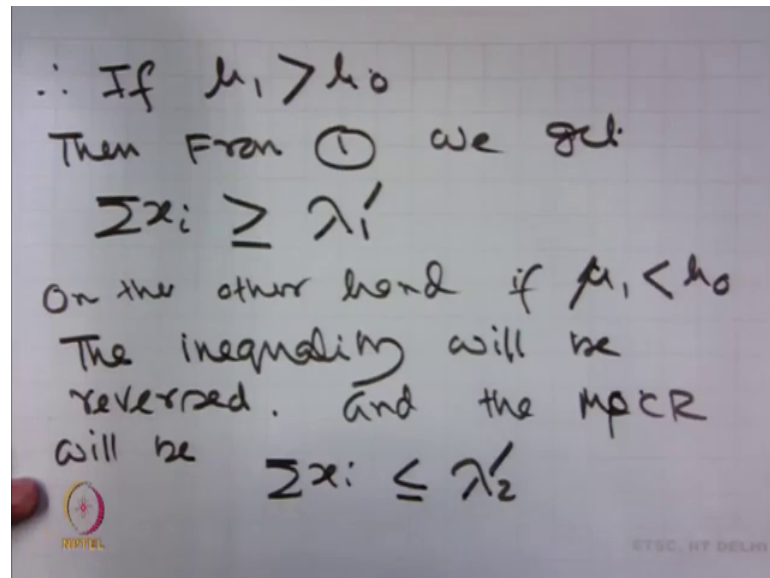
$$\text{or } n(\mu_1^2 - \mu_0^2) - 2\sum x_i (\mu_1 - \mu_0) \geq K_1$$

$$\text{or } n(\mu_1^2 - \mu_0^2) - C_1 \leq \frac{C_1}{2\sum x_i} (\mu_1 - \mu_0)$$

Or by taking natural log minus 1 upon 2 sigma square into sigma x i square minus 2 mu 1 sigma x i plus m mu 1 square minus sigma x i square plus 2 mu 0 sigma x i minus n mu naught square is greater than equal to some constant say K 1 or minus 1 upon 2 sigma square. Now, this gets cancelled and we have n times mu 1 square minus mu naught square minus 2 sigma x i into mu 1 minus mu naught is greater than equal to K 1 or n times mu 1 square minus mu naught square minus 2 sigma x i mu 1 minus mu naught is less than equal to some constant say C 1.

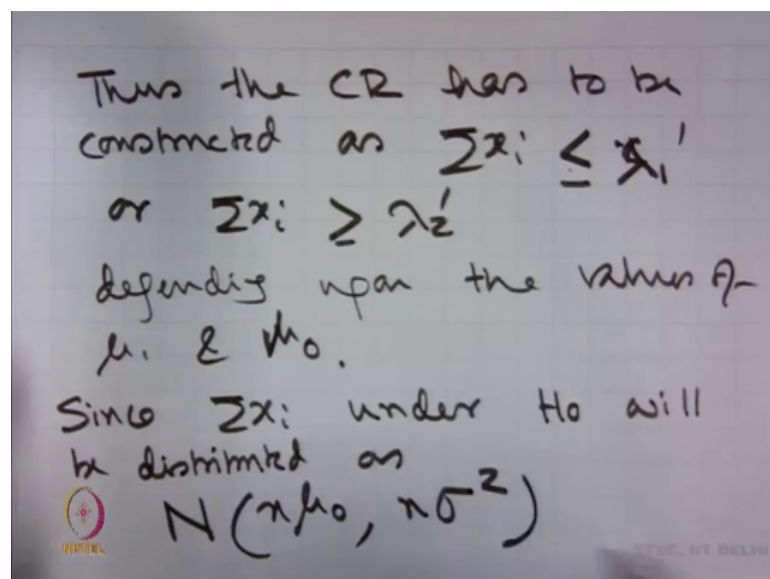
This comes because this minus sign is there therefore, the inequality will be reversed and when we multiply by 2 sigma square we get a different constant C 1 or in mu 1 square minus mu naught square minus C 1 is less than equal to 2 sigma x i into mu 1 minus mu naught. So, let us call it 1.

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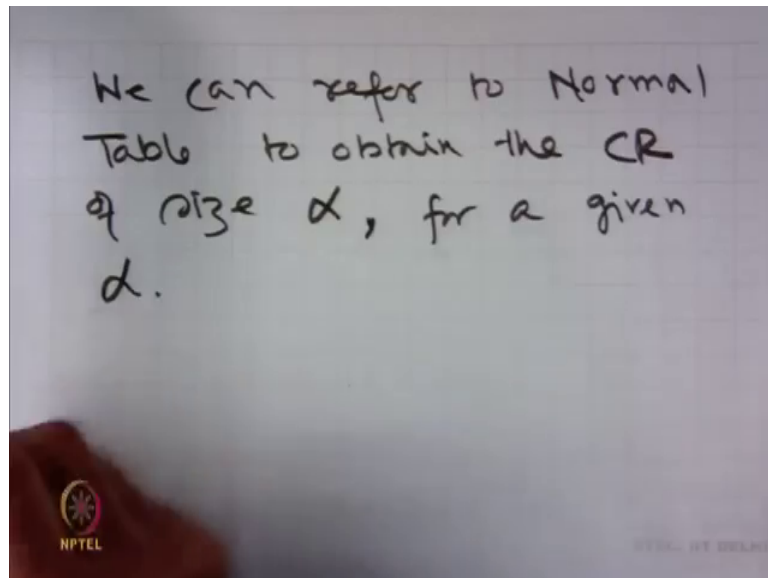
Therefore, if μ_1 is greater than μ_0 , then from 1 we get if we divide both the sides by $\mu_1 - \mu_0$ which is positive we get that $\sum x_i$ has to be greater than equal to some constant λ_1' . As you can see from here $\mu_1 - \mu_0$ is positive therefore, if I divide both the sides then we get $\sum x_i$ has to be greater than some constant. On the other hand if μ_1 is less than μ_0 , then if we divide both sides by $\mu_1 - \mu_0$. The inequality will change will be reversed and the MPCR will be $\sum x_i$ less than equal to some constant λ_2' .

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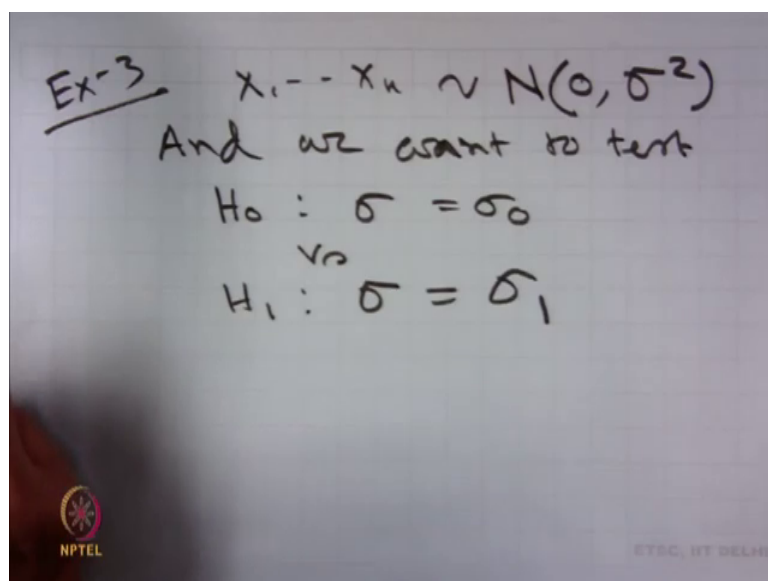
Thus, the critical region has to be constructed as σx_i less than equal to some constant λ_1 for σx_i greater than equal to some constant λ_2 prime, depending upon the values of μ_1 and μ naught. Since σx_i under H naught will be distributed as normal with $N \mu$ naught and the variance is going to be $n \sigma$ square.

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We can refer to normal table to obtain the critical region of size alpha, for a given alpha.

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Now, suppose we want to test for sigma. Suppose we have the following problem x_1, x_2, \dots, x_n are from normal with mean 0 and variance σ^2 . And we want to test $H_0: \sigma^2 = \sigma_0^2$ versus $H_1: \sigma^2 = \sigma_1^2$. Note that we have kept the mean to be 0. This is only for ease of calculation if that normal distribution has the mean known which is μ then instead of x_1, x_2, \dots, x_n we can always look at y_1, y_2, \dots, y_n , where $y_i = x_i - \mu$ and therefore, we will come back to normal $0, \sigma^2$.

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By NP Lemma the MPCR will be

$$\frac{\left(\frac{1}{\sqrt{2\pi}\sigma_1}\right)^n e^{-\frac{1}{2\sigma_1^2} \sum x_i^2}}{\left(\frac{1}{\sqrt{2\pi}\sigma_0}\right)^n e^{-\frac{1}{2\sigma_0^2} \sum x_i^2}} \geq K$$

or $\left(\frac{\sigma_0}{\sigma_1}\right)^n e^{-\frac{1}{2} \left(\frac{1}{\sigma_1^2} - \frac{1}{\sigma_0^2}\right) \sum x_i^2} \geq K$

$$\left(\frac{\sigma_0}{\sigma_1}\right)^n (\sigma_0 - \sigma_1) - \frac{1}{2} \left(\frac{1}{\sigma_1^2} - \frac{1}{\sigma_0^2}\right) \sum x_i^2 \geq K_1$$

By Neyman-Pearson Lemma the most powerful critical region will be 1 upon root over 2 pi sigma 1 whole to the power n into e to the power minus 1 upon 2 sigma 1 square sigma x i square upon 1 over root over 2 pi sigma naught whole to the power n, e to the power minus 1 upon 2 sigma naught square sigma x i square.

This has to be greater than equal to K or sigma naught upon sigma 1 whole to the power n, e to the power minus half; 1 upon sigma 1 square minus 1 upon sigma naught square into sigma x i square is greater than equal to K. Or taking log in log sigma naught minus sigma 1 minus half 1 upon sigma 1 square minus 1 upon sigma naught square sigma x i square is greater than equal to some constant which will come through log operation on K. So, let me call it K_1 .

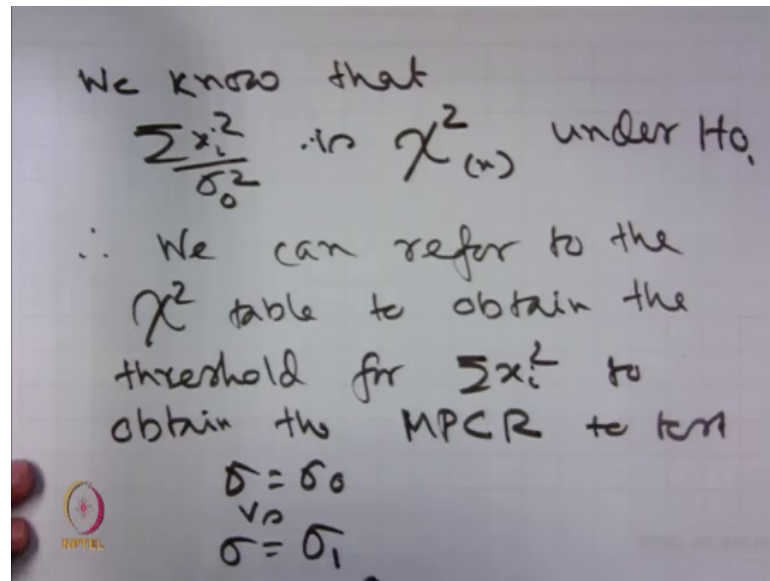
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$$\begin{aligned}
 & \text{Or } n \ln(\sigma_0 - \sigma_1) - K_1 \\
 & \qquad \geq \frac{1}{2} \left(\frac{1}{\sigma_1^2} - \frac{1}{\sigma_0^2} \right) \sum x_i^2 \\
 & \text{Or } \sum x_i^2 \left(\frac{\sigma_0^2 - \sigma_1^2}{2\sigma_0^2\sigma_1^2} \right) \leq \frac{n \ln(\sigma_0 - \sigma_1) - K_1}{C_1} \\
 & \therefore \text{If } \sigma_0 > \sigma_1 \\
 & \quad \text{The MPCR will be } \sum x_i^2 \leq \lambda_1 \\
 & \text{If } \sigma_0 < \sigma_1 \\
 & \therefore \text{The MPCR will be } \sum x_i^2 \geq \lambda_2
 \end{aligned}$$

Or $n \ln(\sigma_0 - \sigma_1) - K_1$ is greater than equal to half of $\left(\frac{1}{\sigma_1^2} - \frac{1}{\sigma_0^2}\right) \sum x_i^2$. Or $\sum x_i^2 \left(\frac{\sigma_0^2 - \sigma_1^2}{2\sigma_0^2\sigma_1^2}\right)$ is less than equal to $\frac{n \ln(\sigma_0 - \sigma_1) - K_1}{C_1}$. Let me call it C_1 .

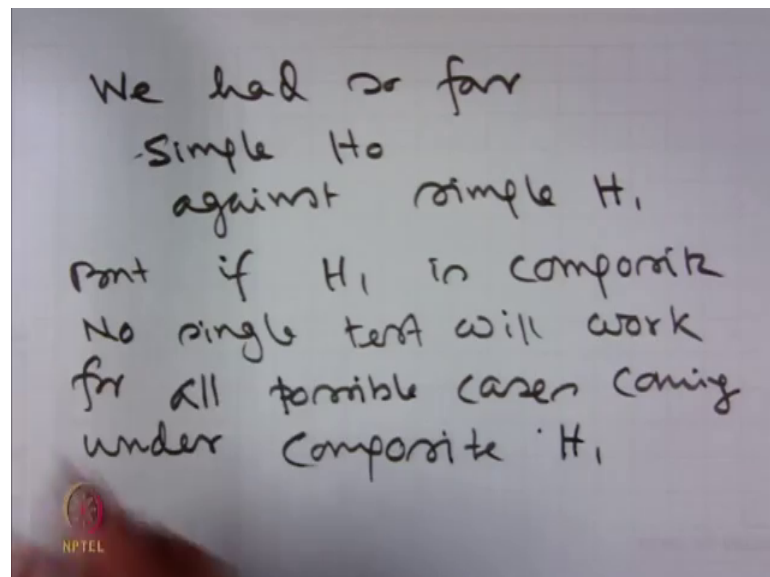
Therefore, if $\sigma_0 > \sigma_1$ then this term becomes positive the MPCR will be $\sum x_i^2 \leq \lambda_1$. And if $\sigma_0 < \sigma_1$ then because of the negative sign the inequality will be reversed and therefore, the MPCR will be $\sum x_i^2 \geq \lambda_2$.

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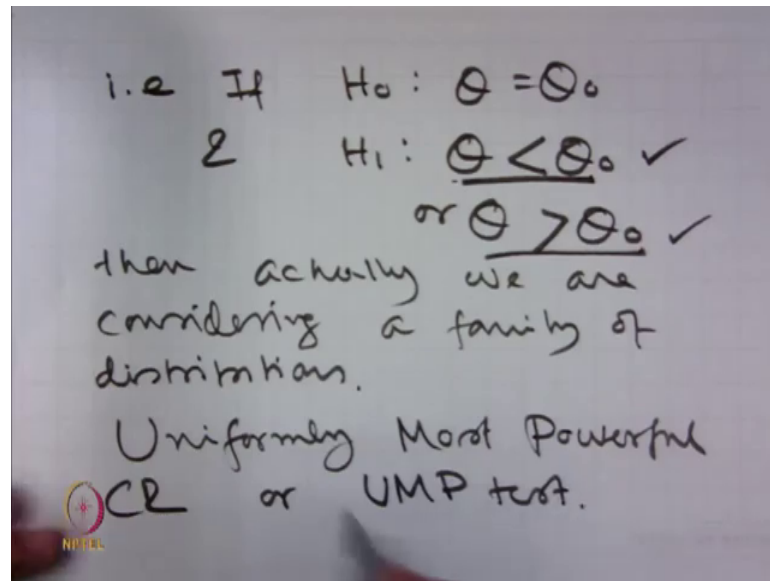
Now, we know that $\sum x_i^2 / \sigma_0^2$ is chi square with n degrees of freedom under H_0 . Therefore, we can refer to the chi square table to obtain the threshold for $\sum x_i^2$ to obtain the most powerful critical region to test $\sigma = \sigma_0$ versus $\sigma = \sigma_1$.

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In all the examples that I have given so far we had so far simple H_0 against simple H_1 . But if H_1 is composite, no single test will work for all possible cases coming under composite H_1 .

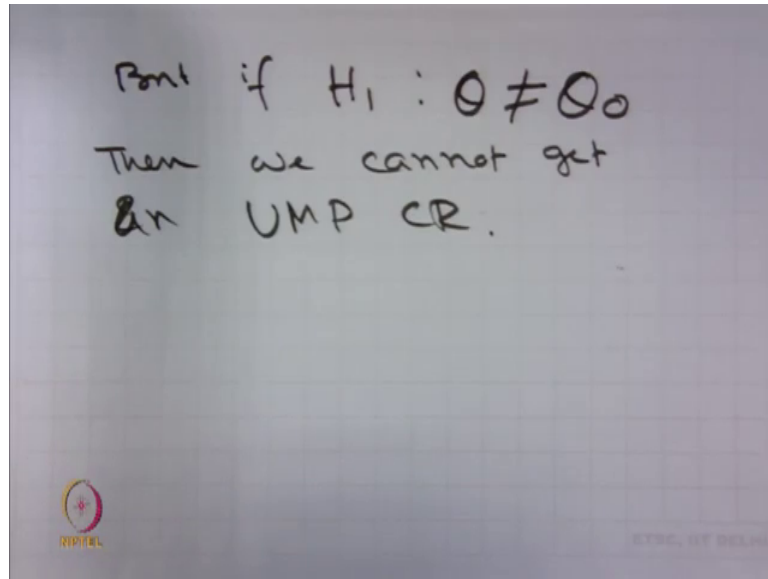
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That is, if H_0 is $\theta = \theta_0$ and H_1 is $\theta < \theta_0$ or $\theta > \theta_0$ then actually we are considering a family of distributions.

But in some cases as we have seen in all the examples, when we are creating the most powerful critical region actually the value of θ_1 has not come into picture in constructing the MPCR. Therefore, if we are testing alternatives of this type that is only one sided inequality then the MPCR that we have obtained there that we will work for all the family of distribution satisfying this or satisfying this. And therefore, the most powerful critical region obtained in that way can be called uniformly most powerful critical region or corresponding test is called uniformly most powerful test.

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But if H_1 is of the form $\theta \neq \theta_0$ then we cannot get a uniformly most powerful critical region in general.

Ok students, with that I conclude my series of lectures on statistical inference. Over the last 21 lectures we started with some basic probability distributions and we have stressed upon several topics including order statistics, theory of estimation and of course, testing of hypothesis. I hope you found that course useful and if you have any doubts or queries please do not hesitate to contact me or my TAs to clarify your doubts.

Thank you very much.