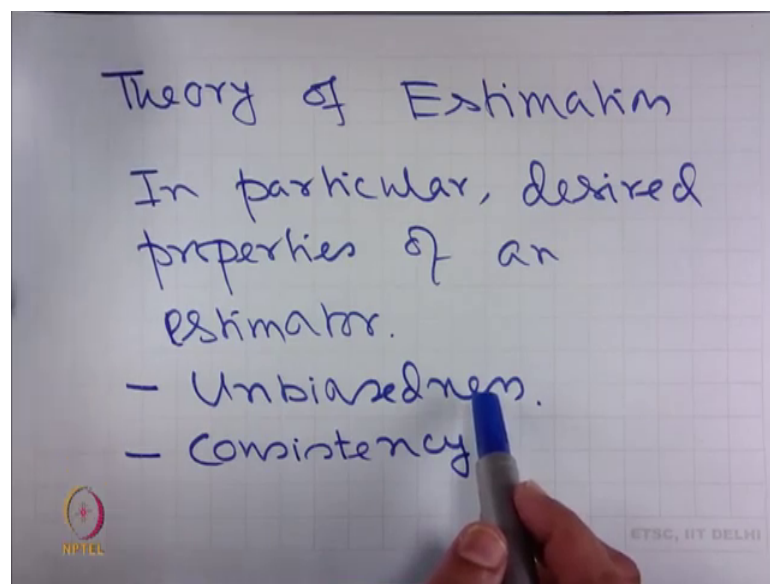


**Statistical Inference**  
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**Department of Mathematics**  
**Indian Institute of Technology, Delhi**

**Lecture – 13**  
**Statistical Inference**

Welcome students to the MOOCS lecture series, on Statistical Inference. This is lecture number 13 and you know that we are currently discussing Theory of Estimation.

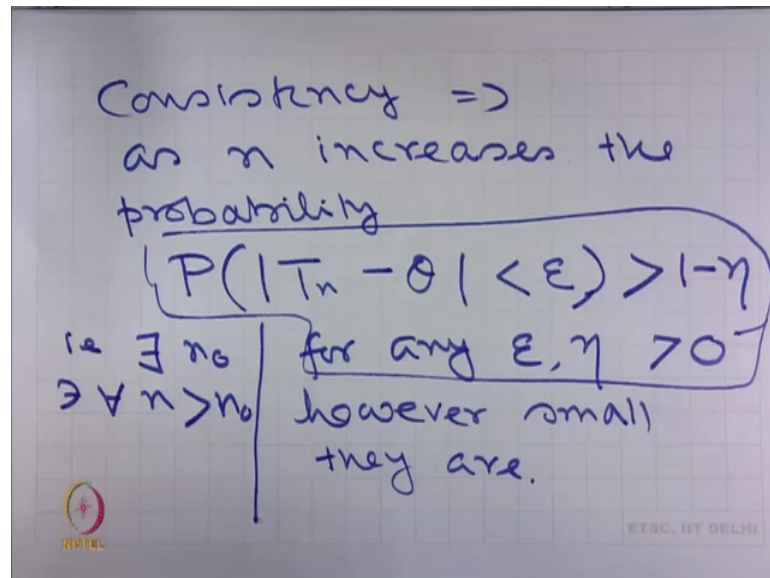
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In particular we are looking at properties of or desired properties of an estimator, in this respect, we have discussed two properties namely unbiasedness and consistency and estimator  $T = T(x_1, x_2, \dots, x_n)$  is said to be unbiased to estimate parameter  $\theta$ , if the expected value of  $T$  is  $\theta$ .

And we have seen that there are many unbiased estimators, but all of them are not consistent. Why consistency is important?

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Consistency implies as  $n$  increases the probability that is  $P$  of modulus of  $T_n$  minus  $\theta$  less than  $\epsilon$  becomes greater than  $1 - \eta$  for any  $\epsilon$  and  $\eta$  greater than  $0$ ; however, small they are or in other words there exist  $N$  such that for all  $N$  greater than  $N$  such that this property holds; that means, after certain number, if we choose samples of size beyond that, then we know that the probability that the estimator  $T_n$  will come arbitrarily close to the parameter  $\theta$  that it is estimating and that probability is going to be as close to one as possible.

And we have seen sufficient conditions for consistency that if the expected value of  $T_n$  converges to  $\theta$  and the variance of  $T_n$  goes to  $0$  as  $N$  goes to infinity, then we know that, that particular  $T_n$  is going to be consistent for  $\theta$ .

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Natural question is:  
If  $T_n$  is unbiased for  $\theta$ , will  $g(T_n)$ , where  $g$  is a function, be unbiased for  $g(\theta)$ ?

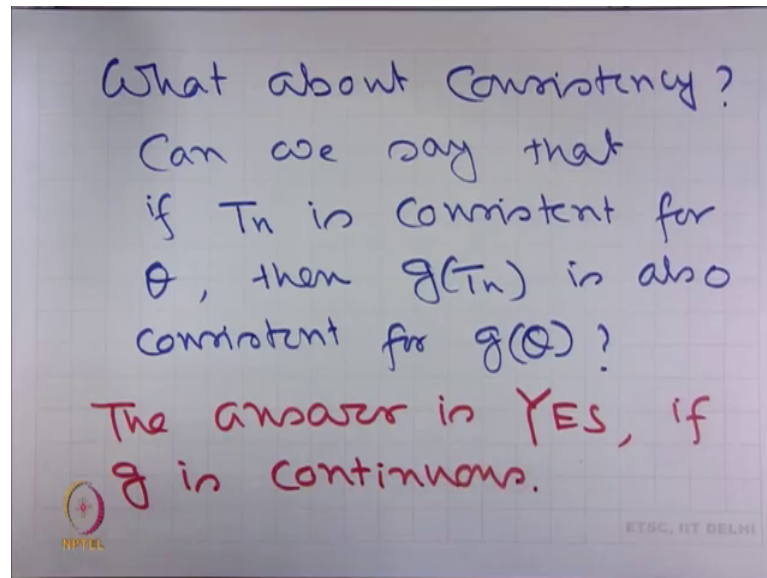
The answer is NO  
e.g. We know  $E(X^2)$   
 $= \text{Var}(X) + (E(X))^2$   
for any variable  $X$ .

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Natural question is, if  $T_n$  is unbiased for  $\theta$  will  $g$  of  $T_n$ , where  $g$  is a function be unbiased for  $g$   $\theta$ . It is a very natural question, because many a time we try to estimate not exactly the  $\theta$ , but some function of  $\theta$ . Can we use the unbiasedness property to get an unbiased estimator for  $g$   $\theta$ ?

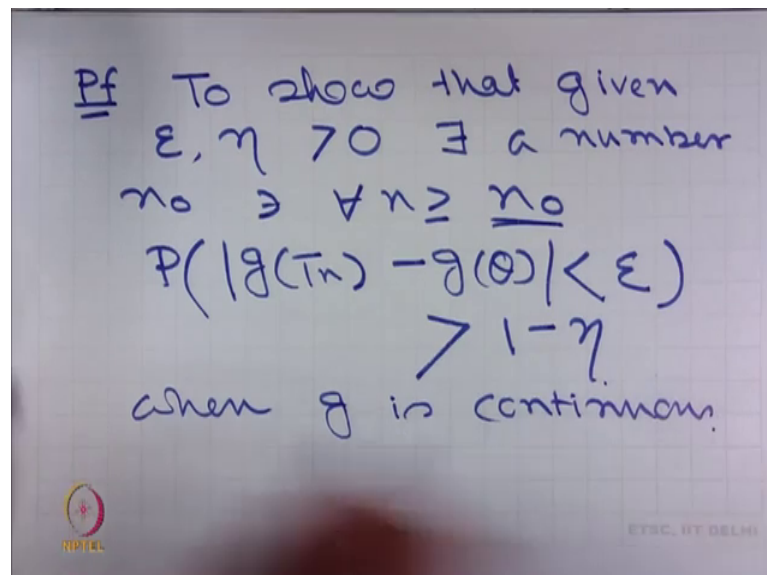
Unfortunately the answer is no. For example, we know expected value of  $X$  square is equal to variance of  $X$  plus expected value of  $X$  whole square for any variable  $X$ . Therefore, if we focus on  $T_n$ , an expected value of  $T_n$  is  $\theta$ , we see that expected value of  $T_n$  square is going to be variance of  $T_n$  plus  $\theta$  square and therefore, we can see that  $X$  square is not unbiased for  $\theta$  square, because there is the bias in the form of variance of  $X$ . What about consistency?

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Can we say that if  $T_n$  is consistent for  $\theta$  then  $g$  of  $T_n$  is also consistent. For  $g$   $\theta$  and the answer is yes, if  $g$  is continuous proof to show that given  $\epsilon$  and  $\eta$  greater than 0, there exist a number  $n$  naught, such that for all  $n$  greater than equal to  $n$  naught probability.

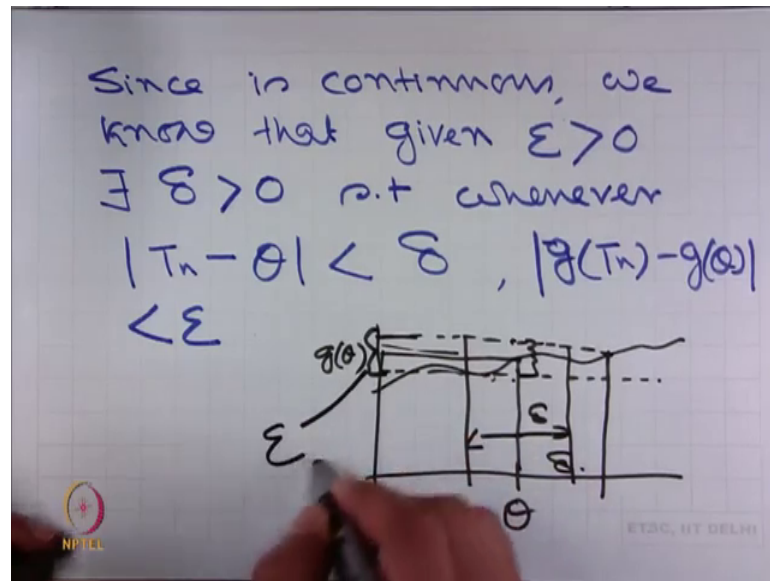
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Modulus of  $g T_n$  minus  $g \theta$  less than  $\epsilon$  is greater than  $1 - \eta$ .

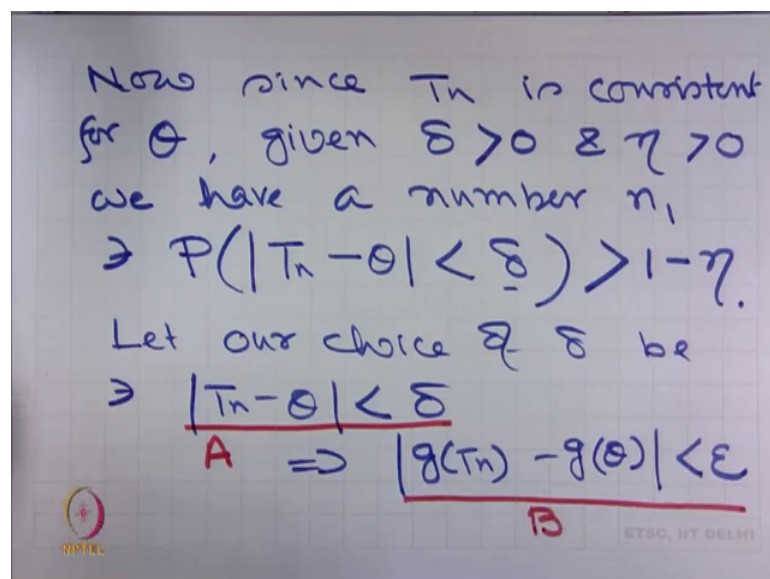
So, we have to identify such a  $n$  naught when  $g$  is continuous, since  $g$  is continuous by definition of continuity.

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We know that given epsilon greater than 0 there exist delta greater than 0, such that whenever modulus of  $T_n$  minus theta is less than delta modulus of  $g(T_n)$  minus  $g(\theta)$  is less than epsilon or suppose, this is theta and  $g$  is a continuous function and this is  $g(\theta)$  and suppose, we create an epsilon neighborhood of around the  $g(\theta)$  then we see that for all  $n$ , in this neighborhood sorry, for all values of the variable say  $X$  in the neighborhood. Let us call it delta, which may be like this. So, let us call it delta we can see that modulus of  $g(\theta)$  minus  $g(T_n)$  is within the limit epsilon. So, this is our epsilon.

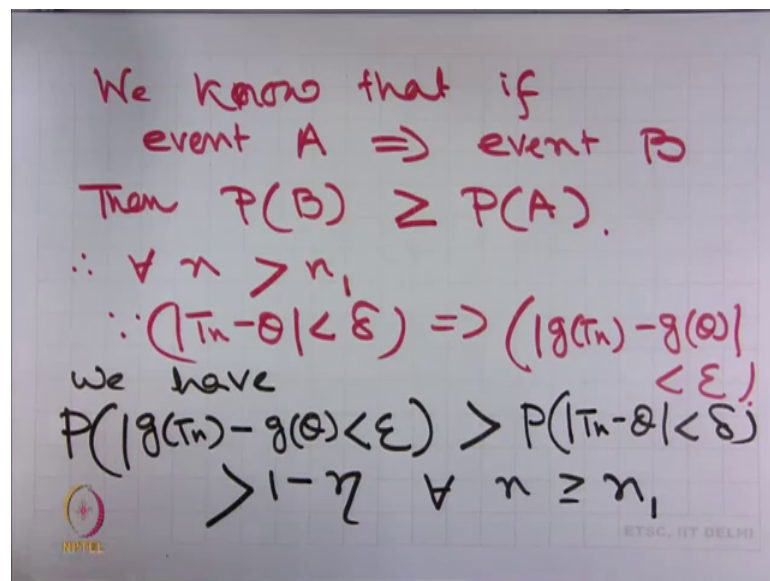
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Now, since  $T_n$  is consistent for  $\theta$  given  $\delta > 0$  and  $\eta > 0$ , we have a number  $n_1$  such that probability modulus of  $T_n$  minus  $\theta$ , less than  $\delta$  is greater than  $1 - \eta$ . Let our choice of  $\delta$  be such that  $T_n$  minus  $\theta$  less than  $\delta$  implies  $g(T_n) - g(\theta)$  less than  $\epsilon$ .

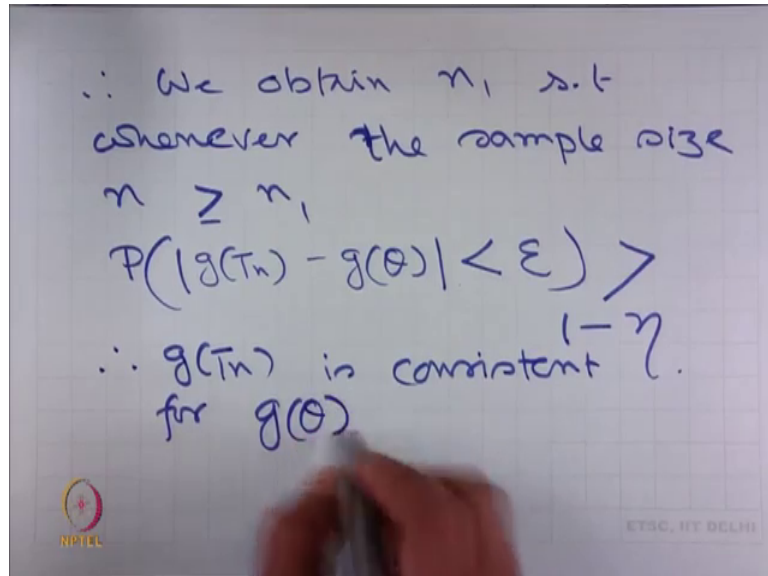
So, let us call this event A and let us call this event B.

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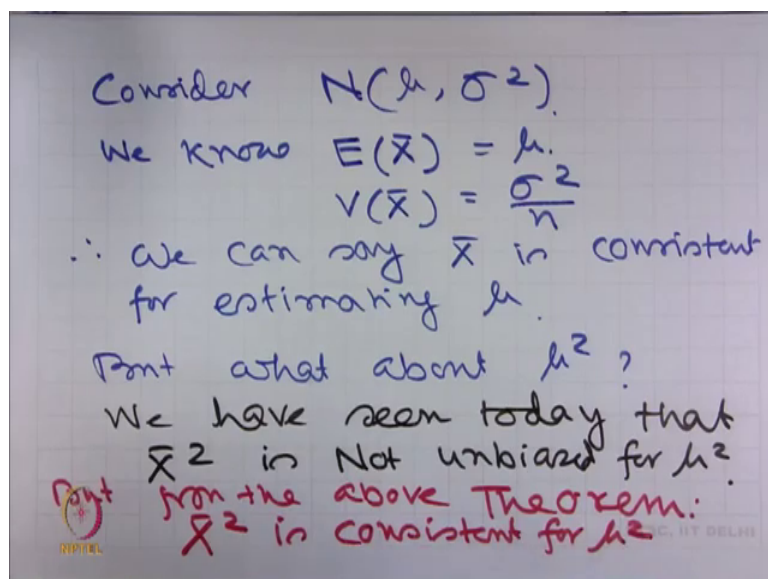
We know that, if event A implies event B then probability of B is greater than equal to probability of A. Therefore, for all  $n$  greater than  $n_1$  since  $T_n$  minus  $\theta$  less than  $\delta$  implies  $g(T_n) - g(\theta)$  less than  $\epsilon$ , we have probability  $g(T_n) - g(\theta)$  less than  $\epsilon$  is greater than probability  $T_n$  minus  $\theta$  less than  $\delta$ , which is greater than  $1 - \eta$  for all  $n$  greater than equal to  $n_1$ .

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Therefore, we obtained  $n_1$ , such that whenever this sample size  $n$  is greater than equal to  $n_1$  the probability  $g(T_n) - g(\theta) < \epsilon$  is greater than  $1 - \eta$  therefore,  $g(T_n)$  is consistent for  $g(\theta)$ . What is the advantage consider?

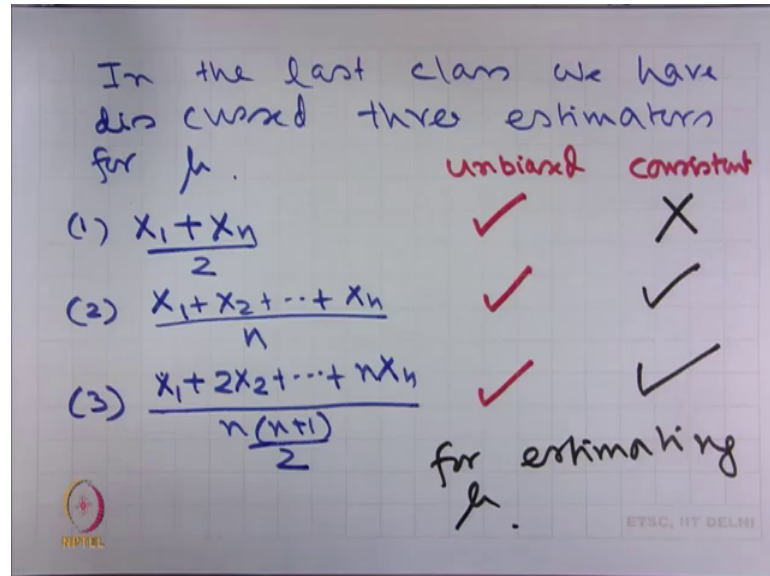
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Normal  $\mu$   $\sigma^2$ , we know that the expected value of  $\bar{X}$  is equal to  $\mu$  variance of  $\bar{X}$  is equal to  $\sigma^2/n$ . Therefore, we can say  $\bar{X}$  is consistent for estimating  $\mu$ , this we get from the theorem that approved in the previous class, but what about  $\mu^2$ ?

We have seen today that  $\bar{X}$  is not unbiased for  $\mu$ , but from the above theorem what we get that  $\bar{X}$  is consistent for  $\mu$ .

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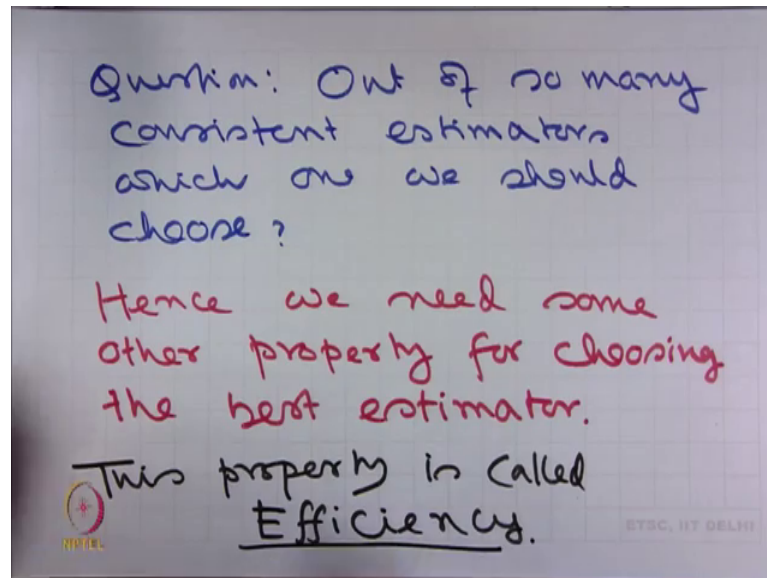


In the last class, we have discussed three estimators for  $\mu$  one is,  $\frac{X_1 + X_n}{2}$ , second is  $\frac{X_1 + X_2 + \dots + X_n}{n}$  and third one is  $\frac{X_1 + 2X_2 + \dots + nX_n}{\frac{n(n+1)}{2}}$ .

If we look at these three estimators, we found that this is unbiased, this is unbiased and this is also unbiased in estimating  $\mu$ , but this is not consistent, but this is consistent and this is consistent for estimating  $\mu$ , which you know is the population mean. Therefore, there is a question; we have many unbiased estimators, among them, some of them are consistent and that can be a large number, because we can choose estimators in such a way. So, that they become consistent, that is they are various reduces to 0 as  $n$  goes to infinity. Therefore, the question is out of so, many consistent estimators which one we should choose?



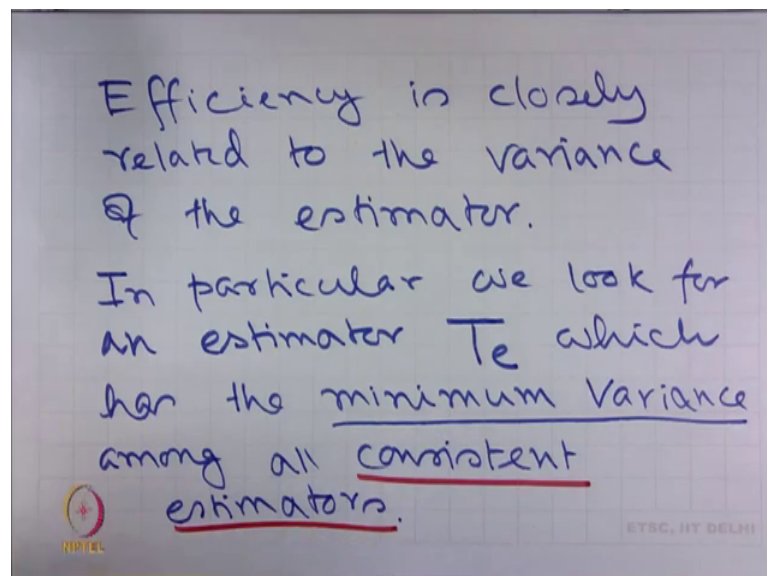
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This is a very pertinent question, because we have been given several estimators, we need to choose the best one among them and we have already observed that, just unbiasedness and consistency, they do not give us a single choice of estimator to estimate a particular parameter  $\theta$ . Just that in the previous case, we have used  $\theta$  is equal to  $\mu$  the population mean.

Therefore, we need some other property for choosing the best estimator, this property is called efficiency.

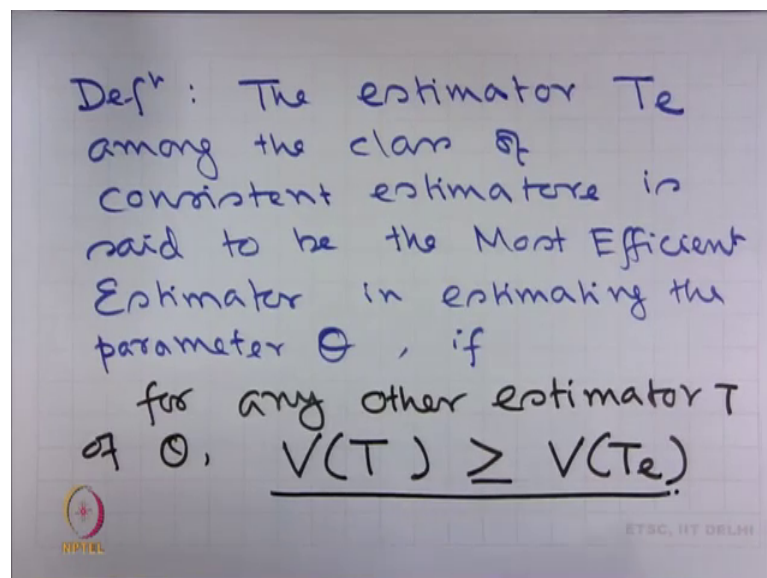
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Now, efficiency is closely related to the variance of the estimator, if the variance is less; that means, if the square of the deviation from the expected value is less then the advantage is that we are more sure that we are closer to the estimate and therefore, the more efficient an estimator is the less is its variance.

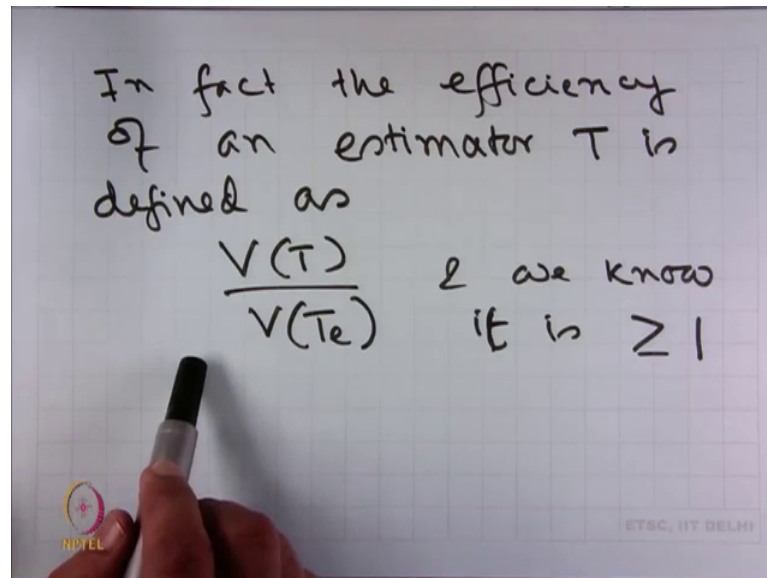
In particular, we look for an estimator  $T_e$ , which has the minimum variance among all consistent estimators and it makes sense that we are only searching minimum variance among the consistent estimators, because if an estimator is not consistent, then we cannot say that it is coming close and close to the parameter  $\theta$  as  $n$  goes to infinity, because this variance is not decreasing as  $n$  increases.

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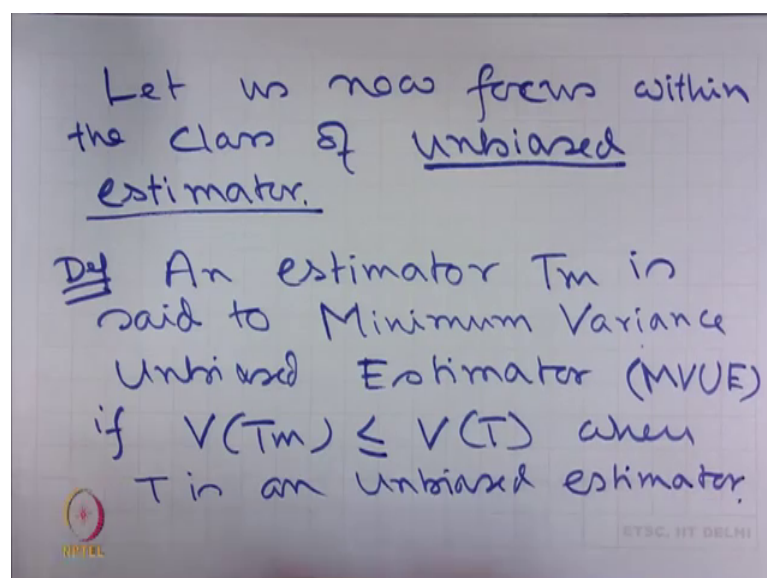
The estimator  $T_e$  among the class of consistent, estimators is said to be the most efficient estimator in estimating the parameter  $\theta$ . If for any other estimator  $T$  of  $\theta$ , variance of  $T$  is greater than; equal to variance of  $T_e$  that is the estimator  $T$ , whose variance is the minimum among all other consistent estimators that is called the minimum variance estimator or most efficient estimator.

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In fact, the efficiency of an estimated  $T$  is defined as variance of  $T$  divided by variance of  $T_e$  and we know it is greater than equal to 1, it will be 1, if variance of  $T$  is same as variance of  $T_e$  otherwise, as  $T_e$  as the minimum variance, variance of  $T$  has to be greater than variance of  $T_e$  and therefore, its efficiency has to be greater than equal to 1 for this class.

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Let us now focus within the class of unbiased estimator and estimator.  $T_m$  is said to the minimum variance unbiased estimator that is MVUE, if variance of  $T_m$  is less than

equal to variance of  $T$ , where  $T$  is an unbiased estimator that is in the class of unbiased estimator, we are looking for an estimated  $T_m$ , whose variance is minimum.

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Consider  $\frac{x_1 + \dots + x_n}{n} = T_1$

$\&$   $\frac{x_1 + 2x_2 + \dots + nx_n}{\frac{n(n+1)}{2}} = T_2$

Let us check their variances:

$$V(T_1) = \frac{\sigma^2}{n}$$

$$V(T_2) = \frac{1}{\left(\frac{n(n+1)}{2}\right)^2} (1^2 + 2^2 + \dots + n^2) \sigma^2$$

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Consider  $X_1 + X_n$  by  $n$  and  $X_1 + 2X_2 + \dots + nX_n$  by  $n + 1$  by  $2$ . Let us call them  $T_1$  and  $T_2$  respectively. We know both are unbiased for  $\mu$ , we know both are consistent. So, let us check they are variances, variance of  $T_1$  as we know is  $\sigma^2$  by  $n$  and variance of  $T_2$  is equal to, we have calculated in other day, but let me do it again. This is  $1$  upon  $n$  into  $n + 2$  by  $2$  whole square into  $1$  square plus  $2$  square up to  $n$  square  $\sigma^2$ , which is equal to  $\sigma^2$  into  $n$  into  $n + 1$  into  $2$  by  $3$ .

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$$\begin{aligned} &= \sigma^2 \times \frac{n(n+1)(2n+1)}{6} \\ &\quad \frac{1}{\left(\frac{n(n+1)}{2}\right)^2} \\ &= \sigma^2 \times \frac{2}{3} \times \frac{2n+1}{n(n+1)} \\ &= \frac{2\sigma^2}{3} \times \frac{2n+1}{n(n+1)}. \end{aligned}$$

Question is which one is minimum  $V(T_1)$  or  $V(T_2)$

This is 4, which goes to the numerator 6 is there into 2 n plus 1 upon n into n plus 1 is equal to 2 sigma square by 3 into 2 n plus 1 upon n into n plus 1.

Question is, which one is minimum variance of T 1 or variance of T 2?

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If we take the ratio

$$\frac{V(T_1)}{V(T_2)} = \frac{\frac{\sigma^2}{n}}{\frac{2\sigma^2}{3} \times \frac{2n+1}{n(n+1)}}$$
$$= \frac{3}{2} \times \frac{n+1}{2n+1} = \frac{3n+3}{4n+2}$$

$\therefore$  when  $n=1 \rightarrow \frac{1}{1}$   
 $n=2 \rightarrow \frac{9}{10}$   
 $n=3 \rightarrow \frac{12}{14}$

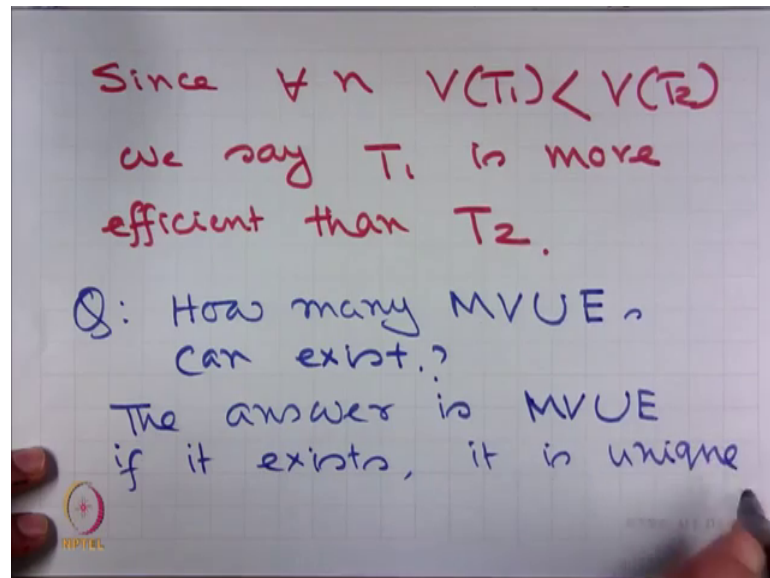
if  $n > 1$  then ratio  $< 1$

$\therefore n > 1 \Rightarrow \frac{V(T_1)}{V(T_2)} < 1$

If we take the ratio variance of T 1 upon variance of T 2, this is equal to sigma square by n upon 2 sigma square by 3 into 2 n plus 1 upon n into n plus 1 is equal to 3 by 2 n plus 1 upon 2 n plus 1 is equal to 3 n plus 3 upon 4 n plus 2.

Therefore when  $n$  is equal to 1, the ratio is 6 by 6 is equal to 1,  $n$  is equal to 2. The ratio is 9 by 10,  $n$  is equal to 3. The ratio is 12 by 14 and we see that, if  $n$  is greater than 1, then ratio is less than 1 and as  $n$  increases. We can, we can see that the numerator is increased by 3, but denominator is being increased by 4 therefore, for all  $n$  greater than 1 variance of  $T_1$  upon variance of  $T_2$  is less than 1 or in other words.

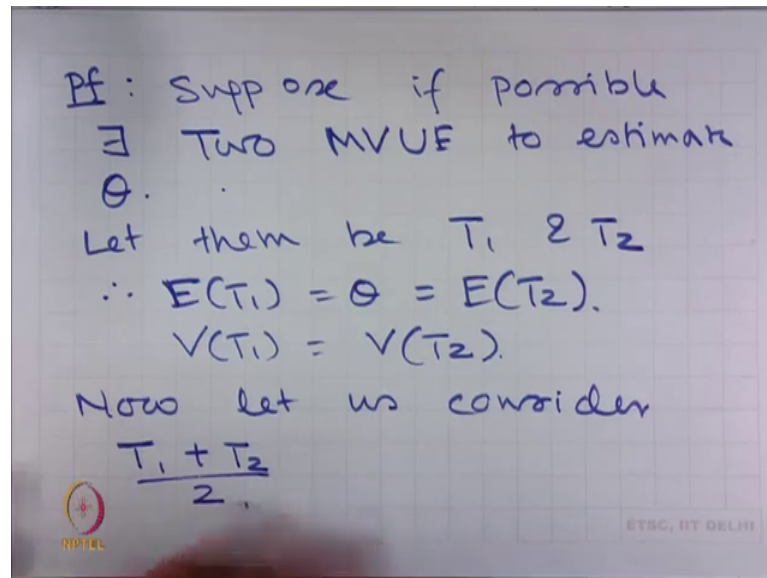
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Since, for all  $n$  variance of  $T_1$  is less than variance of  $T_2$ . We say  $T_1$  is more efficient than  $T_2$ .

Question, the obvious question is how many minimum variance unbiased estimators can exist? Can there be more than one minimum variance unbiased estimator to estimate a parameter  $\theta$ ? The answer is MVUE; if it exist it is unique proof.

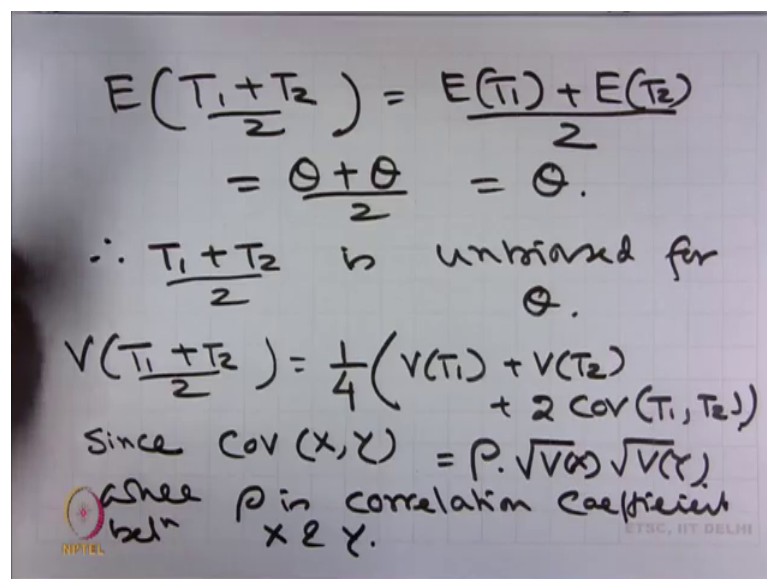
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Suppose, if possible there exist 2 MVUE to estimate theta. Let them be  $T_1$  and  $T_2$  therefore, expected value of  $T_1$  is equal to theta is equal to expected value of  $T_2$  and variance of  $T_1$  is equal to variance of  $T_2$ .

Now, let us consider  $T_1$  plus  $T_2$  by 2, expected value of  $T_1$  plus  $T_2$  by 2 is equal to expected value of  $T_1$  plus.

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Expected value of  $T_2$  upon 2 by linearity of expectation is equal to theta plus theta by 2 is equal to theta. Therefore,  $T_1$  plus  $T_2$  by 2 is unbiased for theta and variance of  $T_1$

plus  $T_2$  by 2 is equal to 1 by 4 times variance of  $T_1$  plus variance of  $T_2$  plus 2 times covariance of  $T_1 T_2$ .

Since, covariance of  $X$  comma  $Y$  is equal to  $\rho$  times root over variance of  $X$  root over variance of  $Y$ , where  $\rho$  is correlation coefficient between  $X$  and  $Y$ .

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We find:

$$V\left(\frac{T_1+T_2}{2}\right) = \frac{1}{4} \left( V(T_1) + V(T_2) + 2\rho\sqrt{V(T_1)}\sqrt{V(T_2)} \right)$$

$\therefore V(T_1) = V(T_2)$   
as both are MVUE  
by assumption

$$V\left(\frac{T_1+T_2}{2}\right) = \frac{1}{4} \left( 2V(T_1) + 2\rho V(T_1) \right)$$

$$= \frac{1}{4} \left( 2V(T_1) + 2\rho V(T_1) \right)$$

We find variance of  $T_1$  plus  $T_2$  by 2 is equal to 1 by 4 variance of  $T_1$  plus variance of  $T_2$  plus 2 times  $\rho$  root over variance of  $T_1$  root over variance of  $T_2$ . Since, variance of  $T_1$  is equal to variance of  $T_2$  as both are MVUE by assumption variance of  $T_1$  plus  $T_2$  by 2 is equal to 1 by 4 into variance of  $T_1$  plus variance of  $T_1$  replacing variance of  $T_2$  by variance of  $T_1$  plus 2  $\rho$ . This is root over variance of  $T_1$ ; this is root over variance of  $T_2$ . So, I can write it as variance of  $T_1$  is equal to 1 by 4 times two variance of  $T_1$  plus 2  $\rho$  into variance of  $T_1$ .



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$$\begin{aligned} &= \frac{1}{2}(V(T_1) + \rho V(T_1)) \\ &= V(T_1) \frac{1+\rho}{2} \checkmark \\ \text{Now } \frac{T_1+T_2}{2} &\text{ is unbiased for } \theta \text{ \& } T_1 \text{ is MVUE} \\ \therefore V\left(\frac{T_1+T_2}{2}\right) &\geq V(T_1) \\ \text{or } V(T_1) \frac{1+\rho}{2} &\geq V(T_1) \\ \text{or } \frac{1+\rho}{2} &\geq 1 \quad \text{or } \rho \geq 1 \quad \therefore \boxed{\rho=1} \end{aligned}$$

This is equal to half into variance of  $T_1$  plus rho variance of  $T_1$  is equal to variance of  $T_1$  into  $1 + \rho$  by 2.

Now,  $\frac{T_1 + T_2}{2}$  is unbiased for  $\theta$  and  $T_1$  is minimum variance unbiased estimator. Therefore, variance of  $\frac{T_1 + T_2}{2}$  is greater than variance of  $T_1$  or variance of  $\frac{T_1 + T_2}{2}$  is greater than equal to variance of  $T_1$  or  $1 + \rho$  by 2 is greater than equal to 1 or  $\rho$  greater than equal to 1 and since,  $\rho$  is a correlation coefficient and therefore, the  $\rho$  lies only between minus 1 to 1. We can say that  $\rho$  is equal to 1. What does it mean?

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$\therefore$  The correlation bet<sup>n</sup>  $T_1$  &  $T_2$  is 1  $\therefore T_1$  is linear with  $T_2$   
i.e.  $T_1 = a + bT_2$  where  $b > 0$   
 $\therefore V(T_1) = b^2 V(T_2)$   
or  $V(T_1) = b^2 V(T_1)$   
 $\therefore b^2 = \pm 1$  but since  $b > 0$   
we have  $b = 1$

It means that the correlation between  $T_1$  and  $T_2$  is 1 and therefore,  $T_1$  is linear with  $T_2$  that is  $T_1$  is equal to  $a + bT_2$ , where  $b$  is greater than 0.

Therefore, variance of  $T_1$  is equal to from here,  $b$  square times variance of  $T_2$  or variance of  $T_1$  is equal to  $b$  squared times variance of  $T_1$  has both of them are MVUE therefore, they are equal therefore,  $b$  squared is equal to plus minus 1, but since  $b$  greater than 0. We have  $b$  is equal to 1.

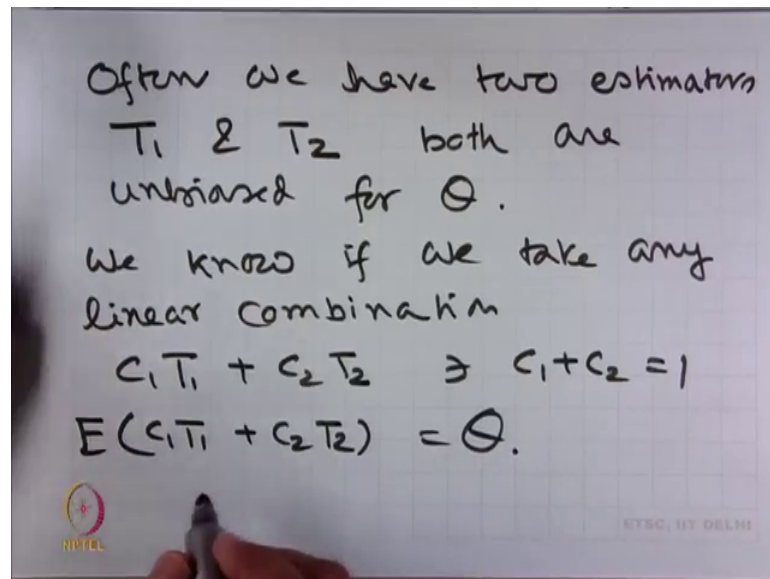
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$\therefore T_1 = a + bT_2$  — (\*)  
 $= a + T_2$  ✓  
 $\therefore E(T_1) = a + E(T_2)$   
 $\therefore$  We have  $a = a + a$   
or  $a = 0$   
 $\therefore$  Putting in (\*)  
 $T_1 = T_2$  QED

Therefore,  $T_1$ , which is equal to  $a + b T_2$  is actually  $a + T_2$  therefore, expected value of  $T_1$  is equal to  $a + \text{expected value of } T_2$  and since, both are unbiased for  $\theta$  therefore, we have  $\theta = a + \theta$  or  $a = 0$ .

Therefore, putting in star  $T_1$  is equal to  $T_2$  Kvd; that means, there can be only one minimum variance unbiased estimator now often.

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We have two estimators  $T_1$  and  $T_2$ , both are unbiased for  $\theta$ , we know if we take any linear combination  $C_1 T_1 + C_2 T_2$  such that  $C_1 + C_2 = 1$  then expected value of  $C_1 T_1 + C_2 T_2$  is equal to  $\theta$  therefore, given any two unbiased estimators. We can generate any number of unbiased estimators for  $\theta$  by taking a linear combination  $C_1 T_1 + C_2 T_2$  and that is going to be unbiased.

The question is should we try all of them to find out which one among them is most efficient or in other words among the all, the linear combinations  $C_1 T_1 + C_2 T_2$ , such that  $C_1 + C_2 = 1$ , which one should we take as the most efficient of them.

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So let  $V(T_1) = \sigma_1^2$   
 $V(T_2) = \sigma_2^2$   
& corr. coeff  $(T_1, T_2) = \rho$   
Then  $V(c_1 T_1 + c_2 T_2)$   
 $= c_1^2 \sigma_1^2 + c_2^2 \sigma_2^2 + 2\rho c_1 c_2 \sigma_1 \sigma_2$   
Putting  $c_2 = 1 - c_1$   
We have  $V(c_1 T_1 + c_2 T_2) =$   
 $c_1^2 \sigma_1^2 + (1 - c_1)^2 \sigma_2^2 + 2\rho(1 - c_1)c_1 \sigma_1 \sigma_2$

So, let variance of T 1 is equal to sigma 1 square variance of T 2 is equal to sigma 2 square and correlation coefficient of T 1 and T 2 is equal to rho then variance of C 1 T 1 plus C 2 T 2 is equal to C 1 square sigma 1 square plus C 2 square sigma 2 square plus 2 times rho C 1 C 2 sigma 1 sigma 2.

Putting C 2 is equal to 1 minus C 1, we have variance of C 1 T 1 plus C 2 T 2 is equal to C 1 square sigma 1 square plus 1 minus C 1 whole square sigma 2 square plus 2 rho 1 minus C 1 into C 1 sigma 1 sigma 2.

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We need to find out  $c_1$  for which this maximum.  
∴ we differentiate wrt  $c_1$  & equate with 0.  
We have variance  $\rightarrow$   
 $c_1^2 \sigma_1^2 + (1 - 2c_1 + c_1^2) \sigma_2^2 + 2\rho c_1 (1 - c_1) \sigma_1 \sigma_2$   
Differentiating wrt  $c_1$ , we have

So, we need to find out  $C_1$  for which this is maximum. Therefore, we differentiate with respect to  $C_1$  and equate with 0, we have variance is equal to  $C_1^2 \sigma_1^2 + (1 - C_1)^2 \sigma_2^2 + 2 C_1 (1 - C_1) \rho \sigma_1 \sigma_2$ .

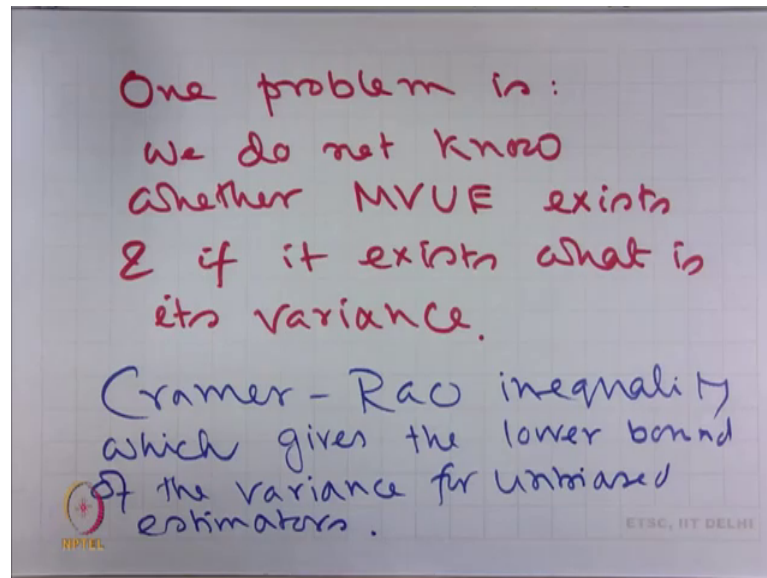
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$$\begin{aligned}
 & 2C_1\sigma_1^2 - 2\sigma_2^2 + 2C_1\sigma_2^2 \\
 & + 2\rho\sigma_1\sigma_2 - 4\rho C_1\sigma_1\sigma_2 = 0 \\
 \text{or } & 2C_1\sigma_1^2 + 2C_1\sigma_2^2 - 4\rho C_1\sigma_1\sigma_2 \\
 & = 2\sigma_2^2 - 2\rho\sigma_1\sigma_2 \\
 \therefore C_1 = & \frac{\sigma_2^2 - \rho\sigma_1\sigma_2}{\sigma_1^2 + \sigma_2^2 - 2\rho\sigma_1\sigma_2} \\
 C_2 = & 1 - C_1
 \end{aligned}$$

By differentiating with respect to  $C_1$  we have  $2 C_1 \sigma_1^2 - 2 \sigma_2^2 + 2 C_1 \sigma_2^2 + 2 \rho \sigma_1 \sigma_2 - 4 \rho C_1 \sigma_1 \sigma_2$  is equal to 0 or  $2 C_1 \sigma_1^2 + 2 C_1 \sigma_2^2 - 4 C_1 \rho \sigma_1 \sigma_2 = 2 \sigma_2^2 - 2 \rho \sigma_1 \sigma_2$ . Therefore,  $C_1$  is equal to  $\frac{\sigma_2^2 - \rho \sigma_1 \sigma_2}{\sigma_1^2 + \sigma_2^2 - 2 \rho \sigma_1 \sigma_2}$  or this is the coefficient for  $T_1$  and  $1 - C_1$  should give you  $C_2$ , which is the coefficient for  $T_2$  such that  $C_1 T_1 + C_2 T_2$  will have the minimum variance.

Therefore, when we have two such estimators, we can always choose in which linear combination we should take them so, that the variance of the summation is minimum.

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One problem is, we do not know whether MVUE exist and if it is exist, what is it? Is variance to find the answer? To this we have something called Crammar Rao inequality, which gives the lower bound of the variance for unbiased estimators. This is not for all estimators, there is a class of estimators, which observe certain properties, which we call regularity or we say regularity assumption under those assumptions. Crammar Rao inequality gives a lower bound for the variance of the unbiased estimators, what is the advantage?

The advantage is that if we find a particular estimator, whose variance is equal to that bound, then we are sure that this is the minimum variance unbiased estimator, because there cannot be any other unbiased estimator, whose variance is lower than that of course, if we focus on the set of estimators, which observe the regularity assumptions. In the next class I shall deal with these and solve certain problems students, see you in the next class.

Thank you so, much.